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ABSTRACT

This study presents data from the 1979 Income Survey Development Program (ISDP) on lifetime work interruptions and examines the relationship between work interruptions and earnings. Descriptive data showing the extent to which men and women have experienced work interruptions are presented, followed by an analysis of the impact of work interruptions on earnings. The tables in the report show the proportion of persons with work interruptions and the proportion of potential work years spent away from work for the reasons surveyed. Work interruptions were reported by sex; by race and Spanish origin; by years of school completed; by occupation; and by life-cycle status. The data were gathered by a panel survey of approximately 9,000 households that were visited at three-month intervals over a period of a year and a half beginning in February, 1979. The data were then inflated to represent the civilian noninstitutional population of the United States and members of the armed forces living off post or with their families on post. The data presented cover persons 21 to 64 years of age who ever worked. (The study concludes that work interruptions explain only a small proportion of the earnings differential between men and women.) (KC)

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Lifetime Work Experience and Its Effect on Earnings:

Retrospective Data From the 1979 Income Survey Development Program

by Joseph J. Salvo and John M. McNeil

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CONTENTS

		Page
Intro	oduction	1
Worl	k Interruptions	1
	By Sex	1
	By Race and Spanish Origin	1
	By Years of School Completed	2
	By Occupation	2
	By Life-Cycle Status	2
Worl	k Interruptions and Earnings Differentials	3
	Previous Research	3
	ISDP Data	3
Сопс	clusions	5
Sele	cted References	5
	TEXT TABLES	
А. В.	Work Interruption History, by Race, Spanish Drigin, and Selected Characteristics: Males . Work Interruption History, by Race, Spanish Drigin, and Selected Characteristics:	6
_	Females	7
C.	Mean Duration of Work Interruptions, by Reason and Sex	7
D.	Mean Values of Regression Variables	8
E. F.	Coefficients of Regression of Log of Hourly Earnings on Specified Explanatory Variables . Standardized Coefficients of Regression of Log of Hourly Earnings on Specified	8
Ģ.	Explanatory Variables	9
	Education Variables	9
	DETAILED TABLES	
1. 2.	Lifetime Work Interruption Status, by Reason, Employment Status, and Sex Proportion of Potential Work Years Spent Away From Work, by Employment Status	10
	and Sex	10
3A.	Lifetime Work Interruption Status, by Reason and Selected Characteristics: Males	11
3B.	Lifetime Work Interruption Status, by Reason and Selected Characteristics: Females	12
4A.	Lifetime Work Interruption Status, by Reason and Selected Characteristics: Employed Males	13
4B.		14
5A.	Proportion of Potential Work Years Spent Away From Work, by Selected Character-	
5B.	istics: Males Proportion of Potential Work Years Spent Away From Work, by Selected Character-	15
6A.	istics: Females	16
6B.	istics: Employed Males	17
ŲŮ,	istics: Employed Females	18



DETAILED TABLES - Continued

7A.	Mean Percent of Potential Work Years Spant Away From Work, by Reason and Selected Characteristics: Persons Who Ever Worked	15
78.	Mean Percent of Potential Work Years Spent Away From Work, by Reason and Selected Characteristics: Employed Persons	20
7¢.	Mean Percent of Potential Work Years Spent Away From Work, by Reason and Selected Characteristics: Full-Time Workers	2
	APPENDIXES	
A. B.	Definitions and Explanations Source and Reliability of the Estimates Facsimile of 1979 ISDP Questions on Work Interruptions	2: 2: 2:



Lifetime Work Experience and Its Effect on Earnings: Retrospective Data From the 1979 Income Survey Development Program

INTRODUCTION

The extent to which persons remain attached to the labor force over the course of their working-age years has important economic and social implications. Differences in labor force attachment between men and women has been cited as one major reason why women earn less than men. This study presents data from the 1979 Income Survey Development Program (ISDP) on lifetime work interruptions and examines the relationship between work interruptions and earnings. Descriptive data showing the extent to which men and women have experienced work interruptions are presented, followed by an analysis of the impact of work interruptions on earnings. The study concludes that work interruptions explain only a small proportion of the earnings differential between men and women.

The 1979 ISDP was a panel survey of approximately 9,000 households that were visited at 3-month intervals over a period of a year and a half beginning in February 1979. The survey. part of the development stage of the new income survey called the Survey of Income and Program Participation (SIPP), was a joint effort of the U.S. Department of Health and Human Services and the U.S. Bureau of the Census. The third-wave questionnaire contained a section on personal history and within that section were questions on lifetime work interruptions. The questions (reproduced in app. C) asked whether the person had ever been away from work for 6 months or longer for each of three reasons: (1) because he or she was not able to find work. (2) because he or she was taking care of home or family, and (3) because he or she was ill or disabled. Beginning and ending dates were recorded for each interruption. A maximum of four interruption periods were identified for each of the three possible reasons for interrupting.

The tables in this report show the proportion of persons with work interruptions and the proportion of potential work years spent away from work for the reasons surveyed. The data presented in this report have been inflated to represent the civilian noninstitutional population of the United States and those members of the Armed Forces living off post or with their families on post. The data presented cover persons 21 to 64 years of age who ever worked. For explanations and definitions of terms, see appendix A.

WORK INTERRUPTIONS

By Sex

There were very large differences between the sexes in lifetime work attachment patterns. Tables A and B show that approximately 26 percent of males had experienced a work interruption of 6 months or more compared to about 72 percent of

females. The mean proportion of potential work years spent away from work was about 3 percent for males and about 31 percent for females.

The disparity in work interruptions between the sexes is primarily the result of the frequency with which women experience familial interruptions. Approximately two-thirds of the women in the universe had familial interruptions compared to less than 2 percent of the men. The effect of this difference on lifetime work experience is heightened by the fact that interruptions for familial reasons tend to be of greater duration than interruptions for other reasons. (See table C.)

There was a small difference between men and women in the proportion who had experienced interruptions due to an inability to find work, approximately 17 percent for men compared to about 14 percent for women. There was not a statistically significant difference in the proportions with disability interruptions,

By Race and Spanish Origin

Black women showed a stronger lifetime attachment to the labor force than did White women. This result held true even though Black women were more likely than White women to have experienced interruptions due to disability or illness or an inability to find work. The reason for the stronger attachment of Black women was due to their relatively low rate of interruption for familial reasons; approximately 44 percent of Black women had one or more familial interruptions compared to about 67 percent of White women. The mean proportion of potential work years spent away from work was about 18 percent for Black women and about 33 percent for White women. A plausible reason for the greater labor force attachment of Black women would seem to be that they have less of an economic option than White women to interrupt work for lengthy periods of time.

Spanish-origin women had overall rates of labor force attachment similar to those of White women. The proportion of Spanish-origin women with one or more interruptions was 75 percent and Spanish-origin women had spent about 28 percent of their potential work years away from work. The comparable figures for White women were not statistically different. There was a marked difference between the two groups in the rate at which they had experienced interruptions due to inability to find work; about 24 percent of women of Spanish origin had experienced such interruptions compared to about 12 percent of White women.

Compared to Black women, women of Spanish origin spent a greater proportion of their potential work years away from work. The gap is the result of a very notable disparity in the proportion who had experienced one or more familial interruptions, about 62 percent for Spanish-origin women compared to about 44 percent for Black women.



1

Black man had a higher overall interruption rate than White men and spent a higher proportion of their botential work years away from work. The percent with interruptions caused by an inability to find work was about 15 percent among White men and 35 percent among Black men. Overall, Black men spent about 7 percent of their potential work years away from work compared to about 3 Percent for White men. The proportion of time spent away from work by Spanish-origin men was not statistically different than the figure for White men.

By Years of School Completed

Education has long been viewed as an important determinant of labor force attachment. The data confirm that, for both men and women, higher levels of education are associated with fewer work interruptions.

Among both men and women, the proportion of persons experiencing work interruptions due to inability to find work decreased with increasing education. The proportion of men who experienced such interruptions ranged from approximately 25 percent among those who did not graduate from high school to about 8 percent among college graduates. A similar pattern was evident for women.

The picture for interruptions due to disability was largely the same. Among both men and women, those with less than a high school education were much more likely to have had an interruption than were college graduates. About one-fifth of the former experienced an interruption compared to 2.5 percent of the latter. Among women, the data indicate the existence of a relationship between level of education and likelihood of having had an interruption for familial reasons, but the key distinction lies between those who have and those who have not graduated from college. Approximately two thirds of women with less than a college education had interruptions for familial reasons compared to about half of those who graduated from college. On average, women college graduates spent about 24 percent of their potential work years away from work compared to 34 percent for those with less than 12 years of schooling. Apparently, the high level of labor force commitment necessary to deter familial interruptions is more salient among college graduates. In general, the relationships observed between work interruptions and education for all ages held up within the individual age groups. For example, among men 21 to 29 years of age, the proportion who had experienced interruptions due to their inability to find work ranged from 35.5 percent among those who had not completed high school to 5.5 percent among college graduates.

Among women, the relationships among age, education, and familial work interruptions are somewhat complex. For the youngest age group, the Proportion who experienced familial interruptions is inversely related to education. This result reflects the choice of women with greater amounts of schooling to postpone childbearing. Among women in the 45- to 64-year age group, women who did not finish high school were less likely to have experienced familial interruptions than were those women who were high school graduates. A plausible explanation for this result is that women with less than a high school

education have less of an economic option to interrupt than do women with a high school education.

By Occupation

Because of the relationship noted above between years of schooling and work interruptions, it would be expected that women in occupations with relatively high educational requirements would have fewer lifetime work interruptions than women in other occupations. There is evidence that such is the case. Among women in white-collar occupations, those who were professional, technical, or kindred workers were somewhat less likely to have interrupted work for familial reasons than were those who were sales or clerical workers.

It should be noted that decisions regarding years of schooling, occupation, and familial interruptions are not independent. Women who expect to be out of the labor force for significant periods of time during their working-age years are likely to make different decisions regarding schooling and occupation than women who expect to minimize labor force interruptions.

For each of the occupation groups, wemen were much more likely than men to have experienced work interruptions and women spent a far greater Proportion of their potential work years away from work. For example, among professional, technical, and managerial workers, the overall interruption rate was 14.7 percent for men and 61 percent for women. Among the same group, men had spent approximately 2 percent of their potential work years away from work while the comparable figure for women was about 24 percent.

By Life-Cycle Status

Tables A and B present interruption data by life-cycle status. The variables chosen to describe life-cycle status included age and whether the person was ever married. Women under 45 years of age were further classified by whether they had children living in the household. The latter variable is not the preferred variable for measuring the impact of childbearing on work patterns, but information on children ever born was not available for this study.

There was an association between age and interruption history for both men and women. Approximately 53 percent of the women under 30 years had one or more work interruptions compared to about 80 percent of the women 30 years and over. The overall relationship was due to the positive association of both familial- and disability-related interruptions with age. Approximately 43 percent of women under 30 years had interruptions for familial reasons compared to about 73 percent of women 30 years and over. The proportion of women with interruptions due to disability was 3.5 percent among women under 30 years. 6.6 percent among women 30 to 44 years, and 16.1 percent among wome. 45 to 64 years.

The positive association between age and disability-related interruptions was also evident in the data for men. The proportion with one or more interruptions due to disability was 3.4 percent among men 21 to 29 years, 8.5 percent among men 30 to 44 years, and 18.2 percent among men 45 to 64 years.

The data in table 8 indicate that both marriage and child-



bearing have prominent effects on the work history patterns of women. Within each age group, ever-married women had spent a much larger proportion of time away from work than never-married women, and women with children had generally spent a larger proportion of their time away from work than women with no children.

Within the 21- to 29-year and 30- to 44-year age groups, women who never married and who had no children had stronger lifetime work attachments than other women. Among women 21 to 29 years of age, the overall rate of interruption ranged from about 21 percent for never-married women with no children to about 81 percent for ever-married women with children. The former group of women had spent only about 6 percent of their potential work years away from work compared to about 34 percent for the latter group. Among women 30 to 44 years of age, the interruption rate ranged from about 33 percent for nevermarried women without children to about 85 percent for evermarried women with children.

