#### DOCUMENT RESUME

ED 060 027

TM 001 110

AUTHOR TITLE

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The Equating of the Scales for the Canadian and

American Scholastic Aptitude Tests.

INSTITUTION SPONS AGENCY Educational Testing Service, Princeton, N.J.

College Entrance Examination Board, New York, N.Y.;

Service for Admission to College and University,

Ottawa (Ontario).

REPORT NO PUB DATE NOTE

PR-71-24 Dec 71

39p.

EDRS PRICE

MF-\$0.65 HC-\$3.29

Comparative Statistics; Correlation; Cross Cultural DESCRIPTORS

Studies; \*Equated Scores; Item Analysis; \*Mathematics; Statistical Analysis; Testing; \*Test

Interpretation: \*Verbal Ability

IDENTIFIERS

Canadian Scholastic Aptitude Test; \*College Board

Scholastic Aptitude Test; CSAT; SAT

#### ABSTRACT

This study examines a common set of equating items for their appropriateness to both American and Canadian populations and, to determine the equations for converting scores from the verbal and mathematical scales of the 1971 form of the Canadian Scholastic Aptitude Test (the future standard continuing scales of the CSAT) to the corresponding standard continuing scales of the Scholastic Aptitude Test. (AG)





COLLEGE ENTRANCE EXAMINATION BOARD RESEARCH AND DEVELOPMENT REPORTS RDR-71-72, NO. 4

PROJECT REPORT
PR-71-24 DECEMBER 1971

# The Equating of the Scales for the Canadian and American Scholastic Aptitude Tests

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The Equating of the Scales for the

Canadian and American

Scholastic Aptitude Tests

bу

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and

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December, 1971

This project was sponsored jointly by the Service for Admission to College and University and the College Entrance Examination Board.

#### Introduction

The scores earned on the verbal and mathematical sections of the College Board Scholastic Aptitude Test (SAT) are reported on a scale that extends from 200 to 800, a scale originally established by setting the mean and standard deviation of candidates tested in April 1941 at the arbitrary values of 500 and 100, respectively. The transformation from raw (formula) scores to scaled scores is a linear one, which preserves all moments of the raw score distribution above the second, that is, it retains the shape of the raw score distribution. The SAT scale is a continuing scale; scores on each form of the test are equated to scores on the preceding forms with the effect that, within the limits of equating error, a score of 563, for example, earned on any one form of SAT-verbal has the same meaning -- represents the same level of ability -- as a score of 563 on any other form of SAT-verbal.

The scores earned on the verbal and mathematical sections of the Canadian Scholastic Aptitude Test (CSAT) of the Service for Admission to College and University (SACU) are also reported on a 200-800 scale. Here, too, the transformation from raw (formula) scores to scaled scores is a linear one. Unlike the SAT-verbal scale, however, which is equated back to April 1941, the 1971 linear scale of the CSAT-verbal was defined on the form actually administered in January 1971 by setting the minimum raw score at 200 and the maximum raw score at 800. The two earlier forms of the CSAT-verbal (introduced in 1969 and in 1970) were scaled in the same way, independently for each form and not equated to each other. However, scores on all future CSAT-verbal forms will be equated to the CSAT-verbal scale for January 1971. Thus, the January 1971 scale becomes the basic reference scale for CSAT-verbal and corresponds directly to the April 1941 reference scale used for SAT-verbal.

Statements similar to these just made for the verbal tests are also appropriate for SAT-mathematical and CSAT-mathematical, respectively.



Although both the CSAT and the SAT are very similarly designed tests of verbal and mathematical aptitude, and although their scales are similar in appearance, it is clear from the foregoing, and also from the fact that no effort has ever been made to link the two systems together, that equalappearing scores in the two systems -- say, again, 563 on CSAT-verbal and SAT-verbal -- might have quite different meanings and reflect quite different levels of ability. Accordingly, the present study was undertaken to examine a common set of equating items for their appropriateness to both populations and, following that examination, to determine the equations for converting scores from the verbal and mathematical scales of the 1971 form of the CSAT (the future standard continuing scales of the SAT.

#### Method

The test forms used in the present equating study were the form of the SAT administered on January 9, 1971 and the form of the CSAT administered on January 20, 1971. As an introduction to the design of the study it will be useful to review the administrative arrangements and content of the two tests. Both tests were administered in five separately timed sections: two operational verbal sections, two operational mathematical sections, and one nonoperational section administered in a time slot reserved for the presentation of pretest or equating items. Table 1 describes the item lengths and timing of the two tests and the order of appearance of the sections in the forms of the tests studied here.

The SAT-verbal contains four item types: antonyms, analogies, sentence completion, and reading comprehension. Typically, all four item types are represented in both of the timed sections of the SAT. The CSAT-verbal items



are of the same type as those in the SAT. Unlike the SAT, however, one of the two separately timed verbal parts of the CSAT contains exclusively discrete items; the other contains exclusively reading comprehension items.

Table 1

Administrative Arrangement and Content

of the CSAT and the SAT

		SAT		1-		CSAT	
Sec	tion	No. of Items	Timing (Min.)	Sec	etion	No. of Items	Timing (Min.)
I	Verbal	40	30	I	Math	20	30
II	Verbal	50	45	II	Verbal	55	30
III	Nonopnl	Variable	30	III	Nonepnl	Variable	30
IV	Math	35	45	IV	Math	20	30
v	Math	25	30	v	Verbal	50	3ō

Except for the data sufficiency items (the SAT has 18; the CSAT has none) the mathematical items in the SAT and CSAT are quite similar. The timing, however, is different. The SAT gives 1 1/4 minutes per mathematical item; the CSAT gives 1 1/2 minutes.

As indicated above, the nonoperational section of both tests (the third time slot in both the SAT form and the CSAT form used in this study) is used for pretesting items and also, in the SAT, for equating successive forms of the tests. In any formal administration of either test more than one non-operational section may be administered in that time slot, with different students taking different nonoperational sections. In order to assure that randomly equivalent groups of students take one such section or another, the test booklets are distributed in "spiralled" arrangement. Thus, if m non-



operational sections are to be administered, every mth student takes the same such section. In the administration of the SAT, for example, it is usual for m to be as high as 25 or 30.

In preparation for the present study arrangements were made between the Service for Admission to College and University (SACU) and the College Entrance Examination Board (CEEB) to exchange 30-minute operational sections and for each of the two organizations to administer the sections of the counterpart organization in its cwn nonoperational position. Thus, in addition to the usual 30-minute pretests and 30-minute equating sections normally assigned to the nonoperational position in the January 1971 SAT, Educational Testing Service (ETS) also administered (in behalf of CEEB) one of the two 30-minute operational CSAT-werbal sections and one of the two 30-minute operational CSAT-mathematical sections taken from the January 1971 CSAT. Correspondingly, in addition to the usual 30-minute pretest sections normally assigned to the nonoperational position in the January 1971 CSAT, SACU also administered the 30-minute operational SAT-verbal section and the 30-minute operational SAT-mathematical section taken from the January 1971 SAT.

