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**ABSTRACT**

Multidimensional item difficulty (MID) is proposed as a means of describing test items which measure more than one ability. With mathematical story problems, for instance, both mathematical and verbal skills are required to obtain a correct answer. The proposed measure of MID is based upon three general assumptions: (1) the probability of answering an item correctly increases monotonically with an increase in each dimension being measured; (2) it is desirable to locate an item at a single point in a multidimensional space; and (3) the best point to use in defining the MID for an item in the multidimensional space is the point where the item is most discriminating and provides the most information about the person being measured. To demonstrate the procedure, the multidimensional extension of the two-parameter logistic (M2PL) model has been selected. M2PL parameter estimates are computed using item response data from a sample of 1,000 students who took the ACT Assessment Mathematics Usage Test. As a basis for comparison, a traditional item analysis and a three-parameter logistic calibration using LOGIST are performed on the same data set. It is concluded that the results are readily interpreted. (GDC)

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that Measure More than One Ability

Mark D. Reckase  
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The Difficulty of Test Items  
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Statistics that describe the characteristics of test items are routinely used to assist in the test construction process. These statistics are often used to help produce equivalent forms of tests and to produce tests according to detailed specifications. Sometimes they are used to select items that will yield test forms that measure a single trait or dimension.

Most, if not all, of the statistics that are commonly used to describe a test item assume that the item measures a single trait or dimension. The unidimensional IRT procedures (see Traub and Lam (1985) for a summary of these) make the assumption of a single trait directly, but even the traditional statistics, such as  $p$  and  $r_{bis}$ , make an implicit assumption that items can be ordered in difficulty along a single continuum and that a single dimension is being measured by the test.

Yet, most items are multidimensional in some sense, and, depending upon the strength of the multiple dimensions, these unidimensional statistics may not be appropriate. Some items measure a fairly strong first dimension with only minor higher order dimensions. Vocabulary and some mathematics computation items may fall into this category. Although there may be many different types of vocabulary words and synonyms, one overall dimension seems to be measured by vocabulary tests. For these types of items, unidimensional statistics seem reasonable.

Some items clearly require more than one distinct ability to arrive at a correct response. Mathematical "story problems" are the most common example of this type of item. Both mathematical and verbal skills are required to obtain a correct answer. Current measures of item characteristics are inappropriate for this type of item. For example, if a proportion correct difficulty index is used to describe items of this type, the ranking of items on difficulty may vary considerably if the examinee sample used to compute the statistics differs more in mathematical or verbal skill.

What is needed is a means of describing the characteristics of an item that takes the dimensionality of the skills required to solve the item into account. Such a statistic can then be used to determine how or if it is possible to compare items that measure different combinations of abilities. The statistic can be used in test construction to ensure that a test matches predefined characteristics and, if desirable, to form tests that measure a single characteristic.

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It is the goal of this paper to describe a means of determining the difficulty of an item that gives useful information when the item measures more than one dimension. This multidimensional measure of difficulty is based on a multidimensional generalization of item response theory concepts. To demonstrate the usefulness of this concept, it will be compared to several commonly used unidimensional item statistics.

### Method

The proposed definition of a measure of multidimensional item difficulty (MID) is based on three general assumptions. First, it is assumed that the probability of answering an item correctly increases monotonically with an increase in each dimension being measured. Second, it is assumed that it is desirable to locate an item at a single point in a multidimensional space. Previous work (Reckase and McKinley, 1983) defined MID as the locus of points of inflection of a multidimensional item response surface (IRS). That definition proved to be unwieldy, prompting the work presented here.

The third assumption is that the most reasonable point to use in defining the MID for an item in the multidimensional space is the point where the item is most discriminating. This is the point where the item provides the most information about the person being measured.

On the basis of these three assumptions, a general technique for determining the MID will be specified. However, in order to demonstrate the procedure, a particular model is needed. For the purpose of demonstration, the multidimensional extension of the two-parameter logistic (M2PL) model has been selected (McKinley and Reckase, 1983a). This model has been selected because it meets the first assumption and because estimation procedures are available for the parameters (McKinley and Reckase, 1983b).

The M2PL model can be described by the following equation.

$$P(x_{ij} = 1 | a_i, d_i, \theta_j) = \frac{e^{(a_i' \theta_j + d_i)}}{1 + e^{(a_i' \theta_j + d_i)}} \quad (1)$$

where  $P(x_{ij} = 1 | a_i, d_i, \theta_j)$  is the probability of a correct response to item  $i$  by person  $j$ ;  $a_i$  is a vector of discrimination parameters;  $d_i$  is a scalar parameter that is related to the difficulty of the item; and  $\theta_j$  is a vector of ability parameters. The exponent can also be expressed as

$$\sum_{k=1}^n a_{ik} (\theta_{jk} - b_{ik})$$

where  $n$  is the number of dimensions,  $a_{ik}$  is an element of  $a_i$ ,  $\theta_{jk}$  is an

element of  $\theta_j$ , and  $d_i = -\sum_{k=1}^n a_{ik} b_{ik}$ . In this form, it is more similar to the

usual expression for the two-parameter logistic model.