WORK INTERRUPTIONS AND EARNINGS DIFFERENTIALS

A major reason for the interest in data on lifetime work experience is the desire to use such data in the analysis of malefemale earnings differentials. The tenets of human capital research have traditionally stressed the importance of work experience patterns as a determinant of earnings. The descriptive data presented in the first part of this report confirm that the lifetime labor force attachment of women is weaker than that of men. Because of interruptions for familial reasons, women have a much higher overall rate of work interruptions than men and they spend a much higher proportion of their potential work years out of the labor force. Such findings have led at least some social scientists to posit that traditional familial responsibilities are one major reason why women earn less than men. This section will describe selected studies of the relationship between work interruptions and earnings and will present an analysis based on the 1979 ISDP data.

Previous Research

A major constraint in early efforts to examine the relationship between experience and earnings was the lack of data on lifetime work experience. More recently, however, a number of studies have been published which exploit the important data which has been made available from the National Longitudinal Surveys of Labor Market Experience (NLS) and the Michigan Panel Survey of Income Dynamics (PSID).

Suter and Miller (1973) were among the first to analyze the retrospective work history data from the NLS. They studied a cohort of women who were 30 to 44 years of age in 1967. Work experience was based on a question which asked about the total number of years in which the person had worked at least 6 months. Suter and Miller concluded that there was a close association between earnings and length of work experience among this cohort of women.

Mincer and Polachek (1974) extended the analysis of the NLS

retrospective data. They specified two reasons why discontinuous work history patterns might lead to lower earnings. First, interruptions in market work lead to lower levels of accumulated human capital. Second, interruptions cause a depreciation in existing human capital. That is, time spent away from market work has a cost beyond the effect of foregone experience. In their analysis, Mincer and Polachek found that the amount of time spent at home had a negative impact on earnings even when experience was also included in the earnings equation. They concluded from their analysis that a depreciation effect does, in fact, exist.

This finding was challenged by Sandell and Shapiro (1978) on the grounds that the NLS data used by Mincer and Polachek were subject to various coding errors. They replicated certain of the Mincer Polachek research using a corrected NLS file and concluded that the original study had overestimated the depreciation effect.

Corcoran (1979) conducted an analysis of the effect of experience and interruptions on earnings using retrospective data from the PSID. One of the major advantages of the PSID data set was that the sample, in contrast to the NLS samples, was representative of the female population 18 to 64 years of age. Corcoran found very little evidence of a depreciation effect. There was no effect for White women and only a minor effect for Black women. Corcoran also argued that restricting the analysis group to women 30 to 44 years of age is likely to overestimate depreciation because many women in this group have recently reentered the labor market and are likely to be affected by misinformation about job opportunities.

More recently, Mincer and Ofek (1982) used NLS data for 30-to 44-year-old married women to reaffirm the depreciation hypothesis. In an analysis of longitudinal (rather than retrospective) data from the NLS, they found that reentry wage rates were lower than wage rates at the time of labor force withdrawal. Furthermore, longer interruptions carried greater wage penalties. They also found, however, that wage rates tended to grow rapidly upon return to work. The observed amount of depreciation, they concluded, is dependent upon the length of the interruption and the length of time spent back in the labor force.

ISDP Data

The effect of work interruptions on earnings was examined by using the data described earlier to construct variables representing interruptions and experience. These variables were included in regressions which related hourly earnings to a set of explanatory variables. The universe for this part of the study consisted of all persons 21 to 64 years of age with wage and salary income during the quarter preceding the interview. Separate regressions were run for men and women, with the log of hourly earnings as the dependent variable.

The interruption and experience variables used in the regressions include the following:



^{&#}x27;Hourly earnings were calculated by dividing total earnings for the 3-month period by the total number of hours worked.

FT

UNEMP I if person had ever experienced an interruption due to an inability to find a job; O otherwise.

OISAB = 1 if person had ever experienced an interruption due to illness or disability: O otherwise.

TIME-AWAY = Duration of all interruptions² as proportion of potential work years.³

EXPER = Number of potential work years minus duration of all interruptions.4

EXPERSQ = The square of EXPER

= 1 if the jobs the person has worked at have usually or always been full-time jobs; O otherwise.

The interruption variables were specified in the above form because it was hypothesized that earnings could be affected by the existence of an interruption as well is by the length of an interruption. Because interruptions due to unemployment or disability had a relatively small effect on the proportion of potential work years spent away from work, they were entered as zero-one dummy variables. Because interruptions for familial reasons had a very strong effect on the amount of time spent away from work, they were allowed to enter the equation through their effect on the TIME-AWAY variable. The general experience variable, EXPER, was entered in its own form as well as in its squared form, EXPERSQ. The inclusion of the squared form was intended to capture the nonlinear effect of experience on earnings. (The returns to experience tend to flatten after some

The education variables included in the regression were designed to take advantage of the ISDP personal history questions on highest degree obtained, vocational training, and types of courses taken in high school. They included the following:

EDUC1 = With an advanced degree

EDUC2 = With a bachelor degree

EDUC3 = High school graduate (reference group)

EDUC4 = Not a high school graduate, with a vocational training certificate

EOUC5 = Not a high school graduate, no vocational training certificate

COURSES = Number of selected academic courses com-

Finally, a set of variables representing marital history were included:

pleted in high school

MARR1 = Married, no marital disruption (reference group)

MARR2 = With a marital disruption (ever widowed.

divorced or separated)

Never married

'A maximum of four interruption periods could be identified for each of three possible reasons for interrupting.

³Potential work years were defined as age minus years of school completed minus 6.

The means for all variables are presented in table O and the regression results are shown in tables E and F. The unstandardized regression coefficients (table E) represent the earnings return to variables included in the equations. The standardized coefficients shown in table F are computed on values of dependent and independent variables which have been standardized so that each variable has a standard deviation equal to 1. This technique makes it possible to use the size of the coefficient as the basis for comparing the relative importance of each of the variables in a given equation. Results are shown for White women and men as well as for all women and men in order to facilitate comparisons with previous studies. Results are also shown for men and women 30 years of age and over with no familial interruptions as an alternative method of examining the influence of work interruptions.

The large differences between the sexes in the degree of work attachment are highly visible in table D. Men had, on average, about 19 years of work experience and had spent only about 2 percent of their potential work years away from work. Women, on the other hand, had 14 years of work experience and had spent about 20 percent of their potential work years away from work. There were small or insignificant differences between men and women in the mean values of the other experience and interruption variables and in the mean values of most of the education and marital history variables. Men, however, were more likely than women to have received advanced degrees and a larger proportion of women than men experienced marital disruptions. The average hourly earnings of all women was \$4.38, about 63 percent as high as the average hourly earnings of \$6.92 for all men.

The regression results confirm the importance of experience as a determinant of earnings. The general experience variables EXPER and EXPERSQ are highly significant for both men and women (table E) and are important relative to other variables in the determination of hourly earnings (table F). Attachment to full-time work also has a significant effect on earnings. The coefficients of the experience variables show that the returns to experience are greater for men than for women.

The interruption variables, in general, have a negative effect on earnings, but the effect is not particularly strong or consistent. The coefficient of TIME-AWAY is significant for both men and women in the equation for persons of all races, but is significant for women only in the equations for White men and women. Interruptions due to illness or disability (OISAB) have a significant negative effect on earnings in five of the equations, but interruptions due to inability to find work have a significant negative effect in only two of the equations.

That an earnings equation contains both experience and interruption variables that are significant is evidence that a depreciation effect does exist. In the equation for men of all races, the experience variables EXPER, EXPERSQ, and FT are highly significant and the interruption variables UNEMP and OISAB are also significant. In the equations for women of all races, the experience variables and the interruption variable TIME-AWAY have highly significant effects on earnings. The conclusion is that a depreciation effect does exist and information about work interruptions will improve those models which attempt to explain earnings.



MARR3

^{*}The ISDP data on employer-specific or job-specific measures of work experience (e.g., tenure with most recent employer/at most recent job) were collected in the fifth wave of the survey and were not available for this study.

The chefficients of the education and maintal history variables are in line with expectations, but two findings should be noted. First, the coefficient of EDUC4 for men is less negative than the coefficient of EDUC5. This finding suggests that a vocational training certificate has a positive effect on earnings. Second, the coefficient of COURSES is highly significant even though other measures of educational attainment are also present in the equation. That is, for the purpose of explaining earnings, it is important to know about the types of courses taken in high school even when we already have information about years of school completed and highest degree obtained.

Table O shows that the mean earnings of women are only about 62 percent of the earnings of men even when the group under study is comprised of persons 30 years of age and over with no familial interruptions. This differential exists even though women in this universe have approximately the same mean years of experience as men. Table E shows why the large differential exists even when the mean values of experience are so close. Among the men in this group, the coefficient of EXPER is highly significant, but among women, the coefficient is not significant.

In general, the standardized regression coefficients (table F) reveal that the work interruption variables are less important than either the general experience or education variables as determinants of earnings. This holds true for both men and women. So, while the work interruption variables do show that a depreciation effect exists, general work experience and education are more critical determinants of earnings.

The earnings equations which have been developed for this report can be used to examine the extent to which differences in work history (experience and interruptions) are related to the earnings gap between men and women. That is, given the coefficients of their own equation, what would the earnings of women be if they had the same mean values as men for the variables measuring experience and interruptions. Table G shows that the effect of assigning to women the mean experience and interruption values of men is to reduce the earnings gap by only 12 percent.

CONCLUSIONS

One purpose of this report is to offer a measure of the often cited difference between men and women in the extent to which work interruptions occur. The survey data show that the difference between men and women in the likelihood of having work interruptions is very great and is apparent even within various schooling and occupational categories.

A second purpose of this report is to examine the relationship between work history and earnings and determine the extent to which differences in work history account for the difference in earnings between men and women. The earnings equations which were constructed with the survey data were found to imply that differences between men and women in general experience and in the extent of work interruptions explain little of the difference in earnings.

Does the fact that differences in work histories explain little of the earnings gap imply that the large remaining gap is due to labor market or societal discrimination? A prudent reply is that the remaining gap may be due to the effect of a number of factors. One factor that has been suggested as a partial reason for the earnings gap is the possibility that some women may choose relatively low-paying jobs if those jobs allow greater flexibility in the carrying out of family-related activities. Other factors which may contribute to the residual gap include variables which are missing from the equation le.g., time spent with same employer, participation in company training programs) or variables which may be poorly measured (e.g., response errors in the reporting of work interruptional. The data available for this study cannot be used to determine the relative importance of such factors versus the factor of discrimination. The data do indicate, however, that the earnings gap between men and women cannot be accounted for by such productivity-related variables as education, general work experience, and work interruptions.