With this arrangement a full 60 minutes of testing material in the verbal area -- representing a total of 95 items -- was administered for purposes of equating to subsamples of both populations in connection with their operational testing. Thirty minutes of this material (40 items) was operational in the SAT (Section I) -- i.e., contributing to the student's formal score -- and nonoperational in the CSAT; thirty additional minutes (55 items) was operational in the CSAT (Section II) and nonoperational in the SAT. Under the same sort of arrangement a full 60 minutes of testing material in the mathematical area -- representing a total of 45 items -- was administered, for the purposes of



equating, to subsamples of both populations in connection with their operational testing. Thirty minutes of this material (25 items) was operational in the SAT (Section V) and nonoperational in the CSAT; thirty additional minutes (20 items) was operational in the CSAT (Section I) and nonoperational in the SAT. As indicated above, all nonoperational sections appeared in Section III of both tests. The operational SAT-verbal section is referred to as SAT-Va and the operational CSAT-verbal section is referred to as CSAT-Vc. The total of the two sections is referred to as Vt. Corresponding notation is used in the SAT- and CSAT-mathematical tests: Ma, Mc, and Mt.

The mathematical sections that were exchanged between the two organizations were quite similar in content, representing a mixture of arithmetic reasoning, algebra, and geometry. The CSAT-verbal section that was administered along with the SAT contained sentence completion items, analogies and opposites only. The SAT-verbal section administered along with the CSAT contained these same item types, but in addition, included reading comprehension items as well.

The procedure followed in developing the conversions between the two testing systems called for two general steps: 1) an examination of the 95 common items in the verbal area and the 45 common items in the mathematical area in order to identify items that appeared to have somewhat "different meaning" for American and Canadian students, and to eliminate such items from the group of common items to be used for equating the two tests; 2) the development of the necessary statistics, based on the remaining items, for equating the verbal and mathematical sections of the two tests.

#### Examination of Item Data

In order to carry out the first of these two steps, two mutually exclusive random samples of the SAT candidates who took Vc were drawn (these

candidates also took Va as part of the operational SAT), and similarly, two mutually exclusive random samples of the CSAT candidates who took Va were drawn (these candidates also took Vc as part of the operational CSAT). Four additional samples were selected for mathematical, and in the same way: mutually exclusive random samples of the SAT candidates who took Mc were drawn (these also took Ma operationally), and similarly, two mutually exclusive random samples of CSAT candidates who took Ma were drawn (tnese also took Mc operationally). Each of the individual American samples consisted of 450 cases; the Canadian samples, however, ranged in size from 405 to 420. four SAT samples are referred to as AVI, AV2, AMI, and AM2; the four corresponding CSAT samples are referred to as CVI, CV2, CMI, and CM2. In addition, samples AVT, CVT, AMT, and CMT were formed. (The letters A and C denote American and Canadian students respectively; the letters V and M denote tests for which these data were to be analyzed: verbal and mathematical; the numerals 1 and 2 are used only to distinguish between random subsamples of students drawn from the same population and taking the same battery of tests. The letter T denotes a combined sample; thus AVT = AV1 + AV2.)

The first step in the analysis -- the examination of the performance of the American and Canadian students on the same items -- involved the use of a technique that will bear detailed description. In comparing two samples of individuals in this way, an item analysis is carried out for each of the two groups and a plot is made of the points represented by the paired  $\Delta$ -values,  $\Delta$  vs  $\Delta$  \*, one point for each of the items (i) under consideration for

<sup>\*</sup>Delta ( $\triangle$ ) is an index of item difficulty for an item.  $\triangle$  = 4z + 13 , where z is a normal deviate corresponding to p, the proportion of examinees answering the item correctly.



which A-values are available. The plot of these points normally appears in the Yorm of an ellipse extending from lower left to upper right, and if the samples are drawn from the same type of population, the scatterplot of these points is a long, narrow one, often representing a correlation as high as .98 or .99. When the samples are different in level, the points will still fall in a long, narrow ellipse, but displaced vertically or horizontally, depending on which group is the abler one. Even when the groups differ in degree of dispersion, the points will still fall in the same type of ellipse, but the ellipse will be tilted at an angle more or less steeper than 450, depending on which sample is the more dispersed. However, when the groups are different in type, the item difficulties will not fall in precisely the same rank order for the two groups, and the correlation represented by the delta-points will be lower than .98 or .99. The items falling at some distance from the plot may be regarded as contributing to the item x group interaction. They are the items that are especially more difficult for one group than for the other, relative to the other items.

The method developed for summarizing the significant features of each plot involves the determination of the major axis of the ellipse represented by the plot and the calculation of the perpendicular distance  $(D_i)$  from each point to that line. The standard deviation of the distribution of these distances is a function of the item x group interaction. As indicated above, the correlation coefficient represented by the ellipse represents the degree to which the items have the same rank order of difficulty in the two groups -- also a representation (inversely) of the item x group interaction.

The equation used for the major axis of the ellipse is a linear one, h = Pg + Q, where



$$P = \frac{(\sigma_{h}^{2} - \sigma_{g}^{2}) \pm \sqrt{(\sigma_{h}^{2} - \sigma_{g}^{2})^{2} + 4r_{gh}^{2}\sigma_{g}^{2}\sigma_{h}^{2}}}{2r_{gh}\sigma_{g}\sigma_{h}}$$

and

$$Q = M_h - PM_g$$

(It is recalled that the variables, g and h, are, respectively, the delta values for the two groups under consideration.) The formula for the perpendicular distance, D, of each point, i, in the plot to the line is given as:

$$D_{i} = \frac{Pg_{i} - h_{i} + Q}{\sqrt{P^{2} + 1}}$$

In preparation for the examination of the delta plots, six item analyses were carried out for the verbal tests, one for each of the four separate verbal samples (AVI, AV2, CVI, and CV2) and one for each of the combined samples (AVT Similarly, six item analyses were carried out for the mathematical tests, one for each of the four separate mathematical samples (AML, AM2, CML, and CM2) and one for each of the combined samples (AVT and CVT). Tables 2 and 3 summarize the data of the item analyses. Table 2, which summarizes the data for the verbal tests, makes it clear that the mean test scores for the Canadian examinees are higher by 1/5 to 1/4 of a standard deviation than are the means for the American examinees. This difference may be observed on the items that were originally Canadian, and also on the items originally American. As expected, the item data support this observation; the mean deltas for the American samples are higher than the mean deltas for the Canadian samples and indicate that the items are more difficult for the Americans than for the Canadians. It is also observed in both the Va and Vc items that the Canadian samples are more homogeneous than the American samples. This finding is also supported by the standard



Table 2

Means and Standard Deviations of Raw Scores and Items

for American and Canadian Samples

Verbal Tests

	;										
		ials			.131	.134	.134		.115	.111	.110
n=55	Data	Biserials	Mean		.518	.534	.522		.555	.534	.533
(Vc);	Item Data	as	2		2.91	2.85	2.86		3.05	3.15	3.08
Canadian Items (Vc); n=55		Deltas	Mean		13.36	13.61	13.51		13.17	13.00	13.13
Canadi	در	Data	S.D.		11.02	11.31	11.18		10,96	10.15	10.61
	Test	Score Data	Mean		19.89	18.77	19.53		21.64	22.08	21.78
		als	S.D.		.131	.132	.126		901.	.095	160.
1=40	ata	Biserials	Mean		264.	.482	.487		.455	.443	.458
(va); $n=40$	Item Data	1 ~ 1	S.D.		2.90	2.79	2.83		3.05	3.08	3.06
American Items		Delts	Mean		15.05	13.23	13.16		12.93	12,85	12.91
Ameri	1	Data	S.D.		8.08	76.7	8 で す。		7,00		
	Test	Score Data	Mean		15,44	14.67	15.06		16.81	17.16	16.92
		No. of	Cases		4.50	450	006		405	410	*677
			Sample	American	AVT	Z.1.∇	AVT	Canadian	ĽΛIJ	GW?	CVI

\*Because of the difficulty in establishing the unique identities of some cases in Samples CV1 and CV2, 40 cases were dropped in the process of combining these samples to form Sample CVT.

