# Estimating the Impact of School Education on Contraception Use among Adolescents Aged 15–19 in Burkina Faso and Nigeria Using a Heckman Correction Model

Miaba Louise LOMPO<sup>1</sup>, Jean-Louis BAGO<sup>2</sup> & Wamadini dite Minata SOURATIÉ<sup>3</sup>

Received: May 23, 2018 Accepted: June 11, 2018 Online Published: June 18, 2018

doi:10.5539/jel.v7n5p31 URL: https://doi.org/10.5539/jel.v7n5p31

#### **Abstract**

Sexually Transmitted Infections (STIs) still raises serious concerns for adolescents' sexual health in west-African developing countries. To this extent, promoting contraception use among sexually active adolescents is a major key to addressing this problem. Yet, the rate of contraception use by adolescents remains surprisingly low in these countries. Using the Demographic Health Survey of Burkina Faso (2014) and Nigeria (2013), this paper examines the influence of school education on contraception use among sexually active male and female adolescents aged 15-19 in Burkina Faso and Nigeria. The standard estimates using Probit regressions suggest that achieving a primary school education increases the probability of a sexually active adolescent to use contraception by 8.26 percentage points (Burkina Faso) and 17.2 percentage points (Nigeria). This effect increases to 20.3 percentage points (Burkina Faso) and 34.7 percentage points (Nigeria) for adolescents with a secondary or higher school education. However, these baseline estimates are biased because adolescents' decision to engage into sexual activity is not random. In light of this, a Heckman Correction Model (HCM) has been applied to account for this selection bias. The results show that the Probit regressions underestimate the effect of education on adolescents' likelihood to use contraception in Burkina Faso and overestimate this effect in Nigeria. In fact, compared to adolescents with no school education, HCM estimates show that adolescents with primary and secondary (or higher) school education have respectively 10.2 and 24.4 percentage points more in the use of contraception in Burkina Faso and 15.1 and 34 percentage points in Nigeria. Together, these results suggest that the exposure to school education increases contraception use among the adolescents in both Burkina Faso and Nigeria.

Keywords: school education, contraception, adolescent, Heckman Correction Model, sexually transmitted infections

#### 1. Introduction

Despite the great popularization of contraceptive methods, the rate of Sexually Transmitted Infections (STIs) in developing countries remains high among male and female adolescents (Morris & Rushwan, 2015). In fact, the statistics on sexual health indicate that up to 37% of all new HIV infections in Sub-Saharan African countries are among adolescents aged 15-24 (UN, 2016). As an explanation to this fact, Dehne, Riedner, Berer, and Organization (2005) shows that adolescents are less likely to use contraception than adults. Analysing the increase of sexual activities among adolescents in Nigeria, Alabi and Oni (2017) argues that the promotion of contraception use among adolescents is one of the most important public health challenges. In a similar study, Biddlecom, Singh, and Munthali (2007) also finds that the contraceptive methods are still underutilized in Burkina Faso.

In the literature, several studies analysed the determinants of adolescents' contraception use, focusing on the access to efficient contraceptive methods (Arcidiacono, Khwaja, & Ouyang, 2012; DiCenso, Guyatt, Willan, & Griffith, 2002; Klick & Stratmann, 2008) and sex education (Cannonier, 2012; Girma & Paton, 2014; Kohler, Manhart, & Lafferty, 2008; Oettinger, 1999; Sabia, 2006; Santelli, Lindberg, Finer, & Singh; Winner et al., 2012)

<sup>&</sup>lt;sup>1</sup> Department of Information and Communication, Laval University, Canada

<sup>&</sup>lt;sup>2</sup> Department of Economics, CRREP, Laval University, Canada

<sup>&</sup>lt;sup>3</sup> Unit of Training and Research in Economics and Management, University of Ouaga II, Burkina Faso Correspondence: Miaba Louise LOMPO, Department of Information and Communication, Laval University, Canada.

as the main determinants of contraception use. Another set of literature focuses on socio-economic and demographic factors such as the less widespread family planning behavior (Khan, Mishra, Arnold, & Abderrahim, 2007) and the higher fertility preferences (Bongaarts, 2003) as the explanation of the low rate of contraception use in developing countries. Yet, since the last decade, policies promoting family planning in developing countries fail to increase contraception use (Cleland, Ndugwa, & Zulu, 2011).

On the other hand, the effect of education attainment as a key determinant of contraception use is less explored in developing countries although a large body of literature suggests that contraceptive use is strongly associated with schooling (Ainsworth, Beegle, & Nyamete, 1996; Beekle & McCabe, 2006; Hogan, Berhanu, & Hailemariam, 1999; Korra, 2002). A recent study by Larsson and Stanfors (2014) shows that the low education attainment in developing countries is an important barrier to understand the importance and all the benefits of contraception use. Using the demographic health surveys of four developing countries (Ghana, Kenya, Madagascar and Zambia), Larsson and Stanfors (2014) finds a positive relationship between women's education and contraception use. Their sampling method excluded women who had not been sexually active. However, by limiting the sample to sexually active women, their results are likely to suffer from selection bias since being sexually active is not random. Heckman (1979) shows that a sample selection bias arises when the sample of study is not random. In our context, the decision to use or not to use contraception is observed only for adolescents who are actually engaged in sexual activities but the decision to enter in sexual activity is not random (Rashad & Kaestner, 2004). This decision depends on a set of personal and social behaviors such as alcohol consumption, smoking and many other factors which are unmeasured (Rashad & Kaestner, 2004). This sample selection causes the standard estimates (probit, logit, linear probability model) to be inconsistent. To access the true causal effect on education on contraception use, a HCM is applied to account for this selection issue. To the best of our knowledge, this is the first paper, to account for the sample selection in estimating the effect of education on contraception use in developing countries.