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Table A. Work Interruption History, by Race, Spanish Origin, and Selected Characteristics: Males

Characteristic			est will one or me sting 6 mouths or			Mean percent of work years spen work for reason	t away from
	jotal numbre (thousands)	All reasons surveyed	Inability to fini work	Family reasons	Hillness or disability	Value	5] andard >- rror
Motes 21 to 65 years of age who ever worked	55,828	15.7	17. 1	1.5	10.6	1, 1	9.1
RACE AND SPANISH ORIGIN ¹				Ţ		ļ	
white Binck. Spanish origin.	49,381 5,627 3,220	.4.2 90.2 34.9	15.2 35.0 22.7	1 2 3.9 1.7	10.7 10.7 15.8	+ 9 6 9 6 1	. t . 3 . 3
YEARS OF SCHOOL COMPLETED							
Less than 12	14,171 29,761 11,896	40.1 74.7 11.0	24.9 17.1 7.0	1.9 1.3 1.6	20.3 9.3 7.4	4.7 1.1 1.1	.1
AGE BY YIARS OF SCHOOL COMPLETED							
21 to 19 years	16,048 2,314 10,104 3,640	91, 5 40, 7 20, 8 6, 9	18 0 18.5 18.5 5.5	1.4 2.2 1.4 .6	4.4 7.4 : 4.2 : 1.2 :	4,5 0.8 5,7 2.5	? . 5 ? 4
10 to 46 years	19,106 3,899 10,728 5,019	23.4 30.7 25.2 9.6	16.2) 24.5 ; 17.8 ; 6.5 ;	1-0 7-1 1-5 1-5	8,5 18.7 6.9 7.0	7.6 5.1 7.8 .0	.1 4
45 to 56 years	.20,674 8,049 9,178 1,247	11.9 41.6 28.5 17.8	17.7 27.1 15.6 12.7	1.5 1.7 .6 5.0	18.2 25.1 17.1 4.3	3,6 4,4 7,1 2,1	.1 .2 .2
OCCUPATION GROUP OF UNITAL 108							
Professional, technical, or managerial. Sales or elected. Urafismen. Operatives. Laborers. Service.	15,040 6,621 12,845 10,254 5,832 1,457	14.7 20.0 28.8 37.5 37.9 25.9	10.7 13.8 18.7 20.8 27.6 14.8	1,7 _9 1,6 _,8 _,2 _,3	9, 1 7, 6 13, 5 14, 9 13, 6 11, 1	7, 3 7, 3 2, 9 3, 9 5, 6 4, L	
LIFE-CYCLE STATES							
21 to 29 years	16,048 9,245 6,801	20.9 19.4 22.1	18.0 17.1 19.2	1.4 1.0 1.8	6,5 3,1 1,7	4.5 7,4 6,6	. 2 . 2
30 to 44 years	19,106 17,825 1,781	: 3, 5 : : 1 41.0	1677 1478 3474	170 1.7 -1	8,5 8,3 11,0	2.6 2.5 5.9	1 1 7
45 to be wears Ever married	20,674 19,882 79,7	11.9 30.9 57.5	17.7 17.2 29.1	1.5 1.4 4.1	18.7 17.4 17.0	1,0 1,8 2,1	. l . 1 . 7

 $^{^{2}\,\}mathrm{persons}$ of Spanish origin may be of asy rase.



Table B. Work Interruption History, by Race, Spanish Origin, and Selected Characteristics: Fernales

and a Serie Co.			nt with one or Sting b wonths o		L	Mess perical o work years spea work for reason	d oaky from
	(atal number (thousands)	112 reasons surveyed	loanility to find sork	Family reasons	illnes or illsautity) white	Standard error
ite in the read of the Storages			,				
· · · · · · · · · · · · · · · · · · ·	57,758	71.9	15.2	۰۱ <u>۱</u>	4.2	177,4	0.2
GOVERNMENT OF A STREET	ļ		Į		Į		
0.70	49,817	73.0	15.4	nn.8	H. 1	42.7 11.6	.?
sports right	6 40.5 1.01a	n3.1 25.0	23.6	41.8 02.44	11.5) 19	27.6	.4 .6
y - Maria and Committee in the	1	1	Ì	ĺ			
74.8 (15.00 f	13.7au	79.5	21.7	n8,5	20.1	11.5	. 1
1	14,80 ; 8,713	73.1 54.3	17.7 8.6	64. t	6.6	17.5 24.2	
YOURS AND ON SCHOOL COMPRESSED							
3 1*	16,804	53.1	17.11	47.5	9.5		. 3
The three trees are a second to the second trees and the second trees are the second trees ar	1.948	20.6	23.1	61.7	5.9	3-1.9	. 6
[+0]	11.650 3,206	56.8 24.1	18.4 8.3	94.9 22.2 (5.1 .1 (22.1	. 5
	19,445	27.5	12.3	12.1	6.5	34.3	. 3
The the Commence	4,000	79.8	20.4	73.6	17.3	14	٠ó٠
I to I to a control to the control t	12, 366 3,018	79.6 65.1	9.9	75.1 58.5	5.7 2.9	34.8 12.7	.3
. On the graphers	21.011	81.7	13.8	73.8	16.1	15.8	.3
1- 5 - 6 to 1 - 1 - 1 - 1 - 1 - 1 - 1 - 1 - 1 - 1	7,233	81.5 83.7	22.0 }	67.6 29.6	.97.8 10.3	33.7	.4
27 GH 05 (F.).	2,489	71.8	5.6	AU, 16	5.4	34.3	. 8
occupation clouds or usual ton	i		ļ				
Soties ware, technical, or imagerial.	11.723	21.0	9.5	55.4	5.4	24.4	- }
STREET OF STREET OF STREET OF STREET	21.782 8,447	75.2 74.8	10.7	69.4 62.9	5.9 (14.5	31.8	.4
* (Coope 20)	950 (0.543	18.1 74.2	21.9 19.7	67.8	15-1	39.7	1.2
# # # # # # # # # # # # # # # # # # #	(4.54)	′4	17.7	b).d	16.5	32.1	•
į	16,505	53.1	17.0	42.5	9.	.0.2	•
71 for Cycars	ł		i	ì	3.5	-	. 3
with this leaves	11.529	65.4 81.1 i	10.3	58.3 76.1	3.4 4.0 {	26.1	. 1
No. 1 boding to a second secon	1,250	13.0	16.6	21.3	En]	9.8	. 5
Server reported processing and a server and	5.274	26.1	18.81	81	1.9	9.1	.4
Signal of the Hillerian Commence of the Commen	1.218 4.05e	43.5 20.8	311,11 15.4	22.0 L 1.9 j	1 4.1	19.4	1.H .4
C1 + (3 - 10 +) + + + + + + + + + + + + + + + + +	19,445	11,5	12,3	72.3	6,6	34.3	.3
sampled	18.)58	80.0	11.3	16.5	4, 2	35.8	
AECC HE MEGALLIAN AND ALCOHOLDS	\$6.255	84.5	12.0	81.7	6,0	38,9	.1
See and Othern Conservation Conservation (Conservation)	4, (0)	64.2	4.1	46.4	8.8	.55.1	.6
Server mana terration and a construction	1.687	35.8	29.3	_7.6 }	6.1	7.4	.7
Anthropic of others and a constraint and the constr	193 894	46.9 33.4	38.7 27.3	32.7	8.5 5.6	12.2 6.4	0,' B.
- 1 - 6- ve-rs	21.011	81.7	11,8	73,8	16.4	15.H	. 3
1 - e esceleda	19.945	83.8	13.2	26.4	15.8	37 H	3
Serve workled	1.965	42.3	34.4	15.0	23.00	13,6	1.0

the consent apertsh origin may be of any race.

Table C. Mean Duration of Work Interruptions, by Reason and Sex

Mean duration shown in years)

Sex	tanbility to		
Males		3.7 9.5	1.1 6.3



Table D. Mean Values of Regression Variables

vortante	Mulea	Pemales	white males	White Cemales	Males 30 and over with no familial in- terruptions	Founder 30 and over with no familial in- terruptions
Interruption and experience		_				
variables:	1					
CNEMP	p. 135	0.117	0.121	2,098	0.137	0.117
PISML	.060	.049	.062	.040	.078	.093
: 11111111:	.017	.203	.015	.220	(7)	(3)
PAPERCALLE	19.256	14.334	19.337	13.868	25.223	24.644
FX91880	535.421	331.227	537.208	312.682	755.661	745.445
rt	.969	.858	.968	.653	.990	.928
Iducation cartantos:					1	
1 013 1	.069	.044	.071	.045	.080	.076
FUCT 2	. 161	.128	.166	-135	.142	. 107
F143 4	.019	.015	.017	-009	.023	.020
F14 5 5	. 204	-157	. 19 [. 145	. 243	.274
counsts	2.170	2.129	2.202	2.167	2.059	1.843
Marital history cariables:						
MARING	.215	.303	. 204	.277	.155	. 388
MARRO	. 154	. 172	. 144	.159	.057	.203
Carottuss variobles:						
Mourly carnings	76.92	*4.38	?7.14	14.42	₹7 5 8	*4.71
top of boarty cardines	1.934	1.478	1.966	1.487	2.026	1.550
Pascighted A	3.157	2,416	2.854	2,101	2,145	593

V Not aboth able.

Table E. Coefficients of Regression of Log of Hourly Earnings on Specified Explanatory Variables

181 and ard	organs	in	Darrett Chasses)

Vartada	imp toyed males	Amp Loved Fema lus	Employed white males	Employed White females	Mates 30 and over with no familial in- terruptions	Females 30 and over with no functial in- terruptions
CSEAD	-0.039 (810.)	(1.002 (.018)	-0.028 (.021)	0.002 (.021)	-0.078 (.018)	0.014
nts/m	125	040	144	- (088	143	~. 183
	(.023)	(.028)	(.025)	(,032)	(.025)	(.0643
((u) - \h.\.\.	+.312 (.122)	128 (.025)	068 (.145)	155 (-028)	(x)	(x)
# NP# IC	.03515	.0:278	.03791	.02495	,03382	-00937
	(.00175)	(.00164)	(.00189)	(.002003	(.00306)	(100600)
UN14 RSO	00058	00042	00065	00046	110056	-,00H14
	000051	(.00005)	{.00005}	(.00085)	(.00005)	(.00012)
	.216	.112	.254	.099	-363	. 372
	(1032)	(.016)	(.035)	(1016)	(1064)	(+048)
Flore \$. 336	.358	.338	.322	.327	.30 L
	(.023)	(.028)	(.023)	(.030)	(.028)	(.051)
Emc2	. 179	.218	.181	.209	.231	.260
	(.016)	(.018)	(.018)	(.021)	(.021)	(.046)
E197.4	069	146	-,002	120	026	415
	(.039)	(.048)	(.044)	(.067)	(.044)	(.092)
EDC5	195	190	173	179	165	244
	(.016)	(.018)	(.016)	(.018)	(.018)	(.035)
(BC88F8,	.038	.044	.034	.052	.045	.070
	(.005)	(.005)	(.005)	(.005)	(.005)	(.00 9)
MARK!	023	.016	038	.038	009	035
	(.014)	(.014)	(.014)	(.014)	(.016)	(.030)
wara	192	009	141	008	+.279	.029
	(.016)	(.018)	(.016)	(.016)	(.030)	(.0351
Canstant	t.318	1.112	1.282	1.098	1.172	.993
H²	.24	. 18	.22	. 19	. 20	. 26

⁻ Represents sero. Y Not applicable.



Table F. Standardized Coefficients of Regression of Log of Hourly Earnings on Specified **Explanatory Variables**

Variable	Males	fema les	white males	Willis femnies	Mates 30 and over with no familial in- terruptions	Funaley 30 and over with no familial in- terruptions
UNEMP	-0.026	0.001	-0.017	0.002	-0.050	0.008
DISAR	056	018	067	037	073	~.098
IIME-AWAY	034	069	007	085	(x)	(x)
FXPER	. 858	. 546	.931	. 582	.698	.203
EXPERSO	653	419	735	~.438	628	. 159
۲۱۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰۰	.072	.084	.086	.075	.070	.178
rnect	. 162	. 158	.167	. 142	. 168	.147
EDPC2	. 125	-156	. 130	.152	. 152	.148
eture4	018	018	001	024	007	108
kbtc5	[49	[48]	131	134	~.150	201
CDIRISES	- 131	- 160	.118	. 188	.154	.223
MARIRZ	018	.016	029	.036	008	031
MARIO	131	007	095	00b	122	.021

Table G. Percent of the Male-Female Earnings Gap Explained by Experience. Interruption. and Education Variables

Variable	hmployed Persons. a(C races	Employed Permons. White	Employed persons 30 years of age and over with no familial interruptions
Hourly chrotoss of mea	:6.92	.27 . 14	<u> </u>
Bourly carniags of women	14.38	14.42	!4.71
Earnings gop	<u>†2.54</u>	\$2.72	12.87
Hourly caraings of women adjusted for experience and interruption variables!	!4.69	14.76	34.85
Offference between actual hourly carnings of women and hourly carnings addisted for experience and interruption variables	÷ .31	₹.34	₹.14
Percent of carnings gap explained by experience and interruption variables	12.2	12.5	4.9
linarly earnings of women adjusted for experience. Intercuption, and education variables?	74.75	14.80	f5.64
Difference between actual boarly carnings of women and hourly carnings adjusted for experience, interruption, and education variables	₹.37	86. 1	t .93
Percent of earnings gap explained by experience. interruption. and education variables	14.6	14.0	32.4

Hourly carnings of women if women Mad the same mean values as men for experience and interruption variables. Experience and interruption variables include (NEMP, BISAB, TIME-AWAY, EXPERS EXPERSO, and FT. Blourly carnings of women if women had the same mean values as men for experience, interruption, and education variables. Education variables include EDUC1, EDUC2, EDUC2, EDUC4, EDUC5, and COURSES.