Table 3

Means and Standard Deviations of Raw Scores and Items

for American and Canadian Samples

Mathematical Tests

								•			
als	S.D.	277	c)n.	.111	.081			.128	.097	רוו	-
ata Biseri	Mean	· ·	040	-,602	.621			. 599	. 622	77.79	- -
Item D	S.D.	, i	2.5I	2.57	去。			26.3	2.86	0 04	- O V
Delta	Mean	. 1	13.56	13.53	13.54			12.50	12.36	ا ا ا	1Z:47
င် )ata	S.D.		5.35	7.98	5.17			4.61	4.67	0	80.
Test Score I	Mean		6,69	£9•9	<b>6.</b> 68			9.03	9.27	1	6. 6.
ν. Γ	S.D.		129	.103	.110			.093	104	,	060
ata Biseria	Mean		.581	.580	. 580			.538	. 565		• 566
Item De	S.D.		2.39	2.55	2.44			3,28	3.05	\ \ \	3.17
1	Mean		13.15	12.98	13.07			12,23	12.14	 	12.16
<b>ئ</b> ا	Data S.D	•	5.93	5.71	5.82			4.68	14.86	)	4.82
Tes	Score		9.90	10.25	10.01			91.61			775* 12.31 4.82
	No. of Cases		450	450	006			1.20		\ H H	*677
	Sample	American	AMI	AMO	AMT		Canadian	LIMU.	<u> </u>	אוייס	CMI
	Test Item Data Test Ltem Data Deltas Score Data Deltas	Test Item Data Test Item Data	Test Test Item Data Test Item Data Item Data Item Data  No. of Score Data Deltas Biseria.  Cases Mean S.D. Mean S.D. Mean S.D. Mean S.D. Mean S.D.	Test         Item Data         Item Data         Item Data         Item Data         Item Data         Item Data           No. of Score Data         Score Data         Deltas         Biserial         Score Data         Biserial         Biserial         Biserial         Biserial         Biserial         Biserial         Mean         S.D.         Mean         S.D.         Mean         S.D.         Mean         Biserial         Mean         Biserial         Mean         S.D.         Mean         Biserial         Mean         S.D.         Mean         Biserial         Mean         S.D.         Mean	No. of Score Data         Score Data         Deltas         Biserials         Score Data         Deltas         Biserials           No. of Score Data         Score Data         Deltas         Biserials         Score Data         Deltas         Biserials           Mean         S.D.         Mean         S.D.         Mean         S.D.         Mean         S.D.         Mean           Mean         5.97         12.15         2.39         .581         .129         6.69         5.35         13.55         2.51         .640           450         10.25         5.71         12.98         2.55         .580         .103         6.67         4.98         13.55         2.57         .602	Test         Item Data         Item Data         Item Data         Test         Item Data         Item Dat	No. of Score Data         Score Data Mean         Item Data S.D.         Item Data Biserials Mean         Test Score Data S.D.         Test Deltas Deltas Biserials Mean         Biserials Score Data Mean         Test Deltas Deltas Biserials Mean         Biserial Biserial Biserial Biserials Mean         S.D.         Mean         S.D.         Mean         Biserial Biserial Biserial Biserial Biserial Biserial Biserial Biserial Biserial Mean         Biserial	Moson         Test         Item Data         Item Data         Test         Item Data         Item Data           Gases         Mean         S.D.         Mean         S.D.         Mean         S.D.         Mean         S.D.         Mean         S.D.         Mean         Biseria.           Mean         S.D.         Mean         S.D.         Mean         S.D.         Mean         S.D.         Mean         Biseria.           Mean         S.D.         S.D.         S.D.	Test Test Test Test Test Test Test Test	Test Test Tem Data Score Data No. of Score Data Nean S.D. Nean S.D	No. of Score Data   Item Data   Score Data   Score Data   Score Data   Biserial   Score Data   Biserial   Score Data   Biserial   Score Data   Biserial   Biserial   Score Data   Biserial   Biserial   Score Data   S.D.   Mean   S.D.   Mean

\*Because of the difficulty in establishing the unique identities of some cases in Samples CM1 and CM2, 60 cases were dropped in the process of combining these samples to form Sample CMT.



deviations of deltas, which are <u>larger</u> for the Canadians than for the Americans.\*

To what extent these various observations can be generalized beyond the observed samples it is impossible to say. At the very most, the data indicate that the CSAT group tested in January 1971, from which the CSAT samples were randomly drawn, is more able on the verbal abilities measured by the Va and Vc items than the SAT group tested in January 1971. It is beyond the scope of these data to give evidence on the extent to which these groups are differently representative of their parent populations.

The data on the biserial coefficients are interesting in that they indicate that items originally American are slightly more discriminating for the Americans ( $\overline{r}_{bis}$  = .49 for AVT) than for the Canadians ( $\overline{r}_{bis}$  = .46 for CVT), while items originally Canadian are slightly more discriminating for the Canadians ( $\overline{r}_{bis}$  = .55 for CVT) than for the Americans ( $\overline{r}_{bis}$  = .52 for AVT). This finding in turn suggests the possibility that the items in the two tests may be homogeneous in content, but specific to the students for whom the items were intended.

Table 3, which summarizes the data for the mathematical tests, repeats some of the findings reported in Table 2. As in the verbal tests, the Canadian students score higher than their American counterparts by almost half a standard deviation on both sets of items, those originally Canadian as well as those originally American. As expected, the means of the item deltas confirm these findings; the mean deltas for the American students are higher than those for the Canadian students, indicating that the items are more difficult

<sup>\*</sup>In general, standard deviations of item difficulties are larger for more homogeneous groups and smaller for more heterogeneous groups.



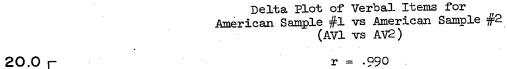
for them than for the Canadian students. Also, as in the verbal tests, the Canadians are more homogeneous as a group than the Americans on both the Ma and Mc subtests. This finding is also supported by the item data, which show a larger standard deviation of deltas for the Canadians. The data on item biserials are less clear for the mathematical items than they are for the verbal items. Although the item biserial correlations for mathematical are higher for the Americans on the items originally American, they are about equal for the two national groups on the items originally Canadian.