The M2PL model defines an IRS that is monotonically increasing in probability as the elements of  $\theta_j$  increase. An example of the IRS defined by the model for the two-dimensional case is shown in Figure 1. This IRS was generated using  $a_{i1} = .75$ ,  $a_{i2} = 1.2$ , and  $d_i = -1$ . This surface clearly meets the first assumption given above.

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Insert Figure 1 about here

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The goal of the analysis to be presented is to determine the point in the  $\theta$ -space at which the IRS has the maximum slope. However, the slope along a surface differs depending on the direction taken relative to the surface. The slope at a point can be quite different in one direction than in another. In order to develop a definition for MID that is easily useable, the slope at a point in the  $\theta$ -space will always be determined using the direction from the origin of the  $\theta$ -space to the point. This choice of a direction has implications that will be described later in the paper.

The procedure to be used to find the point of maximum slope from the origin involves two steps. First, the point of maximum slope in a particular direction will be determined. Then, the slopes in each direction will be analyzed to determine the direction that yields the overall maximum.

In order to facilitate the analysis, the model given in Equation 1 will be translated to polar coordinates. That is, each  $\theta_{jk}$  will be replaced by  $\theta_j \cos \alpha_{jk}$  where  $\theta_j$  is the distance from the origin to  $\theta_j$  and  $\alpha_{jk}$  gives the angle from the  $k$ th axis to the point. Figure 2 shows the relationship between the original and translated parameters for the two dimensional case.

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Insert Figure 2 about here

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The revised expression for the M2PL model is presented in Equation 2 below.

$$P(x_{ij} = 1 | a_i, d_i, \alpha_j, \theta_j) = \frac{e^{(\sum_{k=1}^n a_{ik} \theta_j \cos \alpha_{jk} + d_i)}}{1 + e^{(\sum_{k=1}^n a_{ik} \theta_j \cos \alpha_{jk} + d_i)}} \quad (2)$$

In order to find the maximum slope in a particular direction  $\alpha_j$ , the second derivative is taken with respect to  $\theta_j$  and solved for its zero point.

This analysis gives the point of inflection of the IRS in a particular direction. For the M2PL model, the second derivative with respect to  $\theta_j$  is given by

$$\frac{\delta^2 P(x_{ij} = 1 | a_i, d_i, \alpha_j, \theta_j)}{\delta \theta_j^2} = \left( \sum_{k=1}^n a_{ik} \cos \alpha_{jk} \right)^2 P_{ij} (1 - 3P_{ij} + 2P_{ij}^2) \quad (3)$$

where  $P_{ij} = P(x_{ij} = 1 | a_i, d_i, \alpha_j, \theta_j)$ . Equation 3 is equal to zero when  $P_{ij} = .5$ . Thus, the slope in direction  $\alpha_j$  is at its maximum when the IRS crosses the .5 plane.

The slope of the surface in direction  $\alpha_j$  is given by the first derivative of the model with respect to  $\theta_j$ ,

$$\frac{\delta P(x_{ij} = 1 | a_i, d_i, \alpha_j, \theta_j)}{\delta \theta_j} = P_{ij} (1 - P_{ij}) \sum_{k=1}^n a_{ik} \cos \alpha_{jk} \quad (4)$$

When  $P_{ij} = .5$ , the slope is equal to

$$\frac{1}{4} \sum_{k=1}^n a_{ik} \cos \alpha_{jk} \quad (5)$$

To find the direction of steepest slope, Expression 5 is differentiated with respect to  $\cos \alpha_{jk}$  and solved for zero. However, before performing the differentiation, the constraint that  $\sum_{k=1}^n \cos^2 \alpha_{jk} = 1$  is added to the expression for the slope by setting

$$\cos \alpha_{jn} = 1 - \sum_{k=1}^{n-1} \cos^2 \alpha_{jk} \quad (6)$$

The differentiation results in  $n-1$  equations of the form

$$\begin{aligned} a_{i1} - a_{in} \frac{\cos \alpha_{j1}}{\cos \alpha_{jn}} &= 0 \\ a_{i2} - a_{in} \frac{\cos \alpha_{j2}}{\cos \alpha_{jn}} &= 0 \\ &\vdots \\ a_{in-1} - a_{in} \frac{\cos \alpha_{jn-1}}{\cos \alpha_{jn}} &= 0 \end{aligned} \quad (7)$$

This system of equations can be solved for  $\cos \alpha_{ik}$  in terms of the  $a$ -parameters. The result is

$$\cos \alpha_{ik} = \frac{a_{ik}}{\sqrt{\sum_{k=1}^n a_{ik}^2}} \quad (8)$$

Up to this point, the  $\alpha_j$ -vector was considered as a person parameter that was used to convert from rectangular to polar coordinates. When the direction that yields the maximum slope is found from the system of equations given in (7),  $\alpha$  changes to an item parameter. Therefore, it will be denoted as  $\alpha_i$  in the following equations.