Table 1. Lifetime Work Interruption Status, by Reason, Employment Status, and Sex

exposers to thousands t

Imployment status and sex	ſ	All reasons	.#******	Inability to	find work	[amily re	a source	(1) mess or di	Salil I Lity
	nowher total	Yunter	Percent	Studiore	Percent	Simile- r	(Nego ent	Yugatu-r-	Percent
Porsons 21 to no years of age who ever worked Males	50,878 57,218	15 , 156 64 , 156	25.7 71.9	9,632 8,157	17.1 14.2	828 16.715	1.5 66.1	5.275	;n, € 9. 3
unplayed persons 23 to be vesce of use: Males	47,346 34,590	9, Mai Minbb	20.5 62.7	7,242 4,987	11.4 11.5	415 18,779	,9 9947	3,031 1,954	6.4 5.6

Table 2. Proportion of Potential Work Years Spent Away From Work, by Employment Status and Sex

Dymbers in thousands)

İ			Proport	ton of poli	utial work	years spec	it was tro	work for	reasons su	rveyed	
Employment status and sex	Fotul ուրական	No)II¢	i t pro		\0 1 p+rc	n 24 cent	25 L	5 49 CC#1	50 per and c	
	reporting	Soutere	Percent	Number	Percent	Numbe r	Percent	Number	Percent	Namber	Percent
Persons 21 to be years of age who ever worked;]				
Venaltys	\$1,321 50.1 8 1	41,489 16,105	77.8 32.1	7,590 5,213	14.2	2.556 5.514	4.8	1,111 8,547	2.1	575 14,844	1.1 29.6
Employed persons of to 64 years of age:											
Males	45.544 31.534	37.6. ³ 3	82.6 40.9	5,9117 4,005	13.0 12.7	1, 174 3, 727	3.0. ti.d	426 4,806	,9 15.2	.114 6.085	. 5 19. 3



Table 3A. Lifetime Work Interruption Status, by Reason and Selected Characteristics: Males

	ļ		M)(It une	or more inter	rruptions last	ing 6 months	or more due 1	0	
Characteristic		All reas		Inubilit Ilad w		Familia reason		[liness disabil	
	Total ngsher	Number	Percent	Number	Percent	Number	Percent	Sumber	Percent
Males 21 to be years of age.	55,828	14,355	29.7	9,63.	17.3	818	1,5	5,418	10,6
tau) and spantsh origin		}		1]		
Spanish origin	49, 181 1,627 1,220	11,942 2,263 1,124	24 . 2 40 . 2 34 . 9	7,518 1,970 731	15.2 35.0 22.7	606 217 38	1,2 3,9 1,2	5,305 604 509	10.7 10.7 15.8
NGF.				1	1	j	ļ		
21 to 29 years	16,048 19,106 50,674	4,507 6,594	20.5 23.4 31.9	2,886 3,087 3,657	18.0 16.2 17.7	221 300 308	1.4 1.6 1.5	563 1,618 3,758	3.4 8.5 L8.5
VENUS OF SCHOOL COMPLETED					1	1	1		
12 t + 1 t	14,171 29,761 11,896	5,686 7,368 1,309	40.1 24.7 11.0	3,535 5,159 938	17.3 7.9	206 373 190	1.9 1.1 1.6	2,878 2,754 285	20.3 9.3 2.4
AGE BY YEARS OF SCHOOL COMPACTED									
21 to 29 years; less than 12	2,315 10,105 7,630	940 2,102 249	40.7 20.8 6.9	8, u 1,867 200	15.5 18.5 5.5	52 146 23	2,2 1.4 .e	170 328 29	7.4 3.2 1.2
10 to an years: Less than 12	3,809 10,278 5,019	1,397 2,588 483	36.7 25.2 9.6	933 1,828 326	24.5 17.8 6.5	79 151 70	2.1 1.5 1.4	692 825 107	18.2 8.0 2.0
bess than 12. Li to 15. It and over	8,049 9,378 3,247	3,348 2,670 577	41.6 :8.5 17.8	1,781 1,464 41.	22.1 15.6 12.7	1 35 75 97	1.7 .8 3.0	2,016 1,602 140	25.1 17.1 4.3
OCCUPATION GROUP OF USUAL FOR	{						-		
Professional, feelineal, or managerial	15,040 b,621 12,825 10,254 5,832 3,457	2,205 1,361 3,697 4,047 2,210 881	14.7 20.6 28.8 32.5 37.9 25.5	1.532 913 2.402 2.131 1.612 512	10.2 13.8 16.7 20.8 27.6 14.8	251 61 179 83 129 80	1.7 .9 1.4 .8 2.2 2.3	796 486 1,727 1,532 791 384	5.3 7.3 13.5 14.9 13.6
LUH-eyeta sixus	}	{		1		}	ŀ	1	
Il to 79 coars: Ever married	9,245 6,803	1,796 1,503	19,4 22.1	1,579 1,308	17.1 t9.2	95 126	1.9	290 252	3.1 3.7
SO to a venre: Ever married Never married	10,825 1,281	3,942 326	22.1 41.0	2,646	14.8 34.4	294 1	1.7	1,477 1-1	8.3 11.0
Volto by years: Even marked	19,842 79?	6.139	30.9 57.5	3,429 232	17.2	273 34	1.4	3,464 293	17.4 37.0

^{*}Persons of SpanIsh origin may be of any race.



Table 3B. Lifetime Work Interruption Status, by Reason and Selected Characteristics: Females

			With one	or more inte	eruptions last	ing benontlis	or more due t	1)==	
Characteristic		All rea survey		Innbilit Lind w		Film 11 renso		filmess or disability	
	Total number	Number	Percent	Number	Percent	Number	Percent	N-palaer	Percent
Females 21 to 64 years of age who ever worked	57,258	41,156	71.9	8,157	14.2	36,715	64.l	5,275	9.;
PACE AND SPANISH ORIGIN ¹									
White	49,812 6,402 3,014	36,362 4,039 2,259	73.0 63.1 75.0	6,198 1,757 712	12.4 27.4 23.6	33,297 2,805 1,882	62.4	4,146 1,123 389	8 17.5 12.9
AGE			İ						
31 to 29 years	16,804 19,445 21,011	8,919 15,069 17,168	53.1 77.5 81.7	2,865 2,398 2,893	17.0 17.3 13.8	7,143 14,065 15,507	42.5 72.3 73.8	596 1,268 3,391	3.5 6.6 16.1
VEARS OF SCHOOL COMPLETED			ļ	į					
Less than 12)),74D 34,805 H,71)	10,917 25,504 4,715	79,5 73.3 34.3	2,978 4,430 748	21.7 12.7 8.6	9,418 23,064 4,232	68.5 6t.3 48.6	2,761 2,289 225	20.1 6.6 2.6
AGE BY YEARS OF SCHOOL COMPLETED			Ì				İ		
21 (o.29 years) Less than 12	1,948 11,650 3,206	1,374 6,612 937	70.6 56.8 29.1	452 2,148 266	23.2 18.4 8.3	1,201 5,232 710	. 61.7 44.9 22.2	113 580 2	5.4 4.1 .1
30 to 35 years: Less than 12	4,060 12,766 3,018	3,240 9,863 1,966	79.8 79.8 65.1	827 1,226 344	20,4 9,9 11,4	2,987 9,313 1,765	71.6 75.3 58.5	499 701 86	12.3 5.3 2.9
45 to be years: less than 12	7,733 10,789 2,489	6,302 9,029 1,836	81.5 8J.7 7J.8	1,649 1,055 198	22.0 9.8 5.6	5,230 8,520 1,757	67.6 79.0 70.6	2,147 1,109	27.8 10.3 5.4
OCCUPATION GROUP OF USUAL JOH								; 	
Professional, technical, or managerial	11,723 23,782 489 8,447 950 10,543	7,152 17,888 118 6,122 744 7,818	61.0 79.2 65.1 74.8 78.3 74.2	1,113 2,541 156 1,87) 209 2,079	9,5 10,7 31.5 22.2 21.9 19.7	6,497 16,510 220 5,310 644 6,682	55.4 69.4 45.0 67.9 67.8 63.4	135 1,395 21 1,223 143 1,744	5.4 5.9 4.3 14.3 15.1 16.5
LIFE-CYCLE STATUS			1						
21 to 79 years; byer married; With children	7,779 3,750	6,308 1,216	81.1	1,254 621	[6.] 16.6	5,918 799	76.1 21. J	114 77	4.0 2.1
Never married: With children So children	1,218 4,050	530 845	43.5 20.8	366 624	30.0 15.4	268 157	27.0	39 166	3.3 4.1
in to 24 venes: Ever married: With children	14,255	12.046	84.5	1,707	12,0	11,646	81.7	859	6.0
No children	4,103	2,633	64,2	3/2	9.1	2,336	56.9	363	8,8
With children	193 894	90 2 9 9	46. 9 33.4	75 244	38.7 27.3	63 20	32.7 2.2	16 50	8.5 5.6
6, to 56 Vegra: Ever married	19,945 1,065	16,717 440	83.8 42.1	2,633	11,2 24,4	15,347	76.9 15.0	3,146 265	15.6 23.0

tpersons of Spanish origin may be of any race.



Table 4A. Lifetime Work Interruption Status, by Reason and Selected Characteristics: Employed Males

Combine in thousand of

	1		With one	er were tale	reaptions last	ing 6 nonths	or ware dite.		
estina terristic		All rea sarvey		fitale(1) (Franci I Pera Sed	I	111mess disabil	
	nuoteer	Sumber	Percent	Studen	Percent	Samber	Percent	Sumber	la-pern (
Employed males 21 to be come of agreement	.7 ₃ 540	9,724	20,5	7,232	15.3	415	4,9	1,031	6.4
RWT AND MPANISH ORIGIN ¹				1		i			
Black	52,116 5,517 2,680	8, (35) 1, -73 629	19.3 32.6 25.4	5,770 1,352 504	13.7 19.9 20.1	360 51 35	1.1 1.4	7,76H 253 176	6.6 5.6 7.1
₹c₁}			-	Ì		-	1		
A to 29 years	14,069 17,376 15,906	2,536 1,368 1,819	18.0 19.4 24.0	2, 1/6 2, 163 2, 543	16.5 13.6 16.0	1 93 168 115	,9 1.0 -7	277 1,106 1,647	2.0 6.4 10.4
YEARS OF SCHOOL COMPLETED				ļ					
Less than 17.,	10,834 25,656 10,856	6, 181 7, 381 961	31.2 21.0 8.8	7,190 4,26, 677	21.1 16.6 6.2	196 200 110	1.0 .8 1,0	1,439 1,484 ,00,	12.4 5.8 1.9
AGE, BY VEARS OF SCROOL, COMPLETED	Ì							İ	
21 to 9 cours; Acs Star 12 12 to 1 16 and occurrence	1,804 8,851 3,417	562 1,726 248	31.2 19.5 7.3	506 (1,6,0) 199	28,1 18.3 5.8	1 (107 23	1.2	67 165 44	1.1 1.5 1
Bo to a voices: tiess than 12 17 to 15 Th and over	3.295 4.260 4,816	1,005 1,892 471	30.5 26.4 9.3	045 1,403 31 N	19.6 15.1 6.5	33 65 70	1.0 .7 1.5	517 489 100	15.3 5 2.1
6) to be years; Loss than 1/2	5,736 7,537 2,628	1,814 1,761 242	11.6 23.4 9.2	1,140 1,241 163	19.9 16.5 6.2	70 28 17	1 - 2 - 4 - 6	755 830 62	13 11.6 2.3
OCCUPATION GROUP OF USUAL JOH	ļ					-			
Protessional, technical, or managerial. Sales or clearful. Cracts persons. Operatives. Lationers.	12,873 6,418 (0.978) 8,583 4,506 3,050	1,238 1,072 2,602 2,324 1,319 677	19.4 17.8 23.7 27.1 29.3 22.2	1,022 768 1,882 1,618 1,139 408	8.0 15.1 27.1 19.1 25.3 13.4	96 44 78 77 52 65	.8 .7 .7 .9 1.2 2.1	342 314 1,976 839 256 223	2.1 5.1 9.1 9.1 5.7 7
LIDER-CYCLE STATES							1		
M to 29 years: Ecor parried Seven parried	8,671 5,399	1,538 998	17.7 18.5	1,3°6 950	15.9 17.6	6L 71	.7	191	2.5 1.0
Butto ab years: Even married	16,332	J,021 347	18.5 33.3	2,077	12.7 27.4	168	1.0	1,021 85	6 8.1
n to the yearst: Even majorited	15, 343	3,525 293	23.0 52.6	2,375 168	15.5 30.2	105 La	1.9	1,496 151	9.7 27.1