Figures 1-3 present the delta plots for the verbal items. Figure 1 is the plot of all 95 items for one random sample of American students versus the other (AV1 vs. AV2); Figure 2 is the plot for the same set of 95 items for one random sample of Canadian students versus the other (CV1 vs. CV2), and Figure 3 is the plot for the same set of items for the combined sample of Americans versus the combined sample of Canadians (AVT vs. CVT). It is clear from these figures that the cross-national plot (Figure 3) shows a greater dispersion than does either of the within-national plots (Figures 1 and 2). These observations are supported by the correlations represented by these plots (.976 for AVT vs. CVT, as compared with .990 for AV1 vs. AV2 and .991 for CV1 vs. CV2) and also by the values of  $\sigma_D$  (.458 for AVT vs. CVT, as compared with .287 for AV1 vs. AV2 and .296 for CV1 vs. CV2).

Figures 4-6 present the corresponding delta plots for the 45 mathematical items: Figure 4 is a plot of the deltas for one random sample of Americans versus another (AMI vs. AM2); Figure 5 is a plot of the deltas for one random sample of Canadians versus another (CM1 vs. CM2), and Figure 6 is a plot of the deltas for the combined American sample versus the combined Canadian sample (AMT vs. CMT). Here too the cross-national plot (Figure 6) is more dispersed



Figure 1



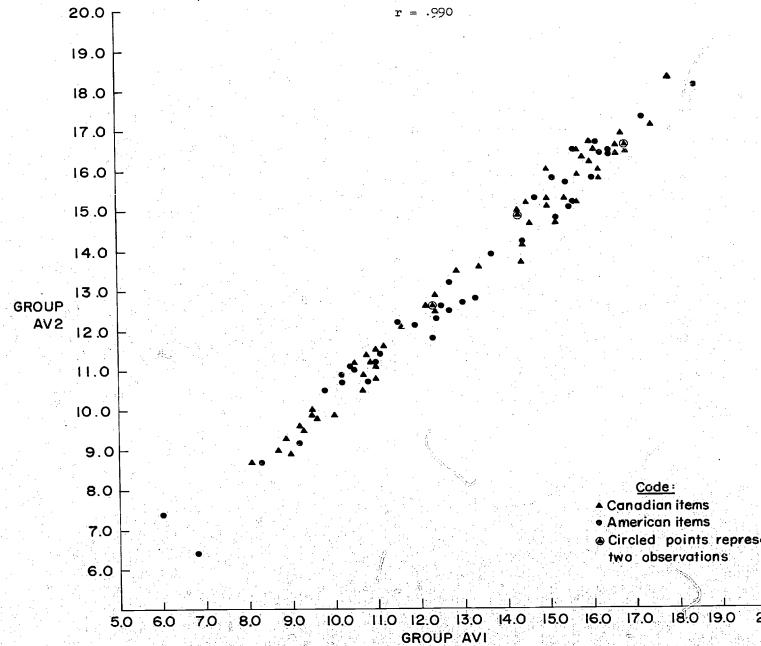




Figure 2

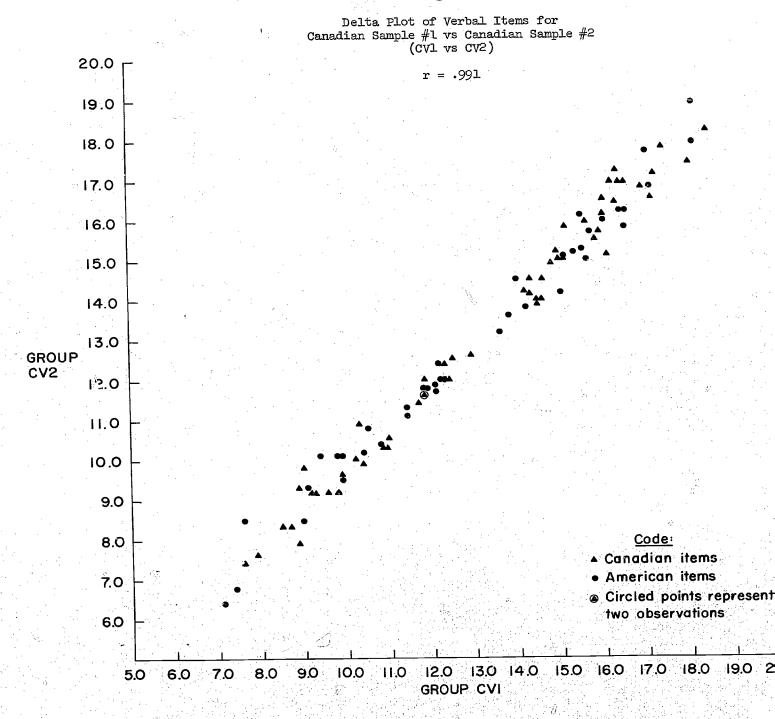




Figure 3

Delta Plot of Verbal Items for Total American Sample vs Total Canadian Sample (AVT vs CVT)

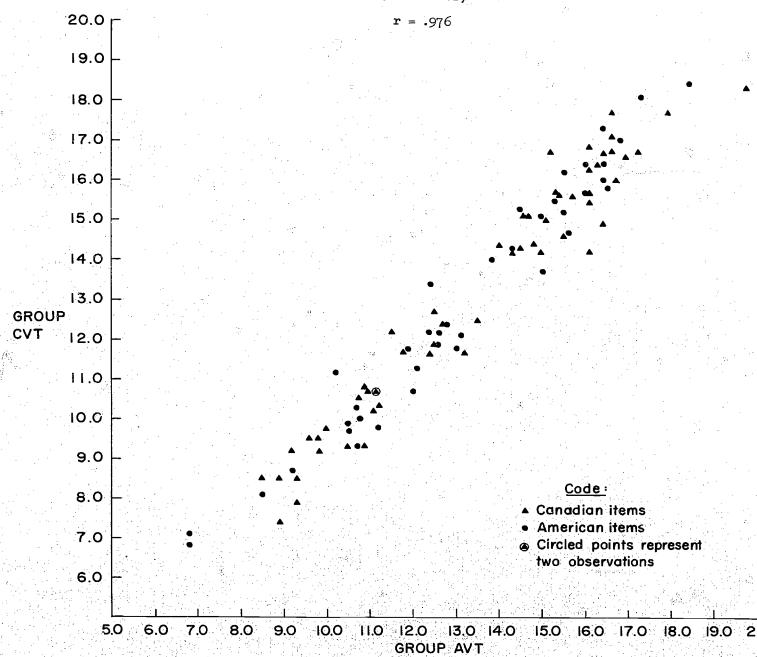


Figure 4

Delta Plot of Mathematical Items American Sample #1 vs American Sample #2 (AML vs AM2)