The above derivation gives the angle from the origin to the point of steepest slope. To determine the distance,  $D_i$ , to the point of steepest slope, Equation 8 can be substituted for  $\cos \alpha_{ik}$  in Equation 2 and the resulting equation can be solved for  $\theta_j$  for  $P_{ij} = .5$ , the value of the probability when the slope is maximum. The result is

$$D_i = \frac{-d_i}{\sqrt{\sum_{k=1}^n a_{ik}^2}} \quad (9)$$

where  $D_i$  is an item parameter.

Thus, to describe the difficulty of a multidimensional item, it is proposed that a set of statistics be reported: the distance to the point of steepest slope in a direction from the origin, and the angles, or direction cosines, needed to describe the direction.

At this point it may prove helpful to give an example of the MID for an item. The item shown in Figure 1 will be used for this purpose. The  $a$ -parameters for the item were  $a_{11} = .75$  and  $a_{12} = 1.2$ , therefore the direction cosines are  $\cos \alpha_{11} = .53$  and  $\cos \alpha_{12} = .85$  corresponding to angles of  $58^\circ$  and  $32^\circ$  with the  $\theta_1$ - and  $\theta_2$ -axes respectively. The distance in that direction to the point of steepest slope is .71. This corresponds to a  $\theta_1$ -coordinate of .37 and a  $\theta_2$ -coordinate of .60. The distance can be interpreted much like a  $b$ -parameter from unidimensional IRT, indicating that the item is fairly difficult for a population centered at the origin.

In order to demonstrate the usefulness of the MID statistics for the analysis of actual test data, the statistics were computed using the item response data from a representative sample of 1000 students who took the ACT Assessment Mathematics Usage Test in February 1983. The M2PL parameter estimates for the 40 items on the test were determined using the MAXLOG program (McKinley and Reckase, 1983b). As a basis for comparison, a traditional item analysis and a three-parameter logistic calibration using LOGIST (Wingersky, Barton, and Lord, 1982) were performed on the same data set.

### Results

The parameter estimates from the M2PL model, the directions, and the distances for a two-dimensional analysis of the ACT Mathematics Usage Test are presented in Table 1. A two-dimensional analysis was performed to keep the demonstration of the procedures simple. Also, the method for determining the appropriate number of dimensions has not been solved.

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Insert Table 1 about here

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The second, third, and fourth columns of Table 1 give the parameter estimates from the M2PL model. Notice that the lower numbered items (the easier items since the items were arranged in order of difficulty) tend to have high values for  $a_{11}$  while the more difficult items tend to have higher values of  $a_{12}$ . The values of  $d_i$  reflect the difficulty of the items.

The next two columns show the direction cosines for each item and the following two columns give the corresponding angles with the  $\theta_1$ - and  $\theta_2$ -axes. Note that the angles must add to  $90^\circ$  once the solution is in two dimensions. The angles with the axes reflect the pattern present in the  $a$ -parameter estimates. The easier items tend to be clustered along the  $\theta_1$ -axis and the harder items tend to have directions close to the  $\theta_2$ -axis.



The last column in the table gives the distance measure for each item. The distances tend to reflect the difficulty ordering of the items.

The results given in Table 1 can be presented graphically as vectors in a two-dimensional space. Such a representation is given in Figure 3. Note that the items with negative values for  $D_1$  tend to cluster along the  $\theta_1$ -axis, while the harder items are more scattered.

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Insert Figure 3 about here

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The traditional difficulty (p-value) and discrimination (biserial correlation) statistics for the forty mathematics items are given in Table 2. The statistics show that the items are arranged in approximate order of difficulty, and that they tend to have uniformly high values of discrimination.

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Insert Table 2 about here

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Table 3 presents the parameter estimates for the three-parameter logistic model. The estimates for the parameters were obtained using LOGIST 5 on the sample of 1000. The a-parameter estimates given in the table tend to be higher for the more difficult items. The b-parameter estimates roughly indicate the ordering of the items according to difficulty.