⁻ Represents vero-



 $^{^{1}\}mathrm{Persons}$ of Spanish origin may be of any race,

Table 4B. Lifetime Work Interruption Status, by Reason and Selected Characteristics: Employed Females

			atth one o	r more later	rapitons lasti	ng to months i	or notes due to	1	
Material System	-	All Pear		- Inablits tlm/s		Famil rea- n		Dimes≤ or disability	
	totat	Situatu-p	Percett	Natulne 2	Percetti	Samtle t	Percent	States	Percent
Cultural to aless 1 to 6.	3, 191	71,088	62.7	4,582	11.1	18,739	>4	1,950	5.6
13c1 AND SPANEORORAN ¹				-			ŀ		
Minute	. 5 54 - , - > 1 1 , 7 34	1848.00 24508 14121	64.0 50.1 64.4	3,317 1,192 395	11.1 26.2 22.6	16,817 1,628 867	57.1 31.8 39.9	1,327 627 178	4.5 13.7 10.2
14.1									
1 2	13 , 200 12 , 751 10 , 550	7, 319 9,073 8,140	39,4 70,4 78,3	1,699 8,4,1 0,6,1	14.8 11.3 13.9	3,342 8,300 7,190	28.4 65.1 69.2	474 640 344	3,0 5.4 8.9
भवन्य ज्ञाल्य स्वयम् वर्	:								
for the formation of the first section of the formation o	0.69° 1.55 6,565	4,003 1,,290 2,751	69,3 69,4 69,5	1 pol 3 2 pol 1 1 pl 2	21.4 12.2 8.5	3,777 12,741 2,421	50.4 57.4 40.0	645 1,219 90	9,6 1.5
301 By ALMS OF SURFOLD COMPUTED	. !	 							
The Charles of the Ch	971 7,737 7,738	u= 1,334 nee	31.7 33.1 25.1	19) 1,201 223	19.9 10.7 8.1	422 2,329 490	43,5 30,1 17,9	2† 317	2.1 4.1 .1
described States	1, 311 11, 416 1, 414	1,63c e, 293 1,595	70,8 74,7 54,6	द्धाः श्रदेक [अप	17.5 10.0 9.4	1,454 5,858 5,898	62.5 69.5 49.8	233 416 37	19.1 4.9 1.8
Teleschen I.	3,016 (,03) (,30)	2,519 2,663 97,	73.5 81 74.7	8,74 5,76 1,061	23.1 9.2 7.7	1,911	70.6 71.7	104 230 21	11.5 8.0 3.9
ong a OVERAN COMPANI SUATE FOR									
There year beauty if ending all ending the state of the s	7,6% 1-,000 -77 -,76% -470 -6,14%	3, 499 9,803 163 3,138 291 1,925	7,07 4,79 4,19 4,49 4,49 9,18	707 1,581 43 1,303 61 1,083	973 1975 1779 2372 1472 1776	3,428 8,818 132 2,601 257 3,279	44.7 60.3 48.2 50.5 58.0 53.4	197 188 5 409 9	2.0 4.1 1.8 8,6 2.2 12.2
1111-1211-51305									
Type expense Ive appropriate Area children Next estable Next estable	1,194 3,037	2,423 739	69.9 24.3	719 367	17.2 12.1	2,615 259	62,5 15,1	1 52 33	3.6 U.I
See a farence	478	25.7 599	29.3 17.9	21.2 400	24.2 11.9	73 46	#.3 2.8	tas	1.0 4.3
Principal Companies (1998) [1998] Conflored: Anthropolitic (1998) Anthropolitic (1998) Anthropolitic (1998)	8,769 4,093	P.923 1,889	78.9 61.1 32.9	971 295 35	11.1 9.5 25.1	6,640 1,640	75.7 53.0	593 140 14	5.6 4.5 10.3
New Hilliam	740	165	22.3	1 38	18,6	ı	1.	4.1	5.1
Note to a protest	9,599 901	7.921 225	82.5 28.0	1,319	13.7	7,128	74,3 8.6	556 53	8.8 10,4

There are in spacish origin bay he of any race.



Table 5A. Proportion of Potential Work Years Spent Away From Work, by Selected Characteristics: Males

Our door of thous ones?

	· T		Proport	ten at pate	uttul work	years spec	at away Iros	work for	reasons su	veyent	
Olignis is milidas		No.	իւ	1 10 9	percent	10 to 74	percekt	25 to 49	percent	50 parcent and over	
	Total -	Studeer	leers out	Sumber	Parrent	Sumber	Percent	Sumber	Percent	Supher	Percent
Water 21 to be yours of a co- kho symmetheters	,H2H.	41,520	75.6	8,751	15.7	3.4tu {	5,1	1.717	1.1	4 4(1	0.8
ICACE AND SPANTSH ORIGIN ¹				1		l			ļ		
Bhile	0, 181 5, 6, 7 5, 270	47 ₄ 054 5 ₄ 394 7 ₄ 096	25.8 60.1 61.1	7,4.74 1,110 760	15.1 20.7 23.6	7,816 5% 239	1.7 9.5 7.4	1, J84 143 [19	7.8 n.1 1.7	25.3 191 6	.t. 3.4 .2
N 1	į	ļ	į				}		1		
H. J. M. Sentherson	164648 194166 204676	12,755 14,639 14,126	79.5 76.6 68.5	851 7,101 6,742	9.3 16.7 22.9	1,350 882 1,235	8+2 4+5 (6-11 (887 321 532	5.5 1.7 2.6	.; j.c. 165 8t	1.5
MAIS OF CHOOL COMPANY				}			}	j			
Less than 1	14.171 29,761 11,895	8,485 72,447 10,587	59.9 75.4 89.0	3.641 4.435 695	25.7 14.9 5.8	1,215 1,785 193	8.0 6.0 6.1	695 840 18 i	4.9 2.8 1.5	135 253 38	1.0 .9 .3
AGE BY ARAGS OF SCHOOL COMPLETED				ļ	į				ļ		
10 9 years; Lees Dan 1	2,315 10,105 3,6 at	1,373 8,002 3,181	59.6 79.7 91.1	405 512 40	13.2 5.1 1.1	410 841 67	18.0 8.3 1.7	163 613 102	7.1 6.1 2.6	56 117 45	2.4 1.4 1.2
90 To so years: Tess (Bah 17)	1,809 10,278 5,019	2,417 7,190 4,536	61.3 74.8 90.4	1,079 1,826 262	27.0 17.8 5.2	182 466 200	4.8 4.5 4.1	112 200 13	2.9 1.9 .3	76 97 •	1,9 .9 -
us to the years: Less shan 12	8,049 9,178 1,247	4,701 6,755 2,671	58.4 72.0 82.2	2,286 2,075 373	28.4 22.1 11.5	629 487 119	7.8 5.2 3.7	418 39 85	5.2 .4 2.0	16 21 1	.?
GCCLDATTON GROUP OF USLAY, JOB		ĺ					į		j		
Probasional, to high it or amaged do	15,060 6,621 12,825 10,254 5,832 3,457	12.851 5.259 9,128 6,917 5,653 2,576	85.4 79.4 71.2 67.5 62.6 74.5	1,361 874 2,303 2,102 1,150 489	9.4 13.7 18.0 20.5 19.7 14.1	421 141 1,045 768 595 161	2.8 5.2 8.7 7.5 10.2 4.7	298 120 313 445 263 197	2.0 1.8 2.4 4.1 4.5 5.7	168 24 36 21 171 45	.7 .4 .3 .2 2.9 1.0
PD 6-0. ATM 2		- 1			}	- 1	}				
21 to the years 1 I year in Consect	6.545 6.803	7,455 (5,300	30.6 77.9	699 113	7.0 1.7	673 655	7.3 9.6	364 542	1.9 8.0	54 194	. b 2 . 9
sb for a years; Ever married	17,805 1,081	15,883 746	77.9 59.0	?.717 382	15.2 29.8	817 66	4.6 5.2	296 25	1.7 2.0	111 52	.6 4. L
Control the graphics Excent majorates	19,882	1 1,79b 337	69.4 62.5	4.504 747	72.7 11.2	1, (O1 120	5.5 16.1	450 79	2.1 10.0	рь 2	.7

popposable zapo,



^{*}Porsous of aparish origin may be of any race.

Table 5B. Proportion of Potential Work Years Spent Away From Work, by Selected Characteristics: Females

(Simbers to thousands)

	L		Proporti	an of pote	ntini work	years apor	it nway from	n work for	reagons see	eveyed	
Charneteristic	Total	No	ne	1 to 9 ;	Percent	10 to 24	Percent	25 ta 49	percent	50 per	
	number	Number	Pervent	Number	Percent	Number	Percent	Number	Percent	Number	Percent
Females 21 to 64 years of age who ever worked	57,258	16,105	28.1	5,502	9.6	8,221	14,4	10,779	18.8	lo,653	29.1
HACE AND SPANISH ORIGIN ¹				1							
Nhite	49,812 6,402 3,014	13,453 2,362 756	27.0 36.9 25.1	4,217 1,262 207	0.5 19.7 6.9	6,982 1,049 825	14.0 16.4 27.4	9,529 1,029 535	19.1 16.1 17.7	15,633 700 693	31.4 10.9 23.0
40A	}	į			1		i	1	ł	}	
21 to 29 years	16,804 19,445 21,011	7,685 4,377 3,843	46.9 22.5 18.3	665 2,193 2,665	4.0 LL.3 L2.7	2,809 2,710 2,658	16,7 13,9 12.7	2,532 3,756 4,492	15,1 19,3 21,4	2,913 6,408 7,352	17.3 33.0 35.0
YEARS OF SCHOOL COMPLETED		ļ		1							
Less than 12. 12 to 1)	13,740 14,805 8,713	2,826 9,301 3,978	20.6 26.1 45.7	1,520 3,500 489	11.1 10.1 5.6	2,383 4,829 1,043	17.3 13.9 12.0	2,791 6,543 1,436	20.j 18.8 16.5	4,221 10,631 1,766	30.7 30.6 20.3
AGE BY YEARS OF SCHOOL COMPLETED				1							
21 to 29 years: fwss than 12	1,948 11,650 3,206	573 5,038 2,274	29.4 43.2 70.9	122 525 25	6.3 4.5 .8	277 1,977 516	14.2 17.0 16.1	542 1,846 172	27.8 15.9 5.4	436 2,764 219	22.0 19.4 6.8
19 to 44 years: Less than 10	4,060 12,366 3,018	822 2,503 1,052	20.3 20.2 34.9	417 1,622 136	10.3 13.1 4.5	694 1,691 336	17.1 13.7 11.1	813 2,338 610	20.0 18.9 20.4	1,313 4,213 878	32.4 34.1 29.0
4) to 64 years: Boss than 12	7,733 10,789 2,489	1,430 1,760 653	18.5 16.3 26.2	990 1,348 329	12.8 12.5 13.2	1,415 1,165 164	t8.3 10.8 6.6	1,425 2,365 656	18.4 21.4 24.4	2,471 4,151 687	32.0 38.5 27.6
OCCUPATION GROUP OF USUAL JUNE										ĺ	
Professional, technical, or managerial. Sales or Clerical. Crafts Dersons Operatives. Laborers. Service.	11,723 23,782 489 8,447 950 10,543	4,571 5,894 171 2,124 208 2,725	39.0 24.8 35.0 25.2 21.9 25.9	823 2,302 48 1,188 17 959	7.0 9.7 9.8 14.1 1.8 9.1	1,618 3,129 62 1,421 90 1,718	13.8 23.2 12.7 16.8 9.5 16.3	2,027 4,717 88 1,686 203 1,829	17.3 19.8 18.0 20.0 21.4 17.4	2,686 7,741 120 2,027 433 3,312	22.9 32.6 24.5 24.0 45.5 31.4
LIFE-CYPLE STATES							•]		
21 to 29 years: Nor married No children Nover married With children No children So children	11,529 7,779 3,750 5,275 1,218 4,056	3,985 1,471 2,514 3,900 688 3,212	34.6 18.9 67.0 13.9 56.5 79.2	545 326 220 121 45	4.7 4.2 5.9 2.3 3,7 1.8	2,270 1,919 351 551 92 441	19.7 24.7 9.4 10.5 7.6 10.9	2,130 1,693 437 404 140 262	18.5 21.8 11.7 7.7 11.5 6.5	2,599 2,370 228 299 253 67	22.5 30.5 6.1 5.7 20.8 1.7
30 to 44 years: Ever married. With children. No children. Never married. With children. So children.	18,358 14,255 4,103 1,087 193 894	3,680 2,210 1,470 697 102 595	20.1 15.5 35.8 64.1 53.1 66.6	2,101 1,527 594 95 32 62	11.4 10.7 14.5 8.7 16.7 6.9	2,525 2,060 466 194 9 188	13.8 14.5 11.4 17.9 4.7 21.0	3,661 3,011 648 95 47 45	19.9 21.1 15.8 8.7, 24.5 5.0	6,391 5,448 925 6	34.8 38.2 22.5 .6 1.0
A5 to 64 years: Lver married	19,945 1,065	3,228 615	16.2 57.8	2,495 224	12.5 21.0	2,603 47	l3. t 4.4	4,357 145	21.8 13.6	7,263 34	36.4 3.2