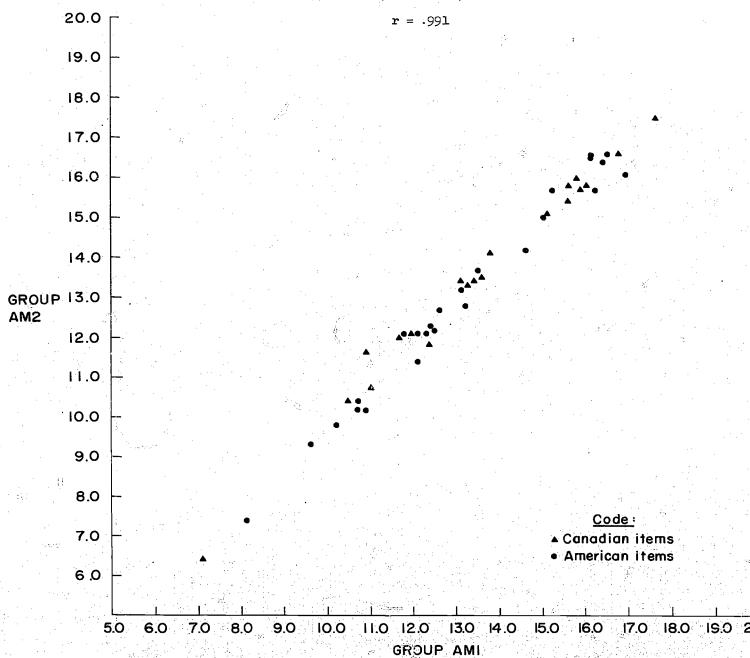




Figure 5

Delta Plot of Mathematical Items
Canadian Sample #1 vs Canadian Sample #2
(CMl vs CM2)

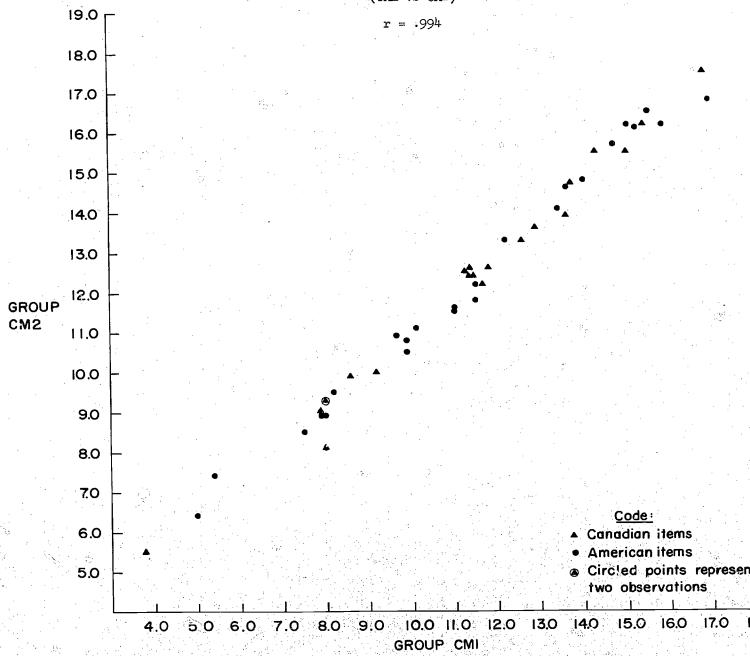
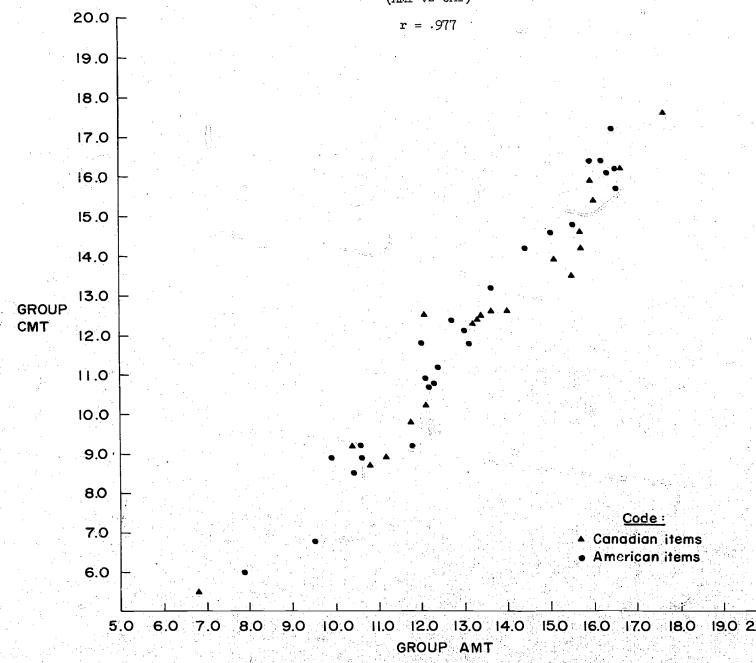




Figure 6

Delta Plot of Mathematical Items
Total American Sample vs Total Canadian Sample
(AMT vs CMT)



than either of the within-national plots (Figures 4 and 5), and, as expected, the statistics based on the plots bear out these observations. (The correlation for AMT  $\underline{vs}$ . CMT is .977, as compared with the correlations of .991 for AMI  $\underline{vs}$ . AM2 and .994 for CMI  $\underline{vs}$ . CM2; the  $\sigma_D$  for AMT  $\underline{vs}$ . CMT is .415, as compared with the  $\sigma_D$  of .233 for AMI  $\underline{vs}$ . AM2 and .242 for CMI  $\underline{vs}$ . CM2.) The six distributions of D-values, corresponding respectively to Figures 1-6, are given in Table 4.

Although these differences are clear and indisputable, certain qualifications are in order. The two sets of within-national samples, it is recalled, were selected randomly, and the differences between them are as small as the method of random sampling and the sample size permit. If systematic samples (e.g., samples drawn from somewhat different geographical areas within the United States or within Canada) had been drawn, we would expect the resulting delta plots to be considerably more dispersed than the observed random within-national samples. In some instances, even larger differences in dispersion of deltas could result from geographical sampling for within-national groups than for the random cross-national samples observed here. Therefore, to draw the general conclusion that between-national comparisons are more dispersed than within-national comparisons is not warranted by these data.

In any case, the dispersions observed in the cross-national plots are too small to warrant the conclusion that the American items, for example, are inappropriate for the Canadians, or vice versa. However, since this issue is of itself not central to the purpose of the present study -- which was to find the equations to be used for converting the scaled scores on the SAT to the scale for the CSAT -- it was not pursued in detail. Instead, the distributions of D-values for the cross-national plots (AVT vs. CVT and AMT vs. CMT) were