This article is based on data from a low-income country (Burkina Faso) and middle-income country (Nigeria) to account for the level of economic development. In fact, Nigeria was ranked as the most developed country in Africa according to the Gross Domestic Product (GDP-2017) while Burkina Faso was ranked as one of the poorest (World Bank, 2017). In addition, the youth literacy rate was estimated at 21.6% in 2013 in Burkina Faso and 66.38 in Nigeria (World Bank, 2014). Interestingly, the statistics in our data reveals that the level of contraception use among adolescents aged 15-19 in Nigeria is 38.4% while it is 26 % in Burkina Faso, which represents a huge difference of 12% percentage points.

In this empirical analysis, Demographic Health Surveys of Burkina Faso (2014) and Nigeria (2013) are used to examine the influence of school education on contraception use among the sexually active adolescents aged 15-19. The study focuses on both male and female adolescents. The estimation strategy comprises two main steps. First, the effect of education is estimated using a standard Probit model. Second, to access a consistent estimate of the effect of education on contraception use, a HCM is applied to account for the sample selection.

In the case of Burkina Faso, the baseline results suggest that having a primary education increases the adolescents' probability of using contraception during intercourse by 8.26 percentage points compared to adolescents without education at all. This effect increases to 20.3 percentage points for adolescents with a secondary education or more. In the case of Nigeria, the effect of education on contraception use is also positive and statistically significant. The extent of this effect is larger in Nigeria compared to Burkina Faso. In fact, the probability that an adolescent with a primary school education will use contraception is 17.2 percentage points higher than the probability of an adolescent who has not school education. It is also observed that having a secondary school education or tertiary education increases the probability that an adolescent will use contraception by 34.7 percentage points. Hence, these results confirm that education has a positive effect on contraception use in Burkina Faso and Nigeria. However, a potential drawback of this standard Probit approach is that it may yield inconsistent estimates because of the selection of an adolescent in sexual activity.

The consistent results from the HCM show that the Probit estimates are biased. In Burkina Faso, the results suggest that having a basic education increased the probability to use contraception by 10.5 percentage point which is higher than the Probit estimates. The effect of secondary education is also higher (24.4 percentage points) than the Probit estimates. In Nigeria, having a primary education increase the probability of using contraception by 15.1 percentage point which smaller than the standard Probit estimates. The HCM estimated effect of secondary and higher education stood at 34 percentage points which is also smaller than the effect estimated in baseline Probit.

The rest of this paper is structured as follows. Section 2 presents the data and identification strategy, Section 3

presents the descriptive statistics and Section 4 presents the estimation results. Section 5 discusses the results and Section 6 provides a conclusion to the paper.

# 2. Data and Identification Strategy

#### 2.1 Data

This article is based on data from Demographic and Health Surveys of Burkina Faso (2014) and Nigeria (2013). It examines the effect of male and female adolescents' education on the likelihood that she/he uses contraception during intercourse. This database contains information about participants' age, level of education, gender, sexual activity, age at first sex, number of sexual partners, exposure to media, access to information about STIs and contraception use. Education is defined as a categorial variable equals 0 if the adolescent reported that he has never been at school, 1 if he has a primary school education and 2 if he has a secondary or higher school education. The dependent variable of interest for contraception use was built using the questionnaire for instance, the respondents were asked about various contraceptive method used. Subsequently, the contraception use is defined as a dummy variable equals 1 if the adolescent reported that he/she uses a contraceptive method during sex and 0 otherwise. The data were restricted to individuals aged between 15 and 19 years old. Our main sample contains 1722 adolescents in Burkina Faso and 3983 adolescents in Nigeria.

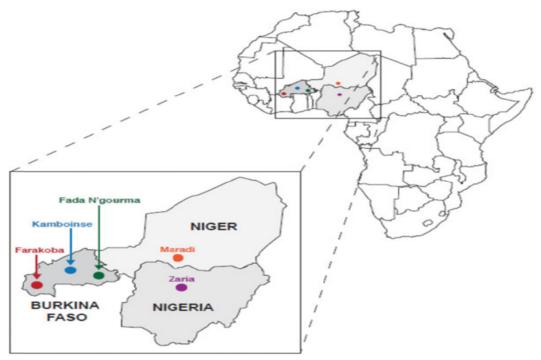


Figure 1. Maps showing location of Nigeria and Burkina Faso

## 2.2 Identification Strategy

This paper uses two identification methods to estimate the effect of an adolescent's education on the likelihood that she is using contraception or not. First, the baseline strategy uses standard Probit regression. Then, the HCM for sample selection approach is estimated as a second identification strategy.

# 2.2.1 Baseline Strategy

Our primary interest is to estimate a standard Probit model. Let i denotes a sexually active adolescent characterized by a level of Education  $Edu_i$  and a vector of other characteristics  $X_i$ .  $X_i$  includes age, gender, age at first sex, number of sexual partners, exposure to media, access to information about STIs and the adolescent's area of residence.