¹Persons of Spanish origin may be of any race.



Table 6A. Proportion of Potential Work Years Spent Away From Work, by Selected Characteristics: Employed Males

(Symbols to thousands)

	1		Proporti	ion of pote	nt fal work	yours spet	t owar troo	work for	Pensons Sur	ya ya d	
theracteristic		No	Dt.	L to 9	percent	10 to 24	percent	25 t u 49	pe rc) #1	D p+ternt and o v (r	
	Tertal - Numitare	Number	Percent	Number	Percent	Number :	Pore, at	Smala r	Percent	Nauts r	Percont
Employed males 71 for his years of age	47,346	37.654	79.5	6,825	14.4	1,939	⊣. i	756	1.6	174	0.4
BACE AND SPANISH ORIGINAL				{						ļ	
Shife	43,116 4,517 2,480	33,981 3,074 1,8>1	80.7 68.1 74.6	5,898 825 538	14.0 18.3 21.7	1,578 342 46	3.B 7.u 1.9	565 194 38	1,3 4,3 1.)	u., 81 7	 1,8 1,8
)L+t			}								
21 to 29 years	14,069 17,376 15,900	11,533 14,008 12,772	82.0 80.6 76.2	753 2,572 3,527	5.4 14.8 22,2	992 721 209	7.1 4.2 1.3	057 38 52	+.7 .3	133 38 -	1.0
YEARS OF SCHOOL COMPLETED			1								
less than 12	10,834 25,636 10,856	7,453 20,305 9,895	68.8 79.2 91.1	2,760 3,563 550	25.5 13.9 5.1	395 1,227 282	3,7 4.8 2.6	181 469 94	1,7 1,8 .9	44 41 50	.4
AGE BY VLARS OF SCHOOL COMPLETIN							·				
21 to 29 years; Less than 12	1,804 8,853 3,412	1,242 7,127 1,164	66.9 80.5 92.7	248 467 40	13.8 5.3 1.2	179 741 62	9,9 8,4 1,8	170 441 101	6,7 1,0 1,0	1 i /8 +>	.8 .4 1.3
30 to 34 years: Less than 12	3,295 9,266 4,816	2,289 7,373 4,345	69.5 79.6 90.?	846 1,496 256	25.7 (16.2 5.3	127 364 202	3.9 3.9 4.2	3 21 12	.1 .? .3 (30 11 -	. 4 . 1 -
45 to 55 graps: tass than 12	5,736 7,537 2,628	3,922 5,805 2,386	68.4 77.0 90.8	1,673 1,624 227	29.2 21.5 8.6	84 109 14	1.5 1.5 .5	13	1.0	: -	- - -
OCH PATTON LAWER OF CSCAD JOB		İ									
Professional, technical, or managerial. Sales or clerital. Crafts persons. Operatives. Lation ps. Service.	12,823 6,018 10,978 8,583 4,506 3,050	11,485 4,946 8,376 6,259 3,218 2,373	89.6 82.2 76.3 72.9 71.4 77.8	1,032 733 1,800 1,687 988 418	8.1 12.2 16.4 19.7 19.7 13.7	169 298 593 514 239 68	1.3 5.0 5.4 6.0 5.3 2.2	135 14 709 120 54 171	1,1 1,9 1,4 1,2 >,6	26 - 3 107 20	
LIFE-CYCLE STATES	ļ	}	ĺ								
Plate 29 venest Eyer Carried Never married	8,671 5,399	7,133 4,400	82.3 81.5	688 55	7.9 1.0	509 487	5.9 9.0	319 342	3.: n.3	21 114	2.1 2.1
30 to 22 years: Ever married Never married	16,332 1,044	13,311 697	81.5 66.8	2,300 271	14.1 26.0	673 52	4. L 5. O	37 2	.:	12 22	. 1 2. 1
4) to 54 years; fyer married	5,343 558	11,848 264	77,2 47,3	3,306 229	21.6 41.0	135 65	_9 [1.7	53	· · · · · · · · · · · · · · · · · · ·	-	-

⁻ to presents with.



22

¹Persons of Spanish origin may be of any race.

Table 6B. Proportion of Potential Work Years Spent Away From Work, by Selected Characteristics: Employed Females

Compacts, and those control

			Proporti	on of pote	ntial work	genra sper	nt away Iron	n work jur	prasons sur	weyed	
Charmert Cle		Non	1.	1 t+9 j	percent	40 to 34	percent	25 to 49	percent	1 Ng De 1 Ng A	
	fortet Humbs P	Sumbs r	Percon	Number 1	Percent	Suntier	Percent	Sumba r	Percent	Samiliary	Percent
Populoved George (T) to	Liv. y 2 ¹²⁸ 1	17,911	17,3	a , 100	12,0	5,247	11.2	1,784	16.7	6,483	L#.7
EACE ASE CRANTER ORDER AS			ļ								
Algreen,	29 ₃ 4594 41, 64 1 ₄ 739	10 ₁ 613 2 ₄ 663 619	36.48 46.49 34.46	3,132 977 138	10,7 21,5 7,9	4,418 738 651	15.0 16.7 37.4	5,273 433 473	17,9 9,5 9,9	6,003 339 139	20.4 7.5 9.1
1)											
Attorney search and the season and t	11,456 12,751 10,400	6,937 3,720 2,254	607,6 297,2 217,7	474 1,769 1,936	ا , 9 . د ا 18 . و	1,987 1,863 1,353	17.4 14.8 13.0	1,035 2,288 2,472	9,0 (18.0) 23.8	1,027 1,081 2,383	8.9 24.2 22.9
VI VICE OF TOTAL										ĺ	
Less Dent L.,	6,697 1,847 6,07	2,056 7,57 3,297	30,7 14,6 14,5	011, 1 878, 5 414	10,6 12,2 7,1	1,223 : 3,302 741	18.3 12.3 12.3	1,24a 3,788 755	18,6 17.3 12.5	1,061 (4,542 826	8,61 8,0% 13,61
क्रक (१८ ४) अ१५ छ। ,त (१००) - एक १५, १३ छ	(ļ	ļ		Ì	;				
1 to 29 v. abs; 1c s. than 12	971 7,737 2,748	.74 4,404 2,059	58,8 56,9 75,0	67 187 29	5,9 5,0 ,9	145 1,357 463	14.9 17.5 16.9	.203 789 147	20,9 40,2 2,4	82 802 173	8,4 10,4 4,8
10 to 1. scarst 14 so thin 1' 15 to 1 16 uid oyet	2;311 8,426 8,426 2;000	ь77 2,133 940	79, 3 25, 3 45, 4	778 1,348 122	12.0 16.0 6,1	549 1,201 154	23.8 14.3 7.5	374 1,565 356	10,2 18,6 17,8	413 2,180 464	18,7 21,9 23,2
to to be real; Toss thin 12),416 5,683 1,300	906 020, 1 328	26,5 18.0 25.3	766 924 273	22.4 16.2 21.0	527 745 102	15,4 13.1 7.9	677 1,438 341	19.8 25.3 26.3	741 1,518 275	15.8 27.5 19.6
OGCUPATION THREE PAR CATAL JOB					Ì						
Oropes Jonel, recipical, or wonterful. Sales or elevical. Crafts presons. Juntates.	7,676 14,109 272 4,760 429 6,142	3,786 4,706 109 1,602 117 2,218	49.3 32.4 39.9 33.7 32.1 36.1	586 1,789 55 926	7.6 12.3 20.2 19.5 - 12,2	1,032 { 2,141 21 792 76 1,078	13.5 14.8 8.4 16.6 17.3 17.6	1,115 2,735 69 808 79 828	14.8 18.9 25.3 17.0 18.5 13.9	1,134 3,137 17 632 137 1,266	14.8 21.6 6.2 13.3 32.1 20.6
LODE CLICATORS			ļ	1	}	ļ		{			
Fig. 19 (co. 29 (co. 28) Fig. 19 (not rich) With children No children No ver married With children No children	7,220 4,184 3,035 4,235 877 1,356	3,559 1,261 2,297 3,378 621 2,757	49.3 30.1 75.7 79.8 70.8 82.2	195 143 202 78 29 48	5.5 4.6 6.7 1.8 3.3	1,537 1,307 231 449 66 386	21.3 31.2 7.6 10.6 7.5 11.5	859 647 211 178 75	11.9 15.5 7.0 4.2 8.6 3.0	870 770 94 152 96	42.4 78.5 3.4 3.6 9.8 1.9
O to 55 years; Fyer married, With children. So shilden. Never married, With children. So children.	11,863 6,770 3,095 878 138 739	3,052 1,848 1,204 668 93	25.7 21.1 38.9 76.1 67.4 77.8	1,676 1,160 530 106 35	14.1 13.2 17.1 12.1 25.4 9,3	1,813 1,420 394 77 8 70	15.3 16.2 12.7 8.8 5.8	2,256 1,795 459 27 2 25	19.0 26.5 14.6 3.1 1.4 3.4	3,004 2,547 508	25.9 29.0 16.4 -
e to 56 years; fver married Never married	9,590; 801	1,678	17.5 72.0	1,790 177	18.7 (22,1	1,344	14.0	2,441 19	23.6 2.6	2,346; 28	24.4

⁻ Represents zero.



⁴Dersons of spanish origin may be of any race.