Table 4 Distributions of D-Values for the Item Delta Plots

	Plots of	Verbal	Items	Plots of	Mathemati	cal Items
	AVl	CVl	AVT	AML	CML	AMT
D-Value	vs. AV2	vs. CV2	vs. CVT	vs. AM2	vs.	vs. CMT
	AV2	<u> </u>			<u> </u>	AMERICAN TO SERVICE OF THE SERVICE O
1.2	* ************************************		1	# - 1		
.1.1			1			
1.0			1			1
0.9			0			0
0.8	, A		3			1
0.7		1	O .		1	1
0.6	1	1	5	1	0	0
0.5	5	2	6	0	1	2
0.4	5	6	4	2	1	5
0.3	11	14	5	2	4	
0.2	9	15	8	8	3	3
0.1	8	16	6	5	10	5
0.0	14	17	10	9	4	6
0.1	12	8	8	6	7	6
0.2	9	5	13	4	9	1
0.3	12	7	14	6	3	5
0.4	4	4	6		.1	0
0.5	2	5	5	0	1	2
0.6	1	2	1	1.00		1
0.7	2	1	3			2
0.8		1	2			1
0.9			1			<b>,</b> O
1.0			1			. 0
1.1			1			1
No. of Items	<u> </u>	– 95 <del>–</del>			<del></del>	
Mean		- 0.0 —	<u> </u>		0.0	
S.D.	. 287	.296	<b>.</b> 458	- 233	.242	•415
Plot Correlation	•990	•991	<b>.</b> 976	•991	•994	•977
					totals of Automotive to	

examined to exclude from the set of items used in equating the CSAT to the SAT those items that were considered to be relatively inappropriate to Americans or inappropriate to Canadians. These items were taken to be the "outliers" (i.e., those with high positive or high negative D-values), contributing heavily to the item x nationality interaction. Accordingly, the items removed from the verbal equating sections were those for which  $-0.55 \ge D_1 \ge 0.55$ ; the items removed from the mathematical equating section were those for which  $-0.52 \ge D_1 \ge 0.52$ . This process resulted in the removal of 20 of the 95 Verbal items (9 Canadian and 11 American items) and 10 of the 45 Mathematical items (4 Canadian and 6 American items).

#### Equating

All available Canadian examinees who had taken Section Va as their nonoperational section and all available Canadian examinees who had taken Section Ma as their nonoperational section, exclusive of those who had been part of the item analysis samples (CV1, CV2, CM1, or CM2), were selected for the equating process. This selection resulted in Canadian samples of 3493 for the verbal equating and 3315 for the mathematical equating. Corresponding American samples of 3495 and 3320 (approximately the sizes of the Canadian samples) were selected at random from the available American examinees who had taken Section Vc or Mc, exclusive (as in the choice of the Canadian samples) of all those who had been part of the item analysis samples (AV1, AV2, AM1 or AM2).

The method used for equating the Canadian and American scales involved first the determination of the equation relating the raw (R - W/4) scores on the operational American SAT to the raw (R - W/4) scores on the operational Canadian



CSAT and substituting into that equation the relationship between raw scores and scaled scores for each of the two testing programs to yield a final relationship between American scaled scores and Canadian scaled scores.

The initial conversion between raw scores on the two systems required first an estimation of the mean and variance for the combined American-Canadian sample on each of the two tests in accordance with a method developed by Tucker (Angoff, 1971, p. 580), as follows:

$$\hat{M}_{x_{t}} = M_{x_{\alpha}} + b_{xv_{\alpha}}(M_{v_{t}} - M_{v_{\alpha}}) , \qquad (1)$$

$$\hat{s}_{x_{t}}^{2} = s_{x_{\alpha}}^{2} + b_{xv_{\alpha}}^{2} (s_{v_{t}}^{2} - s_{v_{\alpha}}^{2})$$
, (2)

$$\hat{M}_{y_{t}} = M_{y_{\beta}} + b_{yv_{\beta}} (M_{v_{t}} - M_{v_{\beta}}) , \qquad (3)$$

and

$$\hat{s}_{y_{t}}^{2} = s_{y_{\theta}}^{2} + b_{yv_{\theta}}^{2} (s_{v_{t}}^{2} - s_{v_{\theta}}^{2}) , \qquad (4)$$

where

x or X = the test (CSAT) taken by Group  $\alpha$  (Canadian),

y or Y =the test (SAT) taken by Group  $\beta$  (American),

v = the equating test, i.e., scores on the items taken by both Group  $\alpha$  and Group  $\beta$  and used in the equating,

and

 $t = the combined group, <math>\alpha + \beta$ .

The notation,  $b_{xv}$ , represents the usual coefficient of regression of variable x on variable v:  $b_{xv} = r_{xv}\sigma_x/\sigma_v$ . (Similarly,  $b_{yv} = r_{yv}\sigma_y/\sigma_v$ .) The estimated values,  $\hat{M}_{x_t}$ ,  $\hat{M}_{y_t}$ ,  $\hat{s}_{x_t}$ , and  $\hat{s}_{y_t}$ , were then substituted in the equation,



$$\frac{\mathbf{Y} - \hat{\mathbf{M}}_{\mathbf{y}_{t}}}{\hat{\mathbf{s}}_{\mathbf{y}_{t}}} = \frac{\mathbf{X} - \hat{\mathbf{M}}_{\mathbf{x}_{t}}}{\hat{\mathbf{s}}_{\mathbf{x}_{t}}} , \qquad (5)$$

to yield the equation,

$$Y = aX + b , (6)$$

converting raw scores on Test X to the scale of the raw scores on Test Y.

In this equation,  $a = \hat{s}_{y_{t}} / \hat{s}_{x_{t}}$  and  $b = \hat{M}_{y_{t}} - a\hat{M}_{x_{t}}$ .

In order to calculate the estimated values given in equations (1) to (4) for the verbal tests, the correlation between the operational test and the 75-item equating section,\* as well as the related means and standard deviations, were prepared for each of the two verbal samples, the sample consisting of the 3495 American students and the sample consisting of the 3493 Canadian students. These statistics are given in Table 5. Distributions of the operational verbal and equating verbal tests for the SAT and CSAT are given in Table 6.

The application of the data in Table 5 to equations (1) to (4) resulted in the following estimated values:

$$\hat{M}_{x_t} = 30.6009$$
,  $\hat{M}_{y_t} = 35.5304$ ,  $\hat{s}_{x_t}^2 = 237.0424$ ,

and

$$\hat{s}_{y_{t}}^{2} = 289.9611$$

<sup>\*</sup>Recall that 20 verbal "outlier" items had been removed from the original group of 95 items administered to both the American and the Canadian examinees.



Table 5

Correlations Between Operational and Equating Tests

for American and Canadian Samples

## Verbal Tests

	American Sample	Canadian Sample
No. of Cases	3495	<b>3</b> 493
Operational Test*  Mean Std. Dev.	31.8710 17.5004	33.9969 14.1405
Equating Test**  Mean Std. Dev.	25.8489 15.1437	32•5663 13•3905
Correlation: Opnl. vs. Equating	•9431	•9572

<sup>\*</sup>The operational SAT-verbal contained 90 items: the operational CSAT-verbal contained 75 items.

<sup>\*\*</sup>The verbal equating test consisted of 75 items (after the removal of 20 "outlier" items).