 $Y_i$  denotes the decision about contraception use (equals 1 if he/she uses contraception and 0 otherwise). The decision can be written as follows:

$$Y_i = \begin{cases} 1 & if \ Y_i^* = \alpha_0 + \sum_{j=1}^2 \alpha_{j1} E du_{ji} + X_i' \gamma + \varepsilon > 0 & (Contraception \ use) \\ 0 & if \ Y_i^* = \alpha_0 + \sum_{j=1}^2 \alpha_{j1} E du_{ji} + X_i' \gamma + \varepsilon < 0 & (No \ contraception \ use) \end{cases}$$
 (1)

where  $Y_i^*$  is the latent variable whose sign determines whether to use or not to use contraception and  $\varepsilon$  is a zero-mean error term.  $\beta$  captures the parameters of covariates that affect the decision that adolescents will use contraception.

The probability that contraception is used can then be written as follows:

$$P(Y_i = 1 | X_i) = P(Y_i^* > 0)$$

$$= P\left(\alpha_0 + \sum_{j=1}^2 \alpha_j E du_{ji} + X_i' \beta > 0\right)$$

$$= P\left(\varepsilon < \alpha_0 + \sum_{j=1}^2 \alpha_j E du_{ji} + X_i' \beta\right)$$

$$= \Phi\left(\alpha_0 + \sum_{j=1}^2 \alpha_j E du_{ji} + X_i' \beta\right)$$

Where  $\Phi(\cdot)$  is the normal distribution. This probability and the parameters of the model are estimated using a maximum likelihood. To deal with the potential clustering of observations at the neighbourhood level, the model is estimated using heteroskedasticity robust standard errors. In order to conduct the interpretations, the marginal effect of each variable on the probability of using contraception are estimated.

However, baseline estimates of Equation (1) are subjected to bias due to selection bias. Indeed, the behavior about contraception of adolescents who are actually engaged in sexual activity is observed. Rashad and Kaestner (2004) shows that the decision to enter in sexual activity is not random and depends on a set of personal and social behaviors such as alcohol consumption and many others which are unmeasured. In a similar analysis, Evans, Oates, and Schwab (1992) shows that peer group influence also affects adolescents' sex behavior. In this context, the standard Probit estimates become subject to sample selection bias. This paper uses a HCM to account for this selection issue.

### 2.2.2 Heckman Correction Model

As suggested in the previous section, the main challenge in identifying the effect of education on the decision to use contraception or not is the presence of potential sample selection issues. Let  $S_i$  denotes the selection variable and equals 1 if an adolescent is engaged in sexual activity and 0 otherwise. In our paradigm, the HCM (Gronau, 1974; Heckman, 1979; Lewis, 1974) assumes that there exists an underlying regression relationship between education and contraception use following Equation (1). However, the sample selection bias arises because the decision to use contraception (or not) is observed only when an adolescent is engaged in sexual activity ( $S_i$ =1). In fact, this decision is not observed for an adolescent who has never engaged in sexual activity ( $S_i$ =0). Given that the decision of adolescents to engage into sexual activity is not random, adolescents who are engaged in sexual activity might share some common characteristics. In such case, the standard Probit estimates are likely to be biased.

The HCM estimation includes the analysis of the following selection equation:

$$S_{i} = \begin{cases} 1 & if S_{i}^{*} = Z_{i}'\pi + \mu > 0 \quad (Sexually \ active) \\ 0 & if S_{i}^{*} = Z_{i}'\pi + \mu < 0 \quad (Not \ sexually \ active) \end{cases}$$
 (2)

where  $Z_i$  includes all of the factors that affect the decision to enter in sexual activity.

The probability to enter in sexual activity can then be written as follows:

$$P(Y_i = 1|X_i) = P(S_i^* > 0)$$
  
=  $P(Z_i'\pi + \mu > 0)$   
=  $P(\mu < Z_i'\pi)$   
=  $\Phi(Z_i'\pi)$ 

where  $\Phi(.)$  is the normal distribution.

HCM accounts for the selection bias (Heckman, 1979) by using a two-step estimator. To access the effect of

education on contraception use in the presence of sample selection, this paper follows Wooldridge (2009) to estimate the model in two stages: The first stage predicts the likelihood of an adolescent to be in sexually active using a Probit model and calculate the predicted inverse Mills ratio. The second stage estimates the decision about contraception including the inverse Mills ratio as a predictor in the model. The presence of the selection bias is captured by the coefficient  $\lambda$  of the inverse of Mills ratio. Typically, if  $\lambda$  is statistically equal to zero, then there is no issue of sample selection. In contrary, if  $\lambda$  is statistically different from zero, then there is a sample selection issue and the baseline estimates are biased.

# 3. Descriptive Statistics

Table (1) presents the summary of statistics. Two noteworthy features concerning the relationship between education and contraception use appear in these statistics. In one hand, one stunning fact between the two countries is the education gap: In Burkina Faso, the statistics suggest that 57.2% of adolescents have no education while only 20% and 22.8% have reached primary and secondary school respectively. This low literacy rate in the main sample reflects the youth literacy rate for the country, estimated at 21.6% in 2013 (World Bank, 2014). In contrast, Nigeria presents a much higher level of adolescents' education. In fact, only 6.67% of adolescents have no education in Nigeria, which represents a huge difference of more than 50 percentage points. For secondary education and higher, the difference is even more stunning.

In fact, Table (1) shows that 81.7 % of adolescents in Nigeria have at least a secondary degree. Compared to Burkina Faso this represents a difference of 58.9 percentage points. Interestingly, the results show that the rate of contraception use is also low among adolescents aged 15-19 in Burkina Faso (26 percent) than in Nigeria (38.4 percent). On the other hand, by carefully considering each country, Figure 2 and Figure 3 reveal a positive association between the level of education and contraception use. In fact, Figure 2 shows that only 10.1% of adolescents without education use contraception while more than the half (52.6%) of adolescents with secondary education use contraception in Burkina Faso. The same feature is observed in Nigeria (Figure 3) where barely 0.7% of uneducated adolescents use contraception. This rate increased to 35.6% when the adolescent has a secondary education. Put together, these two features suggest that education is positively correlated with contraception use in both countries.