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Insert Table 3 about here

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In order to give some basis for comparison for the statistics presented, Pearson product moment correlations were computed for all possible pairs of statistics. The correlations are only reported as an aid for interpretation of the data. In some cases, nonlinear relationships are present between the coefficients so the correlations should be interpreted cautiously. The correlations for all of the statistics are presented in Table 4.

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Insert Table 4 about here

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The strongest relationships indicated in the table of correlations are those among the difficulty indices. The values of  $p$ ,  $d$ ,  $D$ , and  $b$  are all highly interrelated. All seem to be an indicator of the probability of getting an item correct for the group of individuals that were tested.

A second relationship that is of interest concerns the measures of discrimination  $r_{bis}$  and  $a$ . These two values are only correlated .14. An analysis of the scatter plot of these statistics indicated that the low correlation was the result of several difficult items that had relatively low  $r_{bis}$ -values and high  $a$ -values (i.e. Number 34). The  $r_{bis}$ -values for these items may have been deflated by guessing effects.

The direction measure for the items,  $\cos \alpha_{i1}$ , is most highly related to  $p$ , indicating that the dimension measured by the items changes with the difficulty of the item. The high correlations with the other M2PL model parameters are artifacts. For example,  $\cos \alpha_{i1}$  and  $\cos \alpha_{i2}$  are functionally related and must be highly correlated. The  $a$ -parameter from the three parameter logistic model is most highly related to  $a_2$ , indicating that LOGIST is estimating ability from the second M2PL dimension.

### Discussion

The purpose of this paper has been to present a definition for multidimensional item difficulty (MID) and to demonstrate its use for a particular set of test items. The definition presented defines MID as the direction from the origin of the multidimensional space to the point of greatest discrimination for the item and the distance to that point. For the two-dimensional case, two statistics are required to present this information, the angle with one of the axes, and the distance along the vector to the point of maximum discrimination. In  $n$ -dimensions,  $n$  statistics are needed to specify the MID:  $n-1$  angles and a distance. If desired, the coordinates of the point of greatest discrimination could be presented as an alternative definition, but specifying the  $n$ -coordinates was judged not as useful.

The direction and distance is not unique for an item any more than any other IRT-parameters are unique. They are only uniquely defined when the origin and unit of measure of the complete latent space are specified. If a different origin is used, or different units are specified, the direction and distance will change. Work is currently being done to work out the procedures for translating from one specification of a latent space to another. Such procedures will be needed to link multidimensional calibrations or for equating multidimensional tests.

The MID information for the ACT Mathematics Usage Test revealed some interesting information about the test. The easier items tended to be measures of  $\theta_1$  while the harder items tend to give more information about  $\theta_2$ .

From an analysis of items that are most highly related to  $\theta_1$ , it seems that that dimension is related to mathematics problems with a strong verbal component. For example, Item 6 is almost a pure measure of  $\theta_1$  and it clearly requires reading comprehension skills. Item 20 is a harder item of this same type (see Figure 4 for these items).

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Insert Figure 4 about here

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Item 34, on the other hand is almost a pure measure of  $\theta_2$ . Item 9 is an easier version of this type of item. Both of these items require very little verbal comprehension skill. Both of these items are also presented in Figure 4.

In order to compare the difficulty level of items, they must be measuring the same composite of abilities, that is, they must have the same direction. Thus, it is reasonable to compare the difficulties of Items 6 and 20 since they measure in the same direction, but it would not be reasonable to compare the difficulty of Items 20 and 34 because they are measuring quite different things.

Items that measure best in a particular direction can be combined together to form a test that will operate as if it were measuring a single dimension. For example, if items 4, 13, 16, 25, and 33 were selected as a single subtest, that subtest would operate as a unidimensional subtest because all of the items measure the same composite of  $\theta_1$  and  $\theta_2$ . All of these items can also be ordered in difficulty because they all measure the same composite of abilities.

The concept of MID can also be used to select items to measure a person at a particular  $\theta$ -point. The MID definition emphasizes that when selecting an item the direction must be considered as well as the point of maximum discrimination. If information is wanted for all  $\theta$ -dimensions, the item must be selected so that the slope of the IRS is nonzero for any direction parallel to a  $\theta$ -axis. This means that the item direction must not be parallel to any axis. Item 36 is such an item. If one dimension is of interest to the exclusion of the others, items should be selected that have directions parallel to the dimension of interest.

### Conclusions

The purpose of this paper has been to define a measure of item difficulty that is appropriate for items that require more than one ability to achieve a correct response. The definition that has been proposed describes the difficulty of an item as a direction and a distance in the complete latent space. The use of the definition was demonstrated using the multivariate extension of the two-parameter logistic model, but the definition is sufficiently general that it can be used with any model that yields probabilities that increase monotonically with an increase in ability on any dimension.