Table 6
Frequency Distributions for Operational and
Equating Sections of the SAT and CSAT

# Verbal Tests

	American	Sample	Canadian	Sample
Raw (Formula) Score	Operational SAT	Equating Section	Operational CSAT	Equating Section
84 - 85 - 87 - 87 - 87 - 88 - 88 - 88 - 88 - 88 - 88 - 77 - 18 - 88 - 77 - 18 - 1	2 3 11 20 29 56 67 118 146 145 148 145 148 205 202 214 223 204 175 161 157 1157	7 9 19 28 35 52 88 97 181 201 218 244 267 251 254 221 29 210 210 85 29 31	1 10 30 50 54 65 115 137 169 207 210 253 261 265 325 270 235 229 156 158 103 72 57 43 96 3	1 10 14 35 45 57 79 110 127 203 206 297 318 298 297 193 162 110 88 51 30 14 1
No. of Cases	3495	3495	3493 33 <b>.</b> 9969	3493 32•5663
Mean	31.8710	25.8489	55.9969 14.1405	13.3905
Std. Dev.	17.5004	<b>1</b> 5 <b>.1</b> 4 <b>3</b> 7	T4. T402	エン・フラリン



These values were applied to equation (5) to yield the equation,

$$Y = 1.1060 X + 1.6857$$
 (7)

which permitted the conversion of scores from the raw score scale of the CSAT-verbal to the raw score scale of the SAT-verbal. The derivation of the conversion from the CSAT-verbal reporting scale to the SAT-verbal reporting scale is developed as follows:

The linear equation,

$$S_{c} = AX + B , \qquad (8)$$

is the equation by which raw scores on the form of the CSAT used in January 1971 (X) are converted to the CSAT reporting scale (Sc). Similarly, the linear equation,

$$S_{3} = A'Y + B' \qquad , \tag{9}$$

is the equation by which raw scores on the form of the SAT used in January 1971 (Y) are converted to the SAT scale (Sa). Expressing equation (8) in terms of  $X \left[ X = \frac{S_c - B}{A} \right]$  and equation (9) in terms of  $Y \left[ Y = \frac{S_a - B'}{A'} \right]$ , and substituting in equation (6) results in the equation,  $\frac{S_a - B'}{A'} = a \left[ \frac{S_c - B}{A} \right] + b$ , which, when simplified, becomes

$$S_a = \frac{aA'}{A} S_c + A'b + B' = \frac{aA'B}{A}$$
 (10)

Equation (10) is a linear equation with slope equal to aA'/A and intercept equal to A'b + B' -  $\frac{aA'B}{A}$ , and may be used to convert verbal or mathematical scores from the January 1971 converted score scale for the CSAT to corresponding scores on the SAT scale.



In order to derive the numerical conversion from the CSAT-verbal scale to the SAT-verbal scale, the following numerical values for the slopes and intercepts of equations (6), (8), and (9) were applied to the constants in equation (10):

a = 1.1060; b = 1.6857 [from equation (6)],

A = 7.2533; B = 256.0000 [from equation (8)], and

A' = 6.3656; B' = 229.2008 [from equation (9)].

The resulting conversion equation for the verbal test is, therefore,

$$s_a = 0.9706 s_c - 8.5528$$
 (11)

The inverse equation, for converting the SAT-verbal scale to the CSAT-verbal scale, is:

$$S_c = 1.0303 S_a + 8.8119$$
 (11a)

The following equivalencies were determined from equations (11) and (11a):

From Equ	ation (11):	From Equat	ion (lla):
CSAT-verbal Score	Equivalent SAT-verbal Score	SAT-verbal Score	Equivalent CSAT-verbal Score
800	768	800	<b>(</b> 833)*
700	671	700	730
600	574	600	627
500	477	500	5 <b>2</b> 4
400	380	400	421
300	283	300	318
200	(186)*	200	215

More detailed conversion tables for the verbal tests are provided in Appendix I.

<sup>\*</sup>Scores lower than 200 or greater than 800 on the SAT are reported as 200 or 800, respectively. The same procedures will be employed on all future forms of the CSAT.



Directly parallel procedures were followed in deriving the equation for converting scaled scores on the CSAT-mathematical to scaled scores on the SAT-mathematical. In order to calculate the estimated values given in equations (1) to (4) for the mathematical tests, the correlation between the operational test and the 35-item equating section,\* as well as the related means and standard deviations, were prepared for each of the two mathematical samples, the sample consisting of the 3320 American students and the sample consisting of the 3315 Canadian students. These statistics are given in Table 7. Distributions of the operational mathematical and equating mathematical tests for the SAT and CSAT are given in Table 8.

The application of the data in Table 7 to equations (1) to (4) resulted in the following estimated values:

$$\hat{M}_{x_{t}} = 15.7479 ,$$

$$\hat{M}_{y_{t}} = 26.2079 ,$$

$$\hat{s}_{x_{t}}^{2} = 88.7126 ,$$
and
$$\hat{s}_{y_{t}}^{2} = 162.9277 .$$

These values were then applied to equation (5) to yield the equation,

$$Y = 1.3552X + 4.8663$$
 (12)

which permitted the conversion of scores from the raw score scale of the CSAT-mathematical to the raw score scale of the SAT-mathematical. In order to derive the conversion from the CSAT-mathematical scale to the SAT-mathematical

<sup>\*</sup>Recall that 10 mathematical "outlier" items had been removed from the original group of 45 items administered to both the American and the Canadian examinees.



Table 7

Correlations Between Operational and Equating Tests

for American and Canadian Samples

# Mathematical Tests

	American Sample		Canadian Sample
No. of Cases	3320		3315
Operational Test*  Mean Std. Dev.	22 <b>.</b> 8536 13 <b>.</b> 1620	<i>y</i>	18.2477 8.3968
Equating Test**  Mean  Std. Dev.	12.0973 8.5261		16.7493 7.2081
Correlation: Opnl. vs. Equating	•9348		.9219



<sup>\*</sup>The operational SAT-mathematical contains 60 items; the operational CSAT-mathematical contains 40 items.

<sup>\*\*</sup>The mathematical equating test consisted of 35 items (after the removal of 10 "outlier" items).

Table 8

Frequency Distributions for Operational and

Equating Sections of the SAT and CSAT

# Mathematical Tests

Raw	Ame <b>ri</b> can	Sample	Canadian	Sample
(Formula) Score	Operational SAT	Equating Section	Operational CSAT	Equating Section
58 - 57 54 - 57 54 - 55 54 - 55 54 - 45 54 - 45 54 - 45 54 - 45 55 - 47 56 - 47 57 57 57 57 57 57 57 57 57 5	8 8 19 13 31 39 42 70 78 81 105 98 135 170 164 162 180 167 190 188 176 153 132 179 146 100 121 81 62 35 19 7 5	8 24 46 82 91 129 161 156 221 212 234 261 286 293 263 263 220 218 112 32 7 0 1	4 33 23 80 100 125 175 173 200 194 255 306 304 311 276 257 198 128 89 48 25 92	25 24 91 122 149 210 260 301 302 338 368 313 267 213 136 104 53 29 6 3
No. of Cases	3320	3320 10 0073	3315	3315 16•7493
Mean	22.8536	12.0973	18.2477 8.3968	7.2081
Std. Dev.	13.1620	8.5261	0.2900	(•⊂∪⊙⊥

scale, the following numerical values from the slopes and intercepts of equations (6), (7), and (8) were applied to the constants in equation (9):

a = 1.3552; b = 4.8663 [from equation (6)]

A = 13.9550; B = 242.0000 [from equation (8)], and

A' = 8.5922; B' = 273.2077 [from equation (9)].