For other statistics, as far as the exposure to media is concerned, more than the two thirds of adolescents have access to media channels such as radio, TV and/or internet in both countries. The statistics also show that almost a half (48.8% in Burkina Faso and 48.3% in Nigeria) have access to information about STIs. The main sample contains 83% of female adolescents in Burkina Faso and 84.5% in Nigeria. Statistics also reveal that adolescents in Nigeria enter in sexual activity earlier than those in Burkina Faso. On average, adolescents enter into sexual activity earlier in Nigeria (15.66 years old) than in Burkina Faso (16.3 years old).

Table 1. Descriptive statistics

| VARIABLES          | Burkina Faso |        |       |     | Nigeria | Nigeria |        |       |     |     |
|--------------------|--------------|--------|-------|-----|---------|---------|--------|-------|-----|-----|
|                    | N            | mean   | sd    | min | max     | N       | mean   | sd    | min | max |
| Contraception use  | 1,410        | 0.260  | 0.439 | 0   | 1       | 1,844   | 0.384  | 0.487 | 0   | 1   |
| Education level    |              |        |       |     |         |         |        |       |     |     |
| No education       | 1,410        | 0.572  | 0.495 | 0   | 1       | 1,844   | 0.0667 | 0.250 | 0   | 1   |
| Primary            | 1,410        | 0.200  | 0.400 | 0   | 1       | 1,844   | 0.117  | 0.321 | 0   | 1   |
| Secondary and more | 1,410        | 0.228  | 0.420 | 0   | 1       | 1,844   | 0.817  | 0.387 | 0   | 1   |
| Information on STI | 1,410        | 0.543  | 0.498 | 0   | 1       | 1,844   | 0.714  | 0.452 | 0   | 1   |
| Female             | 1,410        | 0.800  | 0.400 | 0   | 1       | 1,844   | 0.714  | 0.452 | 0   | 1   |
| Age dummy          |              |        |       |     |         |         |        |       |     |     |
| 15 years old       | 1,410        | 0.0326 | 0.178 | 0   | 1       | 1,844   | 0.0705 | 0.256 | 0   | 1   |
| 16 years old       | 1,410        | 0.106  | 0.308 | 0   | 1       | 1,844   | 0.0987 | 0.298 | 0   | 1   |
| 17 years old       | 1,410        | 0.199  | 0.400 | 0   | 1       | 1,844   | 0.175  | 0.380 | 0   | 1   |
| 18 years old       | 1,410        | 0.307  | 0.461 | 0   | 1       | 1,844   | 0.312  | 0.464 | 0   | 1   |
| 19 years old       | 1,410        | 0.355  | 0.479 | 0   | 1       | 1,844   | 0.343  | 0.475 | 0   | 1   |
| Rural resident     | 1,410        | 0.621  | 0.485 | 0   | 1       | 1,844   | 0.572  | 0.495 | 0   | 1   |
| Age at first sex   | 1,410        | 16.03  | 1.424 | 10  | 19      | 1,844   | 15.66  | 1.860 | 8   | 19  |
| Number of partners | 1,410        | 0.913  | 0.470 | 0   | 7       | 1,844   | 0.893  | 0.449 | 0   | 3   |
| Access to condom   | 1,410        | 0.563  | 0.496 | 0   | 1       | 1,844   | 0.553  | 0.497 | 0   | 1   |
| Exposure to Media  | 1,410        | 0.842  | 0.365 | 0   | 1       | 1,844   | 0.892  | 0.310 | 0   | 1   |

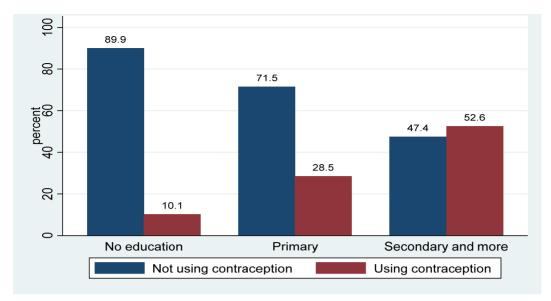


Figure 2. Percentage of adolescents using contraception by level of education in Burkina Faso

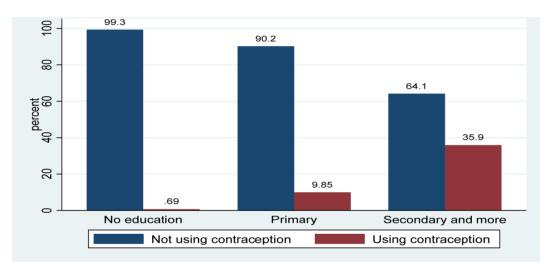


Figure 3. Percentage of adolescents using contraception by level of education in Nigeria

#### 4. Estimation Results

#### 4.1 Probit Estimates

The Probit estimates of Equation (1) are presented in Table (2). Columns (2) and (4) present the marginal effects from Burkina Faso and Nigeria respectively. In the case of Burkina Faso, the results suggest that having a primary school education increases the probability that an adolescent uses contraception during intercourse by 8.26 percentage points compared to an adolescent without education at all. This effect increases to 20.3 percentage points for adolescents with a secondary school education or more. In the case of Nigeria, the effect of education on contraception use also appears positive and even higher compared to Burkina Faso. In fact, the probability that an adolescent with a primary school education will use contraception is 17.2 percentage points higher than the probability of an adolescent who does not attend school. Moreover, being in secondary school or higher increases the probability that an adolescent will use contraception by 34.7 percentage points. Hence, these results suggest that adolescents' exposure to school education has a positive effect on the probability that they will use contraception. In addition, by controlling for the access to information about STI, the estimates suggest that adolescents who have been provided with information about STI are more likely to use contraception than adolescents who have no information on STIs. In addition, the results suggest that female adolescents have more probability to use contraception in Nigeria. However, females have a lower probability to use contraception in

Burkina Faso. The model also controls for age and find that contraception use does not depend on an adolescent' age.