Table 7A. Mean Percent of Potential Work Years Spent Away From Work, by Reason and Selected Characteristics: Persons Who Ever Worked

		Mates		_	Females	
Characterist)e		Mean percent work years spe work for rens	nt hway from		Mean percent work years spe work for reas	of away from
	Somber Othousands)	Value	Stundard error	Sumber (Chousands)	Va Cnc	Standard + From
Persons 21 to 64 years of age whn ever worked	55,828	3.3	0, 1	57, 258	30.4	0.2
RACE AND SPANISH ORIGINA		[
White	49, 381 5, 627 3, 220	2.9 6.9 1.3	.1	49,812 6,402 3,014	32.7 17.6 27.6	. 2 . 3 . 9
AGE.		l				
21 to 29 years	15, 048 19, 406 20, 674	4.5 2.6 3.0	. 2 . 1 . 1	16, 894 19, 445 21, 011	20,7 34.3 35.8	.3 .3 .3
YEARS DE SCHOOL COMPLETED						
Less than 12	14.171 29.761 11.896	4,7 3,2 1,7	. 2 . 1 . 2	13, 740 34, 805 8, 713	33.5 31.5 24.2	.3
AGE BY YEARS OF SCHOOL COMPLETED						
21 to 29 years: Less than 12	2, 314 10, 104 3, 630	6.8 4.7 2.4	. 5 . 2 . 4	1,948 11,659 3,206	30.9 22.3 8.9	. 8 . 3 , 5
30 to 44 years; Leas than 12	3,809 10,278 5,019	1.1 2.8 .0	.4 .21 .2	4,760 12,366 3,018	34.2 34.8 32.2	.6 .3 .7
45 to 64 years: Les4 than 12	8, 049 9, 378 3, 247	4.3 2.4 2.7	.2	7, 733 10, 789 2, 489	33.7 37.7 34.3	.4 .4 .8
OCCUPATION GROUP OF USUAL JOB			ļ			
Orofessional, technical, ormanagerial. Sales or electeal. Crafts persons. Operatives. Laborers. Service.	15, 040 5, 621 12, 825 10, 254 5, 832 3, 457	2.3 2.3 2.9 3.9 3.0 5.0 4.1	. 2 . 2 . 2 . 3 . 4	11, 723 23, 762 489 8, 447 250 10, 543	24. 4 33. 8 26. 8 29. 4 39. 7 32. 1	.3 .2 1.7 .4 t.2
LIFE-CYCLE STATIS						
21 to 29 years: For married. With children. No children. Sever married. With children. So children.	9, 265; (SA) (SA) 6, 803; (SA) (SA)	2.9 (SA) (SA) 6.6 (SA) (SA)	. 2 (NA) (NA) (NA) (SA) (SA)	11, 128 7, 779 3, 750 5, 275 1, 218 4, 056	20.1 33.9 9.5 9.1 19.4 6.0	.4 .5 .4 1.0
30 to 44 years: Ever married With children No children Never married With children No children	17, 825 (SA) (NA) 1, 281 (NA) (NA)	2, 4 (NA) (SA) 5, 9 (NA) (SA)	(NA) (NA) (NA) (NA) (NA)	18, 358 14, 255 4, 103 1, 087 193 994	35.8 38.9 25.3 7.4 12.2 6.4	.3 .3 .6 .7 .7 .1,0
45 to 54 years; Ever married	19, 882 - 792	2.B 7.3	. l . 7	19,945 1,065	37.0 13.6	

SA Not available.

 $^{4}\mathrm{Persons}$ of Spanish origin may be of any race.



Table 78. Mean Percent of Potential Work Years Spent Away From Work, by Reason and Selected Characteristics: Employed Persons

		Ma l es		,) considers	
-harm-ters dis		Mean percent o work years sper work for reaso	M ARAY (PONE		Mean purcent of work years spec work for reason	d iwas from
	Nakper (Cousidats)	vatae	Standard error	Similar (Thomsands)	Value	5 Landard 4 Proc
Employed persons (1) to 64 years of age.	47, 340	2.1	0.1	347 496	22.6	17.2
KACC ASO SPANISH ORIGIN ³		i i			}	
Migre. Black	42.110 4.517 2.480	1.8 6.9 2.0	.1	29, 459 4, 551 1, 739	24.2 12.9 18.2	. 2 . 4 . 1
Arab						
21 to 24 years	14,069 17,376 15,900	3,5 1.6 1.4	.2 .1 .1	11,446 12,741 10,400	12.8 27.5 27.4	.3 .3
VIAMO DE MARON, COMPLETEN					[
fors than 12	10,834 25,656 10,856	2.6 2.2 1.3	.2 .1 .2	4, 697 21, 847 6, 052	23.3 23.7 17.8	.4
AGE BY YEARS OF SCHOOL COMPLETED]	
21 to 29 years; tess than 12	1, 804 8, 853 3, 412	4.5 1.7 2.6	.5 .2 .4	971 7,737 2,748	19.1 15.2 n.7	171 73 73
30 to 4 years; free than 12	3, 295 9, 266 4, 816	2.3 1.6 .9	.3 .2 .2	2,318 8,426 2,004	24.6 25.2 27.8	. 7 . 4 . 4
\$6 to h\$ years; Lass thin 12	5,736 7,537 2,628	2, 0 1, 2 , 5	.2 .1 .2	7,416 5,683 1,300	23.0 30.1 25.6	. 5 . 4 1. 0
OCCUPATION GROUP OF CSUAL JOB						
Professional, technical, ormanagerial. Sales or electral. Craits persons. Operatives. Luborers. Service.	12, 823 6, 018 10, 978 583 4, 506 3, 050	1.1 1.6 2.2 2.6 3.4 3.0	.1 ,2 ,1 ,2 ,3 ,4	7,673 14,509 272 4,760 426 6,142	17.9 25.6 15.2 19.2 54.9 23.5	. 4 . 3 . 2 . 5 . 4 . 4
AFF-CYCLE SINTEN]				
21 to 29 yeaps: Ever marfled	8,671 (SA) (SA) 5,349 (SA) (NA)	2.5 (Na) (Na) 5.2 (Na) (Na)	. 2 (NA) (NA) (NA) (NA)	7, 221 4, 184 3, 037 4, 235 878 3, 357	16.8 24.5 6.1 6.1 10.0 5.0	. 4 . 5 . 5 . 4 . 1, 1
30 to 44 years: Ever marries. With children. No children. Sever married. With children. No children.	16, 332 (NA) (NA) 1, 044 (NA) (NA)	1.5 (NA) (NA) 3.5 (NA) (NA)	.1 (NA) / (NA) / (NA) (NA) (NA)	11.462 8.769 3.093 878 138	29.2 32.4 20.0 7.0 7.1 2.2 4.9	.3 .4 .6 .8 1.9
4') to 04 years: Ever married	15. 343 558	1.3	. l . n	9.599 801	29.2 0.1	.4

NA Not available.



tpersons of Spanish origin may be of any race.

Table 7C. Mean Percent of Potential Work Years Spent Away From Work by Reason and Selected Characteristics: Full-Time Workers

į.		Males			femalics	
Characteristic		Mean percent o work years spe work for reas	nt news from		Menn percent work years spe work for reas	or may from
	Sumber (thousands)	Vallage	Standard orror	Number (Thomsands)	Value	Standari
Persons 21 to 65 years of age working tuffs that	40, 482	1.9	0.1	24, 379	18.0	0.1
RICE AND SOUNDSH ORIGIN [‡]					1	
2641. Black. opaci30 origin	35,95 2 1,984 2,106	1.7 3.8 2.0	.1 .3 .3	20,306 3,557 1,368	19.6 9.2 18.1	. 2 . 4 . 8
VG.F						
A to 25 years	12,279 14,811 13,401	2.9 1.3 1.4	.2 .1 .1	9,077 8,470 6,832	10.6 21.5 23.6	.1 .4 .4
YEARS OF SCHOOL COMPLETED		}				
Less than 12	9,290 22,125 9,067	2.2 2.9 1.5	.2 .1 .2	4,480 15,197 4,502	21. 4 18.5 13. 0	.5 .1 .5
ALL BY YEARS OF SCHOOL COMPLETED						
21 to 29 years- Less than 12	1.523 7.751	3.2 2.9	.4 .2	754 6.07 6	1 14.1 , 12.0	l. 1 . 4
to and over	2,996	2.8	-4	2.247	5.5	-5
P to 44 years: bess than 12	2,876	2.2	.3	1.561	23.4	.9
12 to 10	7,998 3,938	1.5 l.t	.1 .2	5,540 1,369	21.1 21.0	.5
15 to 94 years; Less thun 12	4,892	2.0	.2	2, 165	22 4	. 7
12 to 15	6,377 2,133	1.3	.1	3, 781 886	25 - 2 19 - 7	.5 1. (
OCCUPATION GROUP OF UNDAIL JOB						
trofessional, technical, wrranagerial, soles or element. Crafts persons	10,647 4,983 9,363 7,822 4,017	1.2 1.5 2.0 2.1 3.0		5.485 10.493 246 3.836 3181	13.8 20.4 13.5 17.8 29.1	.4 .3 1.8 .5 2.0
Senvice	2,599	2.6	.4	3,529	17.7	. د
LD) -CYCLE STARCS					Ì	
21 to 29 wars; Iver married. With children. Vo children. Aith child.en. Vi child.en.	7, 730 (SA) (SA) 4,540 (SA) (SA)	2.3 (SA) (SA) 3.9 (N1) (SA)	-2 (SA) (SA) (SA) -4 (SA) (SA)	5.378 2.953 2.425 3.699 717 2.982	14.1 21.9 4.5 5.5 11.2 4.1	.4 .6 .5 .4 :-1
30 to 46 years: ther married	14, 010 (NA)	1.5 (SA)	.1 (SA)	7.726 5.407	23.3 23.8	.4
No children. Vever marked. Aith Children. No children.	(SA) 801 (SA) (SA)	(SA) 2,6 (SA) (SA)	(SA) .5 (SA) ISA)	2,318 744 109 635	17.5 2.6 3.7 2.4	.7 .7 1.6
4) to 64 years: Even married.	12,907 495	1.4 2.8	.1 .6	6, 191 642 :	25.5 5.3	.4 1.0

NY Note in (11 differ)



Horsons of Spanish origin may be of any race.

APPENDIX A

Definitions and Explanations

Population coverage. This report includes the civilian noninstitutional population of the United States and members of the Armed Forces in the United States living off post or with their families on post. Other members of the Armed Forces are excluded.

Age. The age classification is based on the age of the person at his/her last birthday.

Race. Data are presented for two groups identified on the basis of race: White and Black.

Persons of Spanish origin. Persons of Spanish origin were identified by a question that asked for self-identification of the person's origin or descent. Respondents were asked to select their origin (and the origin of other household members) from a flashcard listing ethnic origins. Persons of Spanish origin were those who indicated that their origin was Mexican, Puerto Rican, Cuban, Central or South American, or some other Spanish origin. It should be noted that persons of Spanish origin may be of any race.

Years of school completed. Data on years of school completed were derived from the combination of answers to questions concerning the highest grade of school attended by the person and whether that grade was completed. The questions on educational attainment apply only to progress in "regular" schools. Such schools include graded public, private, and parochial elementary and high schools (both junior and senior), colleges, universities, and professional schools (whether day schools or night schools). Thus, regular schooling is that which may advance a person toward an elementary school certificate, a high school diploma, or a college, university, or professional school degree. Schooling in other than regular schools was counted only

if the credits obtained were regarded as transferable to a school in the regular school system.

Number of selected academic courses completed in high school. The courses that were asked about included (1) algebra, (2) geometry, (3) trigonometry or analytical geometry, (4) chemistry or physics, and (5) 2 or more years of a foreign language.

Employment status. Persons were considered to be employed if they had worked at a job or business at any time during the survey reference period (the 3 months preceding the month of interview).

Occupation group of usual job. Persons were classified according to the kind of work he/she had "done most often" during his/her adult life. The occupation groupings used in this report are based on the classification system used in the 1970 census.

Work interruptions. Classifications of persons by frequency, reason for, and duration of work interruptions was based on responses to the set of questions reproduced in appendix C.

Potential work years. Potential work years were defined to equal current age minus years of school completed minus 6 years.

Years spent away from work. Years spent away from work were defined to equal potential work years minus years spent away from work.

Rate of earnings. Hourly earnings were calculated by dividing total earnings for the 3 month reference period by the total number of hours worked during the period. The latter figure was calculated by multiplying the number of weeks worked by the usual hours worked per week.