The definition of difficulty was applied to test data from the ACT Mathematics Usage Test and the results were shown to be readily interpreted. The uses of the statistics for test construction were also discussed.

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Table 1

Item Parameters, Directions  
and Distance for the Items  
in the ACT Assessment Mathematics Usage Test

Item Number	$a_{i1}$	$a_{i2}$	$d_i$	$\cos \alpha_{i1}$	$\cos \alpha_{i2}$	$\alpha_{i1}$	$\alpha_{i2}$	$D_i$
1	1.81	.86	1.46	.90	.43	25	65	-.73
2	1.22	.02	.17	1.00	.02	1	89	-.14
3	1.57	.36	.67	.97	.22	13	77	-.42
4	.71	.53	.44	.80	.60	37	53	-.50
5	.86	.19	.10	.98	.21	12	78	-.11
6	1.72	.18	.44	.99	.10	6	84	-.25
7	1.86	.29	.38	.99	.15	9	81	-.20
8	1.33	.34	.69	.97	.25	14	76	-.50
9	1.19	1.57	.17	.60	.80	53	37	-.09
10	2.00	.00	.38	1.00	.00	0	90	-.19
11	.87	.00	.03	1.00	.00	0	90	-.03
12	2.00	.98	.91	.90	.44	26	64	-.41
13	1.00	.89	-.49	.75	.66	42	48	.37
14	1.22	.14	.54	.99	.11	7	83	-.44
15	1.27	.47	.29	.94	.35	20	70	-.21
16	1.35	1.15	-.21	.76	.65	40	50	.12
17	1.06	.45	.08	.92	.39	23	67	-.07
18	1.92	.00	.12	1.00	.00	0	90	-.06
19	.96	.22	-.30	.97	.22	13	77	.30
20	1.20	.12	-.28	.99	.10	6	84	.23
21	1.41	.04	-.21	.99	.03	2	88	.15
22	1.54	1.79	.02	.65	.76	49	41	-.01
23	.54	.23	-.69	.92	.39	23	67	1.18
24	1.53	.48	-.83	.95	.30	17	73	.52
25	.72	.55	-.56	.79	.61	37	53	.62
26	.51	.65	-.49	.62	.79	52	38	.59
27	1.66	1.72	-.38	.69	.72	46	44	.16
28	.69	.19	-.68	.96	.27	15	75	.95
29	.88	1.12	-.91	.62	.79	52	38	.64
30	.68	1.21	-1.08	.49	.87	61	29	.78
31	.24	1.14	-.95	.21	.98	78	12	.82
32	.51	1.21	-1.00	.39	.92	67	23	.76
33	.76	.59	-.96	.79	.61	38	52	1.00
34	.01	1.94	-1.92	.01	1.00	90	0	.99
35	.39	1.77	-1.57	.22	.98	78	12	.87
36	.76	.99	-1.36	.61	.79	52	38	1.09
37	.49	1.10	-.81	.41	.91	66	24	.67
38	.29	1.10	-.99	.25	.97	75	15	.87
39	.48	1.00	-1.56	.43	.90	64	26	1.41
40	.42	.75	-1.61	.49	.87	61	29	1.87

Table 2

Proportion Correct and Biserial Correlation  
for each Item on the Mathematics Usage Test

Item Number	p	$r_{bis}$
1	.70	.48
2	.53	.41
3	.60	.48
4	.59	.36
5	.52	.37
6	.56	.49
7	.55	.51
8	.62	.44
9	.50	.49
10	.54	.54
11	.51	.32
12	.60	.53
13	.39	.49
14	.60	.41
15	.54	.47
16	.44	.55
17	.51	.44
18	.51	.49
19	.43	.39
20	.44	.43
21	.46	.45
22	.47	.56
23	.34	.29
24	.35	.53
25	.38	.40
26	.39	.35
27	.41	.60
28	.35	.33
29	.32	.49
30	.30	.45
31	.32	.33
32	.31	.40
33	.30	.42
34	.22	.32
35	.26	.41
36	.25	.47
37	.34	.40
38	.31	.34
39	.22	.39
40	.20	.31