Looking at the number of partners, the results suggest that having one additional partners increases the probability that adolescent use contraception by 28.4 percentage points in Burkina Faso and 31.8 percentage points in Nigeria. In fact, these results are consistent with the fact that having several sex partners increases the risk of STIs and hence, the need of protection. For both countries, adolescents who have better access to condoms tends to use more contraception than their counterpart who do not have access to condoms. This result is consistent with the negative effect of living in rural areas in both two countries. The results also suggest that the exposure to media such as radio, TV, boosts awareness creation and thus increases the probability to use contraception methods in Burkina Faso but not in Nigeria. As far as the age at first sex is concerned, the data suggest that adolescents who enter early in sexual activity tend to use more contraception in Nigeria and less in Burkina Faso.

However, these baseline estimates might be biased because of the sample selection since the entry into sexual activity is not random. The HCM is used to correct for this selection bias in order to estimate consistent effects.

Table 2. Standard Probit estimates

| Dependent variable is Contraception use (0/1) | Burkina Faso |                  | Nigeria      |                  |  |
|---|--------------|------------------|--------------|------------------|--|
|   | (1)          | (2)              | (3)          | (4)              |  |
| VARIABLES                                     | Coefficients | Marginal effects | Coefficients | Marginal effects |  |
| Education Base: No education                  |              |                  |              |                  |  |
| Primary                                       | 0.375***     | 0.0826***        | 0.860***     | 0.172***         |  |
|   | (0.112)      | (0.0259)         | (0.229)      | (0.0379)         |  |
| Secondary and more                            | 0.822***     | 0.203***         | 1.417***     | 0.347***         |  |
|   | (0.118)      | (0.0322)         | (0.213)      | (0.0291)         |  |
| Access to information on STI                  | 0.323***     | 0.0693***        | 0.275***     | 0.0867***        |  |
|   | (0.0959)     | (0.0206)         | (0.0731)     | (0.0229)         |  |
| Female dummy                                  | -0.914***    | -0.196***        | 0.199**      | 0.0627**         |  |
|   | (0.117)      | (0.0231)         | (0.0829)     | (0.0261)         |  |
| Age dummy (Base: 15 years)                    |              |                  |              |                  |  |
| 16 years old                                  | 0.255        | 0.0497           | -0.0977      | -0.0305          |  |
|   | (0.276)      | (0.0518)         | (0.161)      | (0.0505)         |  |
| 17 years old                                  | 0.401        | 0.0813*          | 0.122        | 0.0391           |  |
|   | (0.263)      | (0.0489)         | (0.150)      | (0.0476)         |  |
| 18 years old                                  | 0.225        | 0.0435           | 0.0325       | 0.0103           |  |
|   | (0.263)      | (0.0483)         | (0.144)      | (0.0456)         |  |
| 19 years old                                  | 0.393        | 0.0794           | -0.0500      | -0.0157          |  |
|   | (0.271)      | (0.0503)         | (0.147)      | (0.0462)         |  |
| Age at first sex                              | -0.0831**    | -0.0178**        | 0.0435**     | 0.0137**         |  |
|   | (0.0345)     | (0.00737)        | (0.0203)     | (0.00639)        |  |
| Number of partners                            | 1.324***     | 0.284***         | 1.008***     | 0.318***         |  |
|   | (0.176)      | (0.0338)         | (0.0947)     | (0.0267)         |  |
| Access to condom                              | 0.460***     | 0.0987***        | 0.550***     | 0.174***         |  |
|   | (0.103)      | (0.0217)         | (0.0734)     | (0.0220)         |  |
| Exposure to media                             | 0.259*       | 0.0556*          | 0.126        | 0.0396           |  |
|   | (0.151)      | (0.0323)         | (0.121)      | (0.0382)         |  |
| Residence area                                | -0.462***    | -0.0991***       | -0.205***    | -0.0647***       |  |
|   | (0.0945)     | (0.0197)         | (0.0665)     | (0.0208)         |  |
| Constant                                      | -0.994*      |                  | -3.840***    |                  |  |
|   | (0.576)      |                  | (0.405)      |                  |  |
| Observations                                  | 1,402        | 1,402            | 1,829        | 1,829            |  |

*Note.* Robust standard errors in parentheses \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

## 4.2 HCM Estimates

Table (3) presents the consistent HCM estimates. The statistics of the coefficients  $\lambda$  of the Mills ratio of both Burkina Faso and Nigeria are significant. This confirms the presence of selection bias in the baseline estimates and supports the use of HCM as a consistent identification strategy of the effect of education on contraception use. In Burkina Faso, the results suggest that having a primary school education increases the probability of contraception use by 10.5 percentage points which is higher than the baseline estimates in Table (2). The effect of secondary school education estimated by HCM is also 24.4 percentage points higher than the Probit estimates. In Nigeria, having a primary school education increases the probability of using contraception by 15.1 percentage points which smaller than the standard Probit estimates. The estimate from HCM for secondary and higher school is 34 percentage points which is also smaller than the effect estimated in the baseline Probit. To summarize, the study finds that school education has a positive effect on contraception use for both countries. However, the simple Probit estimates underestimate this effect in the case of Burkina Faso and overestimate the effect in the case of Nigeria.