APPENDIX B

Source and Reliability of the Estimates

SOURCE OF DATA

The estimates in this report are based on data obtained from Interview Period No. 3 of the 1979 Income Survey Development Program (ISDP). The universe for this survey was the civilian noninstitutionalized population and persons in the Armed Forces living off post or with their families on post. The estimation procedure used in this survey involved the inflation of weighted sample results to independent estimates of the civilian noninsitutionalized population of the United States by age, race, and sex. These independent estimates were based on statistics from the 1970 decennial census which were updated to 1979 using statistics on the strength of the Armed Forces, Because data from the ISDP in this report were collected over a 3-month period (August, September, and October 1979), weighting of each third of the sample was carried out independently, each third being inflated to one-third of the population control totals for each month of interview, respectively. The ISDP sample was selected from several sources. The main source of the sample was the Survey of Income and Education (SIE) conducted in 1976. The SIE was chosen as a sample source because the income data collected in this survey could be used to permit oversampling of low and high income households. In addition to the SIE portion, the general population sample also included some housing units reserved for current survey programs, such as the CPS, and a portion selected to represent new construction. Approximately 9,300 housing units were selected for the nationally representative, general population sample. The sample also included approximately 1,000 cases selected from administrative lists of disabled Supplemental Security Income (SSI) recipients and 1,000 cases selected from persons who had applied for the Basic Education Opportunity Grant (BEOG) program, Neither of the nationally representative SSI or BEOG samples were used to derive estimates shown in this report. The ISDP sample housing units were spread over 130 areas in 45 States and the District of Columbia.

About 900 occupied units were visited where interviews were not obtained because the occupants refused to voluntarily participate or were unavailable for some other reason. In addition, about 3 percent of the sample units visited were found to be vacant or otherwise not eligible for interview.

RELIABILITY OF THE ESTIMATES

Since the estimates in this report are based on a sample, they may differ somewhat from the figures that would have been obtained if a complete census had been taken using the same questionnaires, instructions, and enumerators. There are two types of errors possible in an estimate based on a sample survey: sampling and nonsampling. The standard errors provided for this report primarily indicate the magnitude of the sampling error. They also partially measure the effect of some nonsampling errors in response and enumeration, but do not measure any systematic biases in the data. The full extant of the non-

sampling error is unknown. The proportional amount of sampling and nonsampling error present in survey estimates is also unknown. Consequently, particular care should be exercised in the interpretation of figures based on a relatively small number of cases or on small differences between estimates.

Nonsampling variability. Nonsampling errors in surveys can be attributed to many sources, e.g., inability to obtain information about all cases in the sample, definitional difficulties, differences in the interpretation of questions, inability or unwillingness of respondents to provide correct information, inability to recall information, errors made in collection such as in recording or coding the data, errors made in processing the data, errors made in imputing values for missed data, and failure to represent all sample households and persons within sample households fundercoverage).

At present, not much is known about undercoverage in the ISDP. Since the ISDP sample frame is similar to that of CPS, the undercoverage is assumed to be similar. Undercoverage in the CPS results from missed housing units and missed persons within sample households. Overall undercoverage, as compared to the level of the 1970 decennial census, is about 5 percent. It is known that CPS undercoverage varies with age, sex, and race. Generally, undercoverage is larger for males than for females and larger for Blacks and other races than for Whites. Ratio estimation to independent age-sex-race population controls, as described previously, partially corrects for the bias due to survey undercoverage. However, biases exist in the estimates to the extent that missed persons in missed households or missed persons in interviewed households have different characteristics than interviewed persons in the same age-sex-race group. Further, the independent Population controls used have not been adjusted for undercoverage in the 1970 census, which was estimated at 2.5 percent of the population, with similar undercoverage differentials by age, sex, and race as in CPS.

Although the sample frames are similar, there are believed to be differences in coverage between ISDP and CPS. Survey estimates showed that the estimated number of families and husband-wife households was higher for the ISDP than the CPS. However, the CPS had higher estimates of the number of female-headed families and recipients of Aid to Families with Dependent Children (AFDC).

Sampling variability. The standard errors of the survey estimates are primarily measures of sampling variability, that is, of the variation that occurred by chance because a sample rather than the entire population was surveyed. The sample estimate and its standard error enable one to construct confidence intervals, ranges that would include the average result of all possible samples with a known probability. For example, if all possible samples were selected, each of these baing surveyed under essentially the same general conditions and using the same sample design, and if an estimate and its standard error were calculated from each sample, then—



28

- Approximately 68 percent of the intervals from one standard error below the estimate to one standard error above the estimate would include the average result of all possible samples.
- Approximately 90 percent of the intervals from 1.6 standard errors below the estimate to 1.6 standard errors above the estimate would include the average result of all possible samples.
- Approximately 95 percent of the intervals from two standard errors below the estimate to two standard errors above the estimate would include the average result of all possible samples.

The average estimate derived from all possible samples is or is not contained in any particular computed interval. However, for a particular sample, one can say with a specified confidence that the average estimate derived from all possible samples is included in the confidence interval.

Standard errors may also be used to perform hypothesis testing, a procedure for distinguishing between population parameters using sample estimates. The most common types of hypotheses appearing in this report are (1) the population parameters are identical or (2) they are different. An example of this would be comparing the proportion of men with work interruptions for familial reasons versus the proportion of women with work interruptions for familial reasons. Tests may be performed at various levels of significance, where a level of significance is the Probability of concluding that the parameters are different when, in fact, they are identical.

All statements of comparison in the text have passed a hypothesis test at the 0.10 level of significance or better, and most have passed a hypothesis test at the 0.05 level of significance or better. This means that, for most differences cited in the text, the estimated difference between parameters is greater than twice the standard error of the difference. For the other differences mentioned, the estimated difference between parameters is between 1.6 and 2.0 times the standard error of the difference. When this is the case, the statement of comparison will be qualified in some way; e.g., by use of the phrase "some evidence."

Standard error Parameters and their use. To derive standard errors that would be applicable to a wide variety of items and could be prepared at a moderate cost, a number of approximations were required. As a result, two parameters (denoted "a" and "b") were developed for use in calculating standard errors of estimated numbers and proportions. These Parameters have values of ~ .00020905 and 46,012, respectively. They may be used to directly calculate the standard errors for estimated numbers and percentages. Since a number of approximations were made, standard errors computed from these parameters provide an indication of the order of magnitude of the standard error rather than the precise standard error for any specific item. Methods for using these parameters for direct computations are given in the following sections.

Standard errors of estimated numbers. The approximate standard error of an estimated number can be obtained by using formula 1.

$$\sigma_{\chi} = \sqrt{ax^2 + bx} \tag{1}$$

Here x is the size of the estimate and a and b are the parameters mentioned above.

Standard errors of estimated percentages. The reliability of an estimated percentage, computed using sample data for both numerator and denominator, depends upon both the size of the Percentage and the size of the total upon which this Percentage is based. Estimated Percentages are relatively more reliable than the corresponding estimates of the numerators of the Percentages, particularly if the percentages are 50 percent or more. The approximate standard error, $\sigma_{\{X,P\}}$, of the estimated percentage can be obtained by the formula

$$\sigma_{(x,p)} = \sqrt{\frac{b}{x} \cdot p \ (100-p)}$$
 (2)

Here x is the size of the subclass which is the base of the percentage, p is the percentage $(0 \le p \le 100)$, and b is again the parameter mentioned above.

Illustration of the use of standard error parameters. Table A shows that in 1979 there were 9,245,000 ever-married males between 21 and 29 years of age who had ever worked. Using the values of a = -.00020905 and $b^{*}46.012$ in formula 1, the approximate standard error for the estimated number is

$$\sqrt{(-.00020905)(9,245,000)^2 + (46,012)(9,245,000)}$$

 $\stackrel{.}{=} 638,000$

The 68-percent confidence interval as shown by the data is from 8.607,000 to 9.883,000. Therefore, a conclusion that the average estimate derived from all possible samples lies within a range computed in this way would be correct for roughly 68 percent of all possible samples. Similarly, we could conclude that the average estimate derived from all possible samples lies within the interval 7.969,000 to 10.521,000 (using twice the standard error) with 95-percent confidence.

Standard error of a difference. For a difference between two sample estimates, the standard error is approximately equal to the square root of the sum of the squared standard errors of the estimates; the estimates can be of numbers, percents, ratios, etc. As such, the formula for the standard error of a difference is

$$\sigma_{(x-y)} = \sigma_x^{2} + \sigma_y^{2} \tag{3}$$

where σ_X and σ_Y are the standard errors of the estimates x and y. This will represent the actual standard error correctly for the difference between two estimates of the same characteristic in two different areas, or for the difference between separate and uncorrelated characteristics in the same area, if, however, there is a high positive (negative) correlation between the



estimates of the two characteristics, the formula will overestimate (underestimate) the true standard error.

Illustration of the computation of the standard error of a difference. Table A shows that the percentage of ever-married males aged 30 to 44 who had work interruptions due to inability to find work is 14.8, and the percentage for nevermarried males aged 30 to 44 is 34.4. Using formula 2, the standard errors of these percentages are

$$\sigma_{(17.825,000,14.8)} = \sqrt{\frac{46,012}{17.825,000} \cdot 14.8(85.2)} = 1.80$$

$$\sigma_{(1,281,000,34.4)} = \sqrt{\frac{46,012}{1.281,000} \cdot 34.4(65.6)} = 9.00$$

Therefore, the standard error of the estimated difference of 19.6 is

$$\sigma_{(x-y)} = \sqrt{(1.80)^2 + (9.00)^2} = 9.18$$

This means that the 68-percent confidence interval as shown by these data is from 10.42 to 28.78. Therefore, a conclusion that the average estimate derived from all possible samples lies within a range computed in this way would be correct for roughly 68 percent of all possible samples. The corresponding 95-percent

confidence interval is from 1.24 to 37.96. Since the interval does not contain zero, we can conclude with 95-percent confidence that the percentage of never-married males with work interruptions due to inability to find work was higher than that for ever-married males.

STANDARD ERRORS OF ESTIMATED MEANS AND REGRESSION COEFFICIENTS

Estimated standard errors for the means and regression coefficients are provided in the tables and do not need to be calculated by the user. The means and regression coefficients were calculated using SPSS, with the standard formulas used in this statistical package. A constant factor of 2 3 was applied to ordinary variance formulas which used sample weights, but did not fully take into account the sample design. This constant factor was obtained indirectly from direct estimates which were available for variances of proportions and numbers. The factor was selected to give a conservative approximation, which means the variance may be overestimated. This implies that if a difference is significant using this approximation, it would also be significant using the exact variance. However, some significant differences may be missed.



APPENDIX C

Facsimile of 1979 ISDP Questions on Work Interruptions

19 b. Thinking about the jobs that has worked at during 's adult life, hos always worked full time, most often worked full time, most often worked part time?	Always full time ~ SKIP to 20a Always full time ~ SKIP to 20a Most often full time Always part time Always part time
20a. We would like to know about's experiences with unemplayment. Have there been any times in's adult life when was out of work for 6 months or langer because could not find a job?	2 No _ SKIP to 214
b. How often has this happened — just once, σ few times, or many times?	3135 Just once 2 A few rimes 3 Many times
c. In what years was out of work for 6 months or longer because could not find a job? If more than 4 stretches, ask about most recent 4)	(3136) (1) 1 9 1 9 1
21a. Have there been times in 's adult fife when, instead of holding a jab, stayed home to take care of a family or home? (Include only periods lasting 6 months or longer)	3147) 1 (" Yes 2
b. In what years did stay home to take care of a fomily or home? Was there any other time when stayed home 6 months or longer to take care of a family or home?	(3142) (1) (1) (1) (1) (1) (1) (1) (1) (1) (1
c. Have there been times in's adult life when did not work for 6 months or langer because was sick or disable6?	2 [No _ SKIP to 22a
d. In what years didnot work for 6 months or langer becausewas sick or disabled? Was there any other time whendid not work for 6 months or longer becausewas sick or disabled?	(3148) (1) 1 9 1 9 1



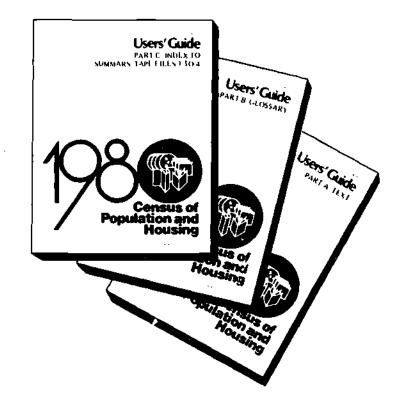
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