Table 3

Estimated Item Parameters  
for the Three-Parameter Logistic Model

Item Number	a	b	c
1	1.08	-.64	.12
2	.66	.08	.12
3	.97	-.25	.12
4	.47	-.29	.12
5	.50	.18	.12
6	.93	-.10	.11
7	1.21	.05	.15
8	.77	-.35	.12
9	1.28	.34	.20
10	.97	-.17	.02
11	.40	.29	.12
12	1.44	-.22	.12
13	1.23	.72	.16
14	.65	-.29	.12
15	.91	.09	.15
16	1.36	.38	.12
17	1.09	.47	.24
18	.85	.03	.06
19	.73	.71	.16
20	.81	.58	.15
21	.84	.44	.13
22	1.46	.30	.14
23	.41	1.61	.12
24	.89	.54	.00
25	.91	1.00	.18
26	.57	.98	.13
27	1.36	.35	.06
28	.93	1.36	.22
29	1.31	.91	.12
30	1.43	1.11	.15
31	1.60	1.44	.23
32	1.34	1.22	.18
33	1.46	1.18	.17
34	2.00	1.55	.15
35	1.73	1.29	.14
36	.79	1.15	.03
37	1.05	1.15	.18
38	1.22	1.45	.20
39	1.25	1.52	.11
40	1.94	1.67	.13

Table 4

Correlation between Item Statistics  
for the Items of the ACT Assessment Mathematics Usage Test

Statistic	Statistic												
	1.	2.	3.	4.	5.	6.	7.	8.	9.	10.	11.	12.	13.
1. p		.38	.75	-.47	.99	.72	-.69	-.71	.71	-.96	-.49	-.95	-.13
2. $r_{bis}$			.80	.16	.41	.31	-.18	-.24	.24	-.49	.14	-.55	-.44
3. $a_{i1}$				-.34	.78	.71	-.66	-.70	.70	-.75	-.21	-.82	-.47
4. $a_{i2}$					-.46	-.81	.87	.86	-.86	.33	.74	.38	.16
5. $d_i$						.74	-.67	-.71	.71	-.94	-.47	-.93	-.12
6. $\cos \alpha_{i1}$							-.92	-.97	.97	-.63	-.69	-.69	-.30
7. $\cos \alpha_{i2}$								.98	-.98	.61	.65	.65	.30
8. $\alpha_{i1}$									-1.00	.63	.69	.68	.31
9. $\alpha_{i2}$										-.63	-.69	-.68	-.31
10. $D_i$											.42	.96	.15
11. a												.40	.24
12. b													.34
13. c													



Figure 1

Item Response Surface  
for the M2PL Model  
with Parameters  $a_1 = .75$ ,  $a_2 = 1.2$ ,  $d = -1$

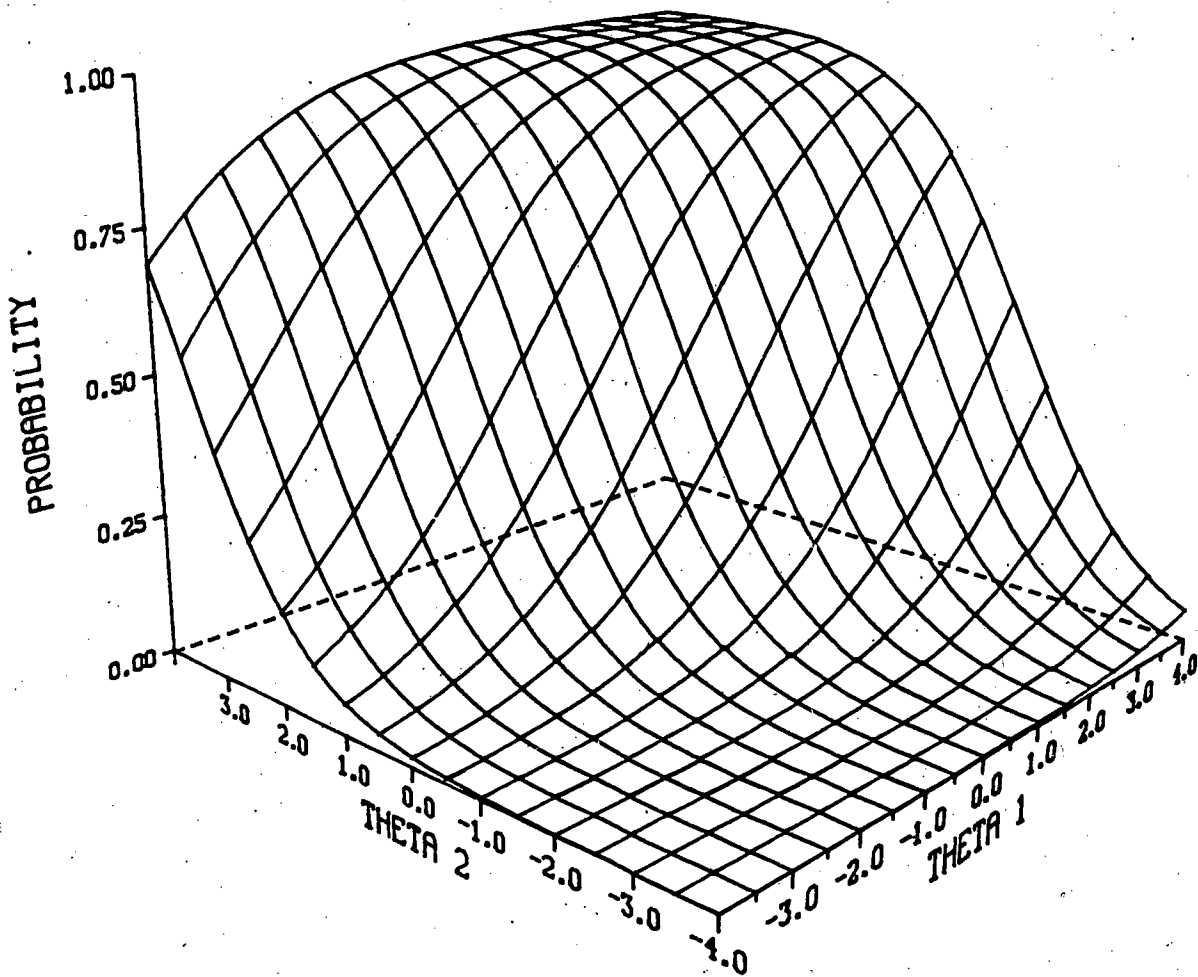


Figure 2

Conversion of M2PL Parameters  
to Polar Coordinates

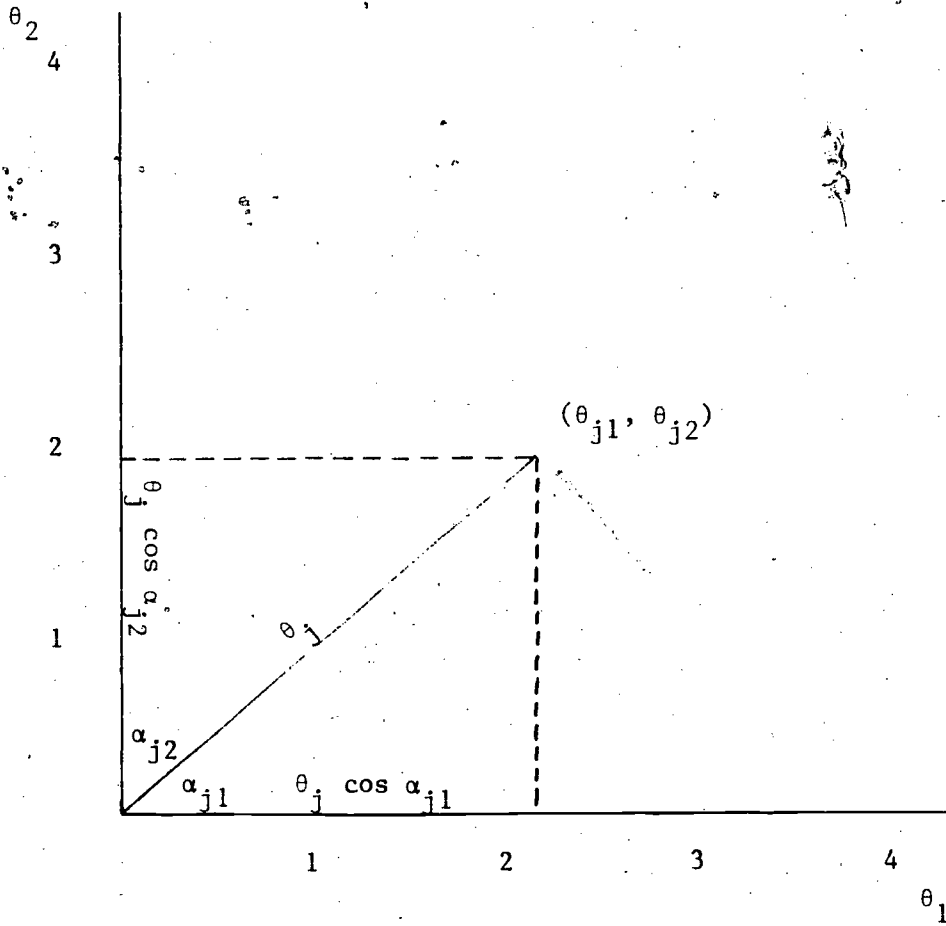


Figure 3.