Table 3. HCM estimates

| Dependent variable is            | Burkina Faso | Nigeria    |  |  |
|----------------------------------|--------------|------------|--|--|
| Contraception use (0/1)          |              |            |  |  |
| Education: Base (No education)   |              |            |  |  |
| Primary                          | 0.102***     | 0.151***   |  |  |
|                                  | (0.0260)     | (0.0500)   |  |  |
| Secondary and more               | 0.244***     | 0.340***   |  |  |
|                                  | (0.0277)     | (0.0435)   |  |  |
| Female                           | -0.287***    | 0.0362     |  |  |
|                                  | (0.0272)     | (0.0261)   |  |  |
| Age group (Base=15 years old)    |              |            |  |  |
| 16 years old                     | 0.0769       | -0.0519    |  |  |
| ·                                | (0.0607)     | (0.0505)   |  |  |
| 17 years old                     | 0.103*       | 0.0217     |  |  |
| •                                | (0.0582)     | (0.0466)   |  |  |
| 18 years old                     | 0.0621       | -0.0116    |  |  |
| •                                | (0.0578)     | (0.0450)   |  |  |
| 19 years old                     | 0.0975*      | -0.0395    |  |  |
|                                  | (0.0584)     | (0.0456)   |  |  |
| Age at first sex                 | -0.0233***   | 0.00632    |  |  |
|                                  | (0.00765)    | (0.00634)  |  |  |
| Number of sex partners           | 0.122***     | 0.162***   |  |  |
| F                                | (0.0311)     | (0.0435)   |  |  |
| Access to condom                 | 0.0910***    | 0.176***   |  |  |
| Tables to Condom                 | (0.0221)     | (0.0232)   |  |  |
| Access to information about STI  | 0.0730***    | 0.0759***  |  |  |
| riccess to information about 511 | (0.0207)     | (0.0229)   |  |  |
| Exposure to media                | 0.0409       | 0.0362     |  |  |
| Exposure to media                | (0.0270)     | (0.0352)   |  |  |
| Rural resident                   | -0.123***    | -0.0801*** |  |  |
| Kurar resident                   | (0.0224)     | (0.0214)   |  |  |
| Constant                         | 0.571***     | -0.266**   |  |  |
| Constant                         | (0.137)      | (0.129)    |  |  |
| Savual activity (0/1)            | (0.137)      | (0.129)    |  |  |
| Sexual activity (0/1)            |              |            |  |  |
| Education: Base (No education)   | 0.125        | 0.650***   |  |  |
| Primary                          | -0.125       | 0.659***   |  |  |
| 2 1 1                            | (0.107)      | (0.236)    |  |  |
| Secondary and more               | -0.547***    | 0.474**    |  |  |
| - ·                              | (0.113)      | (0.218)    |  |  |
| Female                           | 0.449***     | 0.311***   |  |  |
|                                  | (0.0996)     | (0.0740)   |  |  |
| Age group (Base=15 years old)    |              |            |  |  |
| 16 years old                     | 0.774***     | 0.292**    |  |  |
|                                  | (0.196)      | (0.115)    |  |  |
| 17 years old                     | 0.993***     | 0.460***   |  |  |
| <u> </u>                         | (0.193)      | (0.112)    |  |  |

| 18 years old                    | 1.521***  | 0.775***  |  |
|---------------------------------|-----------|-----------|--|
|                                 | (0.192)   | (0.105)   |  |
| 19 years old                    | 1.904***  | 0.926***  |  |
|                                 | (0.191)   | (0.111)   |  |
| Age at first sex                | 0.196**   | 0.294***  |  |
|                                 | (0.0977)  | (0.0733)  |  |
| Access to condom                | 0.169*    | 0.244***  |  |
|                                 | (0.0894)  | (0.0690)  |  |
| Access to information about STI | -0.152    | -0.0894   |  |
|                                 | (0.114)   | (0.120)   |  |
| Exposure to media               | 0.0851    | 0.190***  |  |
|                                 | (0.0966)  | (0.0653)  |  |
| constant                        | -2.740*** | -2.870*** |  |
|                                 | (0.242)   | (0.251)   |  |
| Mills ratio: lambda             | -0.128*** | -0.124*** |  |

*Note.* Robust standard errors in parentheses \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

## 5. Discussion: How Does Education Attainment Increase Contraception Use?

This section focuses on the mechanisms driving the positive relationship between adolescents' education attainment and their contraception use. Four main channels have been identified in the literature to explain the effect of education on contraception use (Gordon, Sabates, Bond, & Wubshet, 2011; Hogan et al., 1999) (Jejeebhoy, 1995; Rosenzweig & Schultz, 1989). These channels include (i) knowledge (ii) access to health services (iii) attitudinal factors and (iv) family planning.

Knowledge: In a study in Ethiopia, Gordon et al. (2011) shows that the effect of education on contraception use is mainly mediated by knowledge. First, education leads to a better comprehension of STIs risks and the benefit of contraceptive choices. Secondly, education increases the knowledge of correct contraception method use and the effectiveness of each contraception method. Knowledge increases adolescents' willingness to use contraception (Anochie & Ikpeme, 2003; Hogan et al., 1999; Rosenzweig & Schultz, 1989). Finally, Anochie and Ikpeme (2003) finds that education address adolescent's misconceptions about effective methods of contraception, leading them to a better understanding of contraceptive methods. In terms, through the increase of knowledge about sexuality and contraception, education promotes the use of contraception among adolescents.