The resulting conversion equation for the mathematical test is, therefore,

$$s_a = 0.8347 s_c + 113.0219$$
 (13)

The inverse equation, for converting the SAT-mathematical scale to the CSAT-mathematical scale, is:

$$S_c = 1.1980 S_a - 135.4042$$
 (13a)

The following equivalencies were determined from equations (13) and (13a):

From Equa	ation (13):	From Equati	ion (13a):
CSAT-math Score	Equivalent SAT-math Score	SAT-math Score_	Equivalent CSAT-math Score
800	<b>7</b> 81	800	<b>(</b> 823)*
<b>7</b> 00	697	700	<b>7</b> 03
600	614	600	583
500	530	500	464
400	447	400	344
300	363	300	22 <sup>1</sup> 4
200	280	200	(104)*

More detailed conversion tables for the mathematical tests are provided in Appendix II.

<sup>\*</sup>Scores lower than 200 or greater than 800 will be reported as 200 and 800, respectively, on all future forms of the CSAT. The same procedures are employed on all forms of the SAT.



The incidental observation is made that the verbal and mathematical tests of the CSAT and the SAT are all quite similar in average difficulty. The estimated formula score means for the verbal tests for the combined SAT-CSAT verbal sample both represent a mean item difficulty of about 0.4 (0.41 for CSAT-verbal and 0.39 for SAT-verbal) in terms of the same combined verbal equating group. The estimated formula score means for the mathematical tests show a slightly greater difference; the mean item difficulty on CSAT-mathematical estimated for the combined mathematical equating group is 0.39 while the mean item difficulty on SAT-mathematical for the same group is slightly greater -- 0.43 -- but still very much in the same range. Although the reason for this difference can only be a matter for speculation, it is possible that part of it, at least, is attributable to the more generous timing allowances in the CSAT.

Ratios of the number of items to the estimated standard deviation for the combined samples were also obtained. These ratios were 4.87 for CSAT-verbal and 5.29 for SAT-verbal, and 4.25 for CSAT-mathematical and 4.70 for SAT-mathematical. These figures suggest the possibility that the mathematical tests yield more discriminations and may therefore be more reliable per unit test than the verbal tests. The figures also suggest the possibility that the tests of the CSAT are slightly more reliable than those of the SAT per unit test.

## Conversion Results

From equations (11) or (11a) and (13) or (13a), and from the conversion tables that follow those equations it is clear that, although the Canadian and American scales are similar in the region of the 500-score, there are detailed differences. The equivalent SAT-verbal scores are lower than the corresponding CSAT-verbal scores throughout the scaled score range, the differences ranging from -14 points at the CSAT-verbal score of 200 to -32 points at the CSAT-verbal



score of 800. At 500 the difference is midway between these extremes, -23 points.

The differences in the mathematical scales are greater, however, and for the most part in the opposite direction. At the CSAT-mathematical scaled score of 200, the difference is as much as 80 points; thus, a student with a score at the 200-score level on CSAT is estimated to be equivalent in ability to a student with a 280 score on SAT. At the opposite end of the scale, the difference is much smaller, but in the other direction; a student at the 800-score level on the CSAT scale would have a score equivalent to a score less than 800 on the SAT scale, namely 781. As expected, the difference at 500 is about midway between these; a score of 500 on CSAT-mathematical is equivalent to a score of 530 on SAT-mathematical.

## Summary

In this equating study the form of the CSAT administered in January 1971 was equated to the form of the SAT also administered in January 1971. The design for equating involved the administration of a one-half hour operational verbal section and a one-half hour operational mathematical section of the CSAT as "experimental" (nonoperational) tests to random samples of the January 1971 SAT group. Similarly, a one-half hour operational verbal section and a one-half hour operational mathematical section of the SAT were administered as "experimental" (nonoperational) tests to random samples of the January 1971 CSAT group. As a result of these symmetrical "cross-over" administrations, random samples of the American and the Canadian groups took the same two sections of verbal test materials. Other random samples of the American and Canadian groups took the same two sections of mathematical test materials.



were then item analyzed, and item difficulty (delta) plots were made to identify "outlier" items -- items that were not equally appropriate for the two groups of examinees. As a result of these analyses, 20 verbal and 10 mathematical items were identified as "outliers" and removed from the equating sections. Data on new samples were developed with the remaining 75 verbal equating items and 35 mathematical equating items to equate raw scores on that form of the CSAT to corresponding scores on that form of the SAT. With these equations in hand and with the equations converting each set of raw scores to its own reporting scale, conversion equations were developed relating the verbal and mathematical scales for the CSAT to the corresponding scales for the SAT.



# Reference

Angoff, W. H., "Scales, Norms, and Equivalent Scores," in R. L. Thorndike (Ed.) Educational Measurement (2nd ed.). Washington: American Council on Education, 1971, pp. 508-600.



Appendix I

Conversions Between CSAT-Verbal Scaled Scores and SAT-Verbal Scaled Scores

119	Equivalent CSAT-Verbal Score	514 503	232 233 233 233 233 233 233 233 233 233
s + 8.8119			
$S_c = 1.$	SAT-Verbal Score	06† 08†	8510 850 850 850 850 850 850 850 850 850 85
ation (11a):	Equivalent CSAT-Verbal Score	(833)*	(8823) (802) (802) 777 761 770 689 677 689 677 687 677 575 575 575 575 575
From Equation	SAT-Verbal Score	800	730 770 770 770 770 770 60 60 60 60 60 60 60 60 60 60 60 60 60
	[관리 ]		
- 8.5528	Equivalent SAT-Verbal Score	79†1 157	(1955) (1955) (1955) (1955) (1955) (1955) (1955) (1955) (1955)
s = 0.9705 s	CSAT-Ver	064	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0
From Equation (11):	Equivalent SAI-Verbal Score	768	758 7799 710 710 710 651 755 755 755 755 755 755 755 755 755 7
From Ec	CSAT-Verbal Score	800	780 770 770 770 770 680 670 670 670 670 670 670 670 670 670 67
	10		

are renorted as 200 or 800, respectively. The same pro-

Appendix II

Conversions Between CSAT-Mathematical Scaled Scores and SAT-Mathematical Scaled Scores

- 135.4042  Equivalent CSAT-Math Score
SAT-Math SAT-Math Score 490 470 470 470 470 470 470 470 47
(15a): s ivalent AT-Math Score (825)* (825)* (827)* (811)* 775 775 775 775 775 775 775 77
From Equation  Equation  Score  800  790  770  780  770  780  770  780  770  780  770  780  770  780  770  780  770  780  770  780  770  780  770  780  770  780  770  780  780  780  680  6
+ 113.0219
CSAT-Math  CSAT-Math  CSAT-Math  BCOTE  490  470  470  470  470  470  470  470
From Equation (13): Sa Equivalent Score 800 781 770 770 770 770 770 770 770 770 770 77
From Eque Score 800 770 770 770 750 650 650 650 650 650 550 550 550 550 5

\*Scores lower than 200 or greater than 800 will be reported as 200 and 800, respectively, on all future