Representation of the ACT Mathematics Usage  
Items as a Direction and a Distance

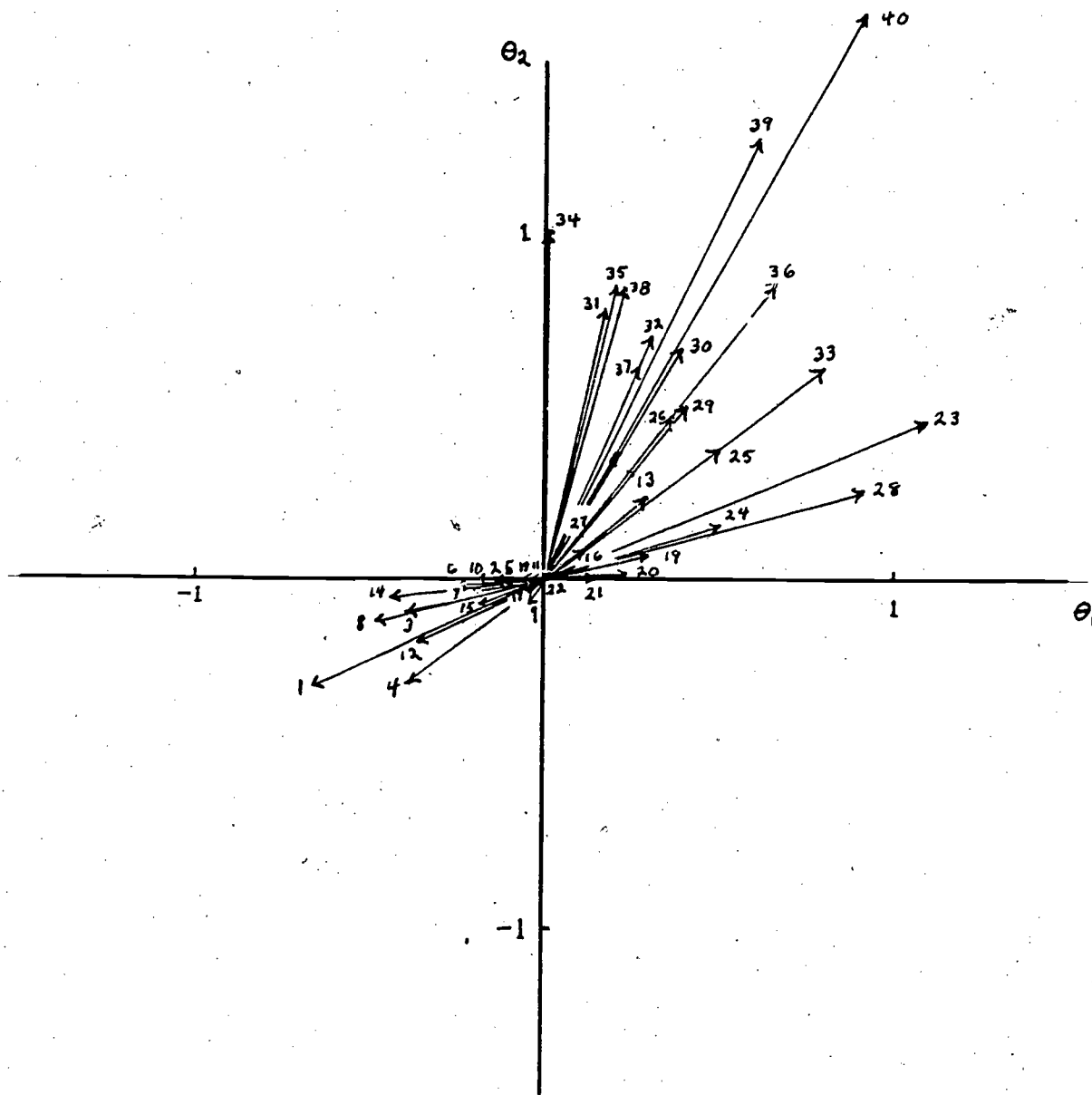


Figure 4

Items Measuring Each of the  
Two Dimensions

$\theta_1$

6. Sheila's salary is \$110 per day. Due to financial problems in her company, her employer has asked Sheila to take a 10% cut in pay. How much will Sheila be earning per day if she takes the cut in pay?

F. \$ 11  
G. \$ 99  
H. \$100  
J. \$109  
K. \$121

20. A serving of a certain cereal, with milk, provides 35% of the potassium required daily by the average adult. If a serving of this cereal with milk contains 112 milligrams of potassium, how many milligrams of potassium does the average adult require each day?

F. 35  
G. 39  
H. 147  
J. 320  
K. 392

$\theta_2$

9.  $|-5| + |6| + (-5) + 6 = ?$

A. -22  
B. -10  
C. 2  
D. 10  
E. 12

34. Which line is parallel to  $y = 3x + 1$  and intersects  $y = 6x - 1$  on the  $y$ -axis?

F.  $y = 3x - 1$   
G.  $y = 2x - 1$   
H.  $y = \frac{1}{3}x - 1$   
J.  $y = \frac{1}{3}x + 1$   
K.  $y = \frac{1}{2}x - 1$