Access to health services: The second channel through which education increase adolescents' contraception use is the access to health services. In fact, adolescents with a higher education are more likely to visit a doctor or a health professional to seek for information and counseling about sexuality and go on to use contraception (Gordon et al., 2011; Jejeebhoy, 1995).

Attitude toward contraception: (Gordon et al., 2011) shows that education leads to positive attitudes of adolescents about contraception by reducing the fears of a particular method of contraception.

Family Planning: Studies on fertility in developing countries indicate that education increases the need for birth control (DeCicca & Krashinsky, 2015) and by this way promotes the use of contraception (Ferré, 2009).

## 6. Conclusion

This paper analyses the effect of school education on contraception use among adolescents aged 15-19 in Burkina Faso and Nigeria. Using the Demographic Health Survey of Burkina Faso (2014) and Nigeria (2013), the baseline estimates use Probit regressions to estimate the effect of education on adolescent contraception. As Larsson and Stanfors (2014), the baseline analysis excludes the adolescents who have never engaged in sexual activity. The research findings show that adolescents who achieved a greater level of education have a higher probability to use contraception. However, this method leads to inconsistent estimates since being sexually active is not random.

Subsequently, to access consistent estimates of the effect of education on contraception use, a HCM is applied to account for this selection bias. The findings suggest that education increase contraception use among adolescents in both Burkina Faso and Nigeria. In Burkina Faso, the results suggest that having a primary education increased the probability to use contraception by 10.5 percentage points which is higher than the Probit estimates. The effect of secondary education is also higher 24.4 percentage points than the Probit estimates. In Nigeria, having a primary school education increases the probability of using contraception by 15.1 percentage points which is smaller than the standard Probit estimates. The HCM's estimated effect of secondary and higher education amounted to 34 percentage point which is also smaller than the effect estimated in baseline Probit.

Several channels such as knowledge, access to health services, attitudinal factors and family planning are advocated to explain the positive relationship between educational attainment and adolescents' probability to use contraception. However, because of the unavailability of data, this study does not focus on sex education particularly but on the exposure to school education in general. It also does not account for quality of education because of missing data on the quality of education in our database. In other words, does the link between education and contraception use depend on the content and the quality of school teaching? An interesting direction for future research will be to examine the relationship between the content and the quality of education and adolescent's contraception use in developing countries.

#### References

- Ainsworth, M., Beegle, K., & Nyamete, A. (1996). The impact of women's schooling on fertility and contraceptive use: A study of fourteen sub-Saharan African countries. *The World Bank Economic Review,* 10(1), 85-122. https://doi.org/10.1093/wber/10.1.85
- Alabi, O., & Oni, I. O. (2017). Teenage Pregnancy in Nigeria: Causes, Effect and Control. *International Journal of Academic Research in Business and Social Sciences*, 7(2), 17-32. Retrieved from http://hrmars.com/hrmars papers/Teenage Pregnancy in Nigeria Causes, Effect and Control.pdf
- Anochie, I., & Ikpeme, E. (2003). The knowledge, attitude and use of contraception among secondary school girls in Port Harcourt. *Nigerian journal of medicine: journal of the National Association of Resident Doctors of Nigeria*, 12(4), 217-220. Retrieved from http://europepmc.org/abstract/med/14768197
- Arcidiacono, P., Khwaja, A., & Ouyang, L. (2012). Habit persistence and teen sex: Could increased access to contraception have unintended consequences for teen pregnancies? *Journal of Business & Economic Statistics*, 30(2), 312-325. https://doi.org/10.1080/07350015.2011.652052
- Beekle, A., & McCabe, C. (2006). Awareness and determinants of family planning practice in Jimma, Ethiopia. *International Nursing Review, 53*(4), 269-276. https://doi.org/10.1111/j.1466-7657.2006.00492.x
- Biddlecom, A. E., Singh, S., & Munthali, A. (2007). Adolescents' views of and preferences for sexual and reproductive health services in Burkina Faso, Ghana, Malawi and Uganda. *African Journal of Reproductive Health*, 11(3), 99-110. https://doi.org/10.2307/25549734
- Bongaarts, J. (2003). Completing the fertility transition in the developing world: The role of educational differences and fertility preferences. *Population Studies*, 57(3), 321-335. https://doi.org/10.1080/0032472032000137835
- Cannonier, C. (2012). State abstinence education programs and teen birth rates in the US. *Review of Economics of the Household, 10*(1), 53-75. https://doi.org/10.1007/s11150-011-9131-8
- Cleland, J. G., Ndugwa, R. P., & Zulu, E. M. (2011). Family planning in sub-Saharan Africa: progress or stagnation? *Bulletin of the World Health Organization*, 89(2), 137-143. Retrieved from https://www.scielosp.org/scielo.php?pid=S0042-96862011000200013&script=sci\_arttext&tlng=en
- DeCicca, P., & Krashinsky, H. (2015). Does education reduce teen fertility? Evidence from Compulsory Schooling Laws.
- Dehne, K. L., Riedner, G., Berer, M., & Organization, W. H. (2005). Sexually transmitted infections among adolescents: the need for adequate health services. Retrieved from http://apps.who.int/iris/bitstream/handle/10665/43221/9241562889.pdf;jsessionid=27F32DD9DFD23249C 7E79A41EDCEDED3?sequence=1
- DiCenso, A., Guyatt, G., Willan, A., & Griffith, L. (2002). Interventions to reduce unintended pregnancies among adolescents: systematic review of randomised controlled trials. *Bmj*, 324(7351), 1426. https://doi.org/10.1136/bmj.324.7351.1426
- Evans, W. N., Oates, W. E., & Schwab, R. M. (1992). Measuring peer group effects: A study of teenage behavior. *Journal of Political Economy, 100*(5), 966-991. https://doi.org/10.1086/261848
- Ferré, C. (2009). Age at first child: does education delay fertility timing? The case of Kenya. Retrieved from https://papers.ssrn.com/sol3/papers.cfm?abstract\_id=1344718
- Girma, S., & Paton, D. (2014). Is education the best contraception? the case of declining teenage pregnancy in england. Retrieved from https://papers.ssrn.com/sol3/papers.cfm?abstract\_id=2462432
- Gordon, C., Sabates, R., Bond, R., & Wubshet, T. (2011). Women's education and modern contraceptive use in Ethiopia. *International Journal of Education*, 3(1), 1. https://doi.org/10.5296/ije.v3i1.622

- Gronau, R. (1974). Wage comparisons—A selectivity bias. *Journal of Political Economy*, 82(6), 1119-1143. https://doi.org/10.1086/260267
- Heckman, J. J. (1979). *Statistical models for discrete panel data*: Department of Economics and Graduate School of Business, University of Chicago Chicago, IL.
- Hogan, D. P., Berhanu, B., & Hailemariam, A. (1999). Household organization, women's autonomy, and contraceptive behavior in southern Ethiopia. *Studies in family planning*, 30(4), 302-314. https://doi.org/10.1111/j.1728-4465.1999.t01-2-.x
- Jejeebhoy, S. J. (1995). Women's education, autonomy, and reproductive behaviour: Experience from developing countries. *OUP Catalogue*. Retrieved from https://ideas.repec.org/b/oxp/obooks/9780198290339.html
- Khan, S., Mishra, V., Arnold, F., & Abderrahim, N. (2007). Contraceptive trends in developing countries. Retrieved from https://www.popline.org/node/199753
- Klick, J., & Stratmann, T. (2008). Do spa visits improve health: evidence from German micro data. *Eastern Economic Journal*, 34(3), 364-374. https://doi.org/10.1057/palgrave.eej.9050038
- Kohler, P. K., Manhart, L. E., & Lafferty, W. E. (2008). Abstinence-only and comprehensive sex education and the initiation of sexual activity and teen pregnancy. *Journal of Adolescent Health*, 42(4), 344-351. https://doi.org/10.1016/j.jadohealth.2007.08.026
- Korra, A. (2002). Attitudes toward family planning and reasons for nonuse among women with unmet need for family planning in Ethiopia: Citeseer.
- Larsson, C., & Stanfors, M. (2014). Women's education, empowerment, and contraceptive use in sub-Saharan Africa: findings from recent demographic and health surveys. *Etude de la Population Africaine*, 28(2), 1022. https://doi.org/10.11564/28-0-554
- Lewis, H. G. (1974). Comments on selectivity biases in wage comparisons. *Journal of Political Economy*, 82(6), 1145-1155. https://doi.org/10.1086/260268
- Morris, J. L., & Rushwan, H. (2015). Adolescent sexual and reproductive health: The global challenges. *International Journal of Gynecology & Obstetrics*, 131(S1). https://doi.org/10.1016/j.ijgo.2015.02.006
- Oettinger, G. S. (1999). The effects of sex education on teen sexual activity and teen pregnancy. *Journal of Political Economy*, 107(3), 606-644. https://doi.org/10.1086/250073
- Rashad, I., & Kaestner, R. (2004). Teenage sex, drugs and alcohol use: problems identifying the cause of risky behaviors. *Journal of health economics*, 23(3), 493-503. https://doi.org/10.1016/j.jhealeco.2003.09.009
- Rosenzweig, M. R., & Schultz, T. P. (1989). Schooling, information and nonmarket productivity: contraceptive use and its effectiveness. *International Economic Review*, 457-477. https://doi.org/10.2307/2526657
- Sabia, J. J. (2006). Does sex education affect adolescent sexual behaviors and health? *Journal of Policy Analysis and Management*, 25(4), 783-802. https://doi.org/10.1002/pam.20208
- Santelli, J. S., Lindberg, L., Finer, L. B., & Singh, S. Recent declines in adolescent pregnancy in the United States: More abstinence or better contraceptive use? Paper presented at the Annual Meeting.
- UN. (2016). *Global AIDS up date*. Retrieved from http://www.who.int/hiv/pub/arv/global-AIDS-update-2016\_en.pdf
- Winner, B., Peipert, J. F., Zhao, Q., Buckel, C., Madden, T., Allsworth, J. E., & Secura, G. M. (2012). Effectiveness of long-acting reversible contraception. *New England Journal of Medicine*, *366*(21), 1998-2007. https://doi.org/10.1056/NEJMoa1110855
- Wooldridge, M. (2009). An introduction to multiagent systems: John Wiley & Sons.
- World Bank. (2014). Education Statistics, country stats. In: World Bank Group.
- World Bank. (2017). Gross Domestic Product per capita, Country profile. In: World Bank Group.

## Copyrights

Copyright for this article is retained by the author, with first publication rights granted to the journal.

This is an open-access article distributed under the terms and conditions of the Creative Commons Attribution license (http://creativecommons.org/licenses/by/4.